

Valuing the impact of Self-rated Health and Social Support on life satisfaction among Hong Kong Chinese population

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Research Article

Keywords: Wellbeing valuation, self-rated health, social support, willingness-to-pay, social impact assessment, Chinese

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30

31 **Abstract**

32 **Background:** Research has highlighted that satisfaction in health and social support are key areas of
33 life affecting individual's wellbeing. Many social and public health initiatives use these two intervention
34 mechanisms to improve individual's wellbeing. For the purpose of cost-benefit assessment, there has
35 been growing interest in expressing these intervention effects in economic terms. However, only a
36 handful of studies have ever estimated these effects in economic terms, and none of which examined in
37 a Chinese context. The aim of this study is to extend this line of valuation work to estimate the implicit
38 willingness-to-pays on the effects of improving individuals' self-rated health status (SRH) and social
39 support (SS) on their life satisfaction in the Chinese population.

40 **Methods:** Using individual's life satisfaction data from a two-wave representative panel survey in Hong
41 Kong ($n = 1,109$), this study first conducted a cross-lagged analysis with structural equation modelling
42 technique to examine the causal effects of SRH and SS on life satisfaction, while simultaneously
43 adjusting their reverse causal influences. The use of this cross-lagged approach was the effort to
44 minimising the endogeneity problem. Then, substituting the respective estimates to the formulae of
45 compensating surplus, the marginal rate of substitution of SRH and SS with respect to individual's
46 equivalised monthly household were estimated and were then expressed as the willingness-to-pays on
47 the effect of improving individuals' SRH and SS on their life satisfaction.

48 **Results:** The cross-lagged analysis ascertained the casual effects of SRH ($\beta = 0.078$, 95%CI: 0.020,
49 0.151) and SS on individuals' satisfaction with life. The sample's marginal rate of substitution of *SRH*
50 and *SS* were found to be 1.28 (95%CI: 0.43, 2.15) and 1.36 (95%CI: 0.23, 2.49) respectively.
51 Translating into the concept of compensating surplus, the implicit monetary values of improving the
52 sample's *SRH* from "poor health" to "excellent health" and their *SS* from "little support" to "a lot of
53 support" are equivalent to an increase in their equivalised monthly household income by HK\$56,000
54 and HK\$39,400 respectively.

55 **Conclusions:** This study has implications for the cost-benefit assessment in wellbeing initiatives for
56 the Chinese population.

57

58 **Keywords:** Wellbeing valuation, self-rated health, social support, willingness-to-pay, social impact
59 assessment, Chinese

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64 **Background**

65 In recent years the concept of individuals' wellbeing has increasingly taken centre stage in public policy
66 discussion. An increasing number of policy thinkers have emphasised a key objective of health and
67 social policy is to improve individuals' wellbeing and collectively raise population welfare [1-4].
68 Worldwide, policy initiatives to enhance population wellbeing have been growing, such as enhancing
69 access to health care and building social support and community cohesion. Interest in assessing the
70 effectiveness of the wellbeing policy has been stacking up, and the preference to capture the impact in
71 some form of economic cost-benefit analysis is common [5, 6].

72 While cost-benefit analysis has been used in a wide array of policy domains, it can be
73 challenging when applied to wellbeing policy, since impacts of many of these initiatives such as
74 improving health status and social support are difficult to measure in economic terms (i.e., non-market
75 goods without direct monetary value). Previous research has developed valuation methods to resolve
76 this challenge, either through the revealed-preference approach or the stated-preference approach [7-
77 10]. However, through decades of research, it is known that multiple factors in the valuation process,
78 such as the cognitive biases caused either by the experimental settings or due to a non-equilibrium
79 market, could affect the valuation decision and hence leads to inadequate estimates [11-15].

80 Against this backdrop, recent research has developed the use of subjective wellbeing data, such
81 as life satisfaction, as an alternative non-market valuation approach [13, 16, 17]. In this context,

82 measures of life satisfaction that capture individuals' appraisals of their overall quality of life with
83 regards to past experience, expectations of the future, and comparison to others [18], are considered to
84 be indicators of one's utility [19]. Surveys that explore how exposure to some non-market
85 circumstances could causally affect individuals' life satisfaction are then interpreted as a direct
86 empirical approximation of how the goods of interest alter individual welfare. Then, by weighting
87 against the causal effect of household income on individuals' life satisfaction, they together form the
88 implicit willingness-to-pay (WTP), or the 'shadow price', of the goods of interest (i.e., additional
89 income to pay or accept in order to compensate for the changes in satisfaction with life for the losses or
90 gains in some particular conditions). Scholars have highlighted that this wellbeing valuation approach
91 does not rest upon individuals' decisions in the valuation process, hence avoiding potential cognitive
92 biases that could have been involved in the stated-preference and revealed-preference methods [20, 21].

93 The use of this wellbeing valuation approach has gained considerable traction in recent years
94 and has started to be applied in estimating the economic term for various health and social conditions.
95 Internationally, a handful of studies have used this wellbeing valuation approach to estimate the effects
96 of improving the self-rated status and social support on life satisfaction [20, 22-26]. However, the
97 estimates derived from these studies varied considerably. For instance, a study in the United Kingdom
98 estimated that the economic value of improving an individual's self-rated health from 'poor health' to
99 'excellent health' is equivalent to increasing his/her annual household income per capita £303,000 [22],
100 other studies reported a substantially smaller estimate (e.g., US\$1,644 - \$1,692 from Chandoevrit &
101 Thampanishvong's study [25]; see Additional file 1). Some variation of the estimates may be
102 methodologically related as some of the prior work ignored the endogeneity issue in their estimation
103 [22, 26, 27], which is a methodological concern that could introduce biases and cause an overestimation
104 of the monetary value of the non-market goods [27, 28]. Also, many of these existing valuation work
105 were conducted in the Western setting. To the best of the author's knowledge, only two studies have
106 used this wellbeing valuation approach to estimate the effects of improving health status and social
107 relationship on life satisfaction in an Asian context [25, 26], and none of which examined in a Chinese
108 context. Hence, to fill this research gap, the aim of this study is to extend the literature by using the

109 wellbeing valuation method to estimate the implicit WTPs of the effects of improving individuals' self-
110 rated health status and social support on their life satisfaction in the Chinese population.

111 The measure of individuals' satisfaction with life has been conceptualised as one of the major
112 indicators of one's overall wellbeing [29, 30], and a substantial body of research from the domains of
113 life literature has postulated that individuals' wellbeing depends on their appraisals on various aspects
114 of daily life activities [31, 32]. Previous research has highlighted that several areas such as health status
115 and social relationships, and among some others (e.g., material wellbeing, work, community etc.) are
116 the key areas that determine individuals' assessment of life [33, 34]. Self-rated health status, a measure
117 of one's overall physical and mental health condition, has been shown to be related to individuals'
118 satisfaction with life. Impaired health is associated with a decline in positive affect and an increase in
119 negative feelings, such as a depressed mood and anger [35]. In addition, with worsening health,
120 individuals' involvement in social activities is likely to be restricted, and they, therefore, have a limited
121 social connection [36, 37]. Extent research has also demonstrated that individuals having greater social
122 support tended to have a higher level of life satisfaction. Social support, which refers to individuals'
123 actual or perceived available social resources coming from their networks, such as family, friends, co-
124 workers, and from other community ties, can enhance bonding and serve as a buffer for negative
125 stressors. Social support as a protective factor has also been found to alleviate the burden of depressive
126 symptoms, generate happiness, and boost overall life satisfaction [38, 39]. This study focused on self-
127 rated health and social support as they are intervention mechanisms commonly used to enhance
128 individual welfare; deriving the monetary value of these two variables provides wide application in
129 cost-benefit policy analysis.

130

131 **Methods**

132 **Sample**

133 Data used in this study were extracted from the Hong Kong Panel Survey for Poverty Alleviation, which
134 is a two-wave household survey with a representative sample recruited through a stratified random
135 sampling strategy by 412 geographical constituency area. The panel survey involved conducting face-

136 to-face interviews with the household head to elicit a variety of information; Each interview lasted
137 approximately 60 minutes. Content of the survey and details of the sampling strategy has been reported
138 elsewhere [40]. The first wave of the panel survey was conducted between September 2015 and April
139 2016, in which 2,002 households were recruited. These households were interviewed again 12 months
140 after the first interview (the second wave), and 1,109 households were retained (retention rate: 50.2%).
141 Given that a notable proportion of the sample recruited in the first wave was lost in the follow-up, the
142 Inverse Mills analysis was conducted to explore the potential sample selection bias, and the result did
143 not indicate a substantial concern (Inverse Mills Ratio: 1.66; $p = 0.678$).

144 This study retrieved data on the sample's socioeconomic and demographic characteristics (i.e.,
145 age, sex, educational attainment, marital status, employment status, housing tenure, household size, and
146 monthly household income), and their status of self-rated health, social support, and satisfaction with
147 life. The sample's self-rated health status was assessed by a single item eliciting their overall health (1
148 = poor health to 5 = excellent health). Previous studies have shown that this is a valid and reliable
149 measure for one's overall physical and mental health condition [41], reflects participation in health risk
150 behaviours [42], and predicts mortality risk [41]. The sample's level of social support was assessed
151 using three items exploring the levels of perceived available support from their family members and
152 friends (1 = little support to 4 = a lot of support). Three types of social support were elicited, including
153 instrumental support, informational support, and emotional support [43]. For life satisfaction, the five-
154 item Chinese version of the Satisfaction with Life Scale (SWLS) was employed to elicit sample's
155 satisfaction with life. The total scores of the SWLS range from 5 (extremely dissatisfied) to 35
156 (extremely satisfied). The Chinese version of the SWLS has been validated and was found to have
157 adequate validity and reliability [44].

158

159 **Statistical analyses**

160 Analysis in wellbeing valuation work can be generally conceptualised as having two-part. The first part
161 involves a regression analysis estimating the effects of some non-market conditions (i.e., self-rated
162 health and social support) on life satisfaction and respectively the effect of an economic indicator (e.g.,
163 household income) on life satisfaction. Estimates of these effects would then form the marginal rate of

164 substitution of the non-market conditions with respect to the economic indicator. As mentioned earlier,
165 one major methodological concern that could introduce biases in this part of estimation is the
166 endogeneity issue [45] and failing to account it can cause an overestimation of the monetary value of
167 the non-market goods [27, 28]. Studies typically used instrumental variables to handle the problem, but
168 the use of the instruments also has its own constraints. Specifically, variables that qualify as an
169 instrument should fulfil the independence assumption, i.e., the instruments should not share causes with
170 the outcome variables (e.g., life satisfaction). Researchers have highlighted that it is very difficult to
171 identify appropriate instruments that meet this assumption [27], as “almost every factor determines life
172 satisfaction” [46].

173 An alternative way to tackle the endogeneity issue in wellbeing valuation is the use of a cross-
174 lagged analysis [47]. A typical cross-lagged model would examine six effects between two variables.
175 For instance, applying in an examination of the relationships between SRH and life satisfaction (LS),
176 two reflects the temporal autoregressive effects of the variable of SRH and LS (β_a : the causal effects
177 of SRH_1 on SRH_2 and LS_1 on LS_2 , respectively) and two cross-sectional correlations between SRH and
178 LS (β_{cor} : the correlations between SRH_1 and LS_1 and SRH_2 and LS_2 , respectively). The remaining two
179 effects examine the reciprocal nature between SRH and LS , which is the causal effect of SRH_1 on LS_2
180 (β_c) and the reverse causal effect of LS_1 on SRH_2 (β_{rc}).

181 In the context of this valuation work, the main subject of interest is the β_c (e.g., the causal effect
182 of SRH_1 on LS_2), as it will be used for calculating the marginal rate of substitution. Specification of the
183 other five effects are ways to reduce the endogeneity problem in the estimation process. Specifically,
184 given the endogeneity bias typically arises from reverse causality (β_{rc}) and omitted influences on the
185 dependent variable (e.g., temporal autoregressions, β_a , and cross-sectional correlations, β_{cor}) [45], the
186 cross-lagged model explicitly addressing them help to reduce these potential influences in the
187 estimation of the causal effect of SRH on LS . The cross-lagged model is computed based on a structural
188 equation modelling (SEM) technique, which is a form of simultaneous regression technique often used
189 in the work of subjective wellbeing valuation. Previous studies have used SEM to examine the causal

190 determinates of life satisfaction [48], and recent valuation studies have used this technique in valuing
191 environmental goods [46].

192 In this study, a cross-lagged model was specified to examine the relationships among the
193 variables of SRH, SS, HI, and LS. The full model analysed all the effects between the pairs of HI, SRH,
194 SS, and LS within and across the two timepoints. In addition, sample's sociodemographic
195 characteristics were hypothesised as covariates and were regressed on the outcome variables (HI, SRH,
196 SS, and LS). To check the model adequacy, the model's comparative fit, parsimony correction, and
197 absolute fit were assessed using the goodness-of-fit indices commonly used in the SEM technique,
198 including the comparative fit index (CFI) and Tucker-Lewis index (TLI), the root mean square error of
199 approximation (RMSEA), and root mean square residual (MRSR). According to Hu and Bentler's
200 principle [49], the model is considered as having an adequate fit with the data when the CFI and TLI
201 are $>.90$, the RMSEA <0.06 , and SRMR $<.08$.

202 For the SEM estimation, data merged across the two waves of the survey was used for model
203 fitting. Sociodemographic characteristics of the sample and their scores of *SRH*, *SS* and *LS* across both
204 waves are reported in Table 1. Samples' equivalised monthly household income was derived from
205 dividing their monthly household income to the square root of the household size. Also, the second
206 wave sample's monthly household income was deflated by the consumer price index for a valid
207 reflection of the real changes of household income. Following convention, standardised coefficients of
208 the cross-lagged model in SEM were reported.

209

210 **Table 1** Sample's socioeconomic characteristics and descriptive statistics on the main variables
211 of interest (N = 1,109)

212

213 Building upon the findings from the cross-lagged analysis, the second stage of valuation procedure
214 was to estimate the WTPs of *SRH* and *SS* on *LS*. Specifically, the estimated causal effects (β_c) of *SRH*₁
215 and *SS*₁ on *LS*₂ as well as *HI*₁ on *LS*₂ were used for computing the marginal rate of substitutions of
216 *SRH* and *SS* with respect to *HI*, respectively. Subsequently, it is substituted into the formulae of
217 compensating surplus [21] expressed as,

218

$$\bar{w} = e^{[\ln(\bar{w}) - \frac{\beta_a}{\beta_b} \Delta x]},$$

219

220 where \bar{w} is sample's median equivalised monthly household income, β_a/β_b was the marginal rate of
221 substitution, and Δx is the change in the non-market condition (i.e., *SRH* or *SS*). Estimates yield from
222 the formulae reflects the WTP of the condition, which can be interpreted as (hypothetically) the amount
223 of additional equivalised monthly household income required to equate the changes in life satisfaction
224 in relation to the increase or decline either in *SRH* or *SS*.

225

226 **Results**

227 **Cross-lagged analysis of the relationship between SRH, SS, HI, and LS**

228 Using the SEM technique, the hypothesised cross-lagged model denoting the complex relationships
229 among *SRH*, *SS*, *LS*, *HI*, and the socioeconomic variables as covariates were examined. Model
230 diagnostics suggested that there was an adequate fit between the specified model and the empirical data,
231 as all the goodness-of-fit indices met the Hu and Bentler's threshold (CFI = .93, TFI = .92, RMSEA
232 = .05, SRMR = .07).

233 Figure 1 summarises the major findings of the cross-lagged analysis on the relationships among
234 *SRH*, *SS*, *HI*, and *LS*. There are a number of noteworthy observations. First, the model indicated that
235 there were strong causal associations among the four variables across the timepoints (the curved arrows
236 in Figure 1; e.g., $HI_1 \rightarrow HI_2: \beta = 0.458, 95\%CI: 0.296, 0.619$). The model also pointed to some significant
237 cross-sectional correlations between the variables of *SRH*, *SS* and *HI*, and *LS* at each wave (i.e., SRH_1
238 $\leftrightarrow SS_1: \beta = 0.107, 95\%CI: 0.037, 0.196$; $SRH_2 \leftrightarrow SS_2: \beta = 0.122, 95\%CI: 0.049, 0.199$; $HI_2 \leftrightarrow LS_2: \beta$
239 $= 0.101, 95\%CI: 0.042, 0.160$, not shown in Figure 1 for the purpose of parsimony). Second, mutual
240 causality was detected among the relationships between *SRH* and *LS* and between *HI* and *LS*.
241 Specifically, the model indicated the casual effects of SRH_1 and HI_1 on LS_2 were significant; one
242 standard deviation increases in *SRH* and *HI* were casually associated with a standard deviation increase
243 in *LS* by the unit of 0.078 (95%CI: 0.020, 0.151) and 0.061 (95%CI: 0.009, 0.112), respectively. At the
244 same time, the reverse causal influences of LS_1 on SRH_2 and HI_2 were also significant. Comparing the

245 magnitude of the estimates, the reverse causal effects of LS_1 on SRH_2 and HI_2 were even greater than
246 their respective causal effects (the effect of LS_1 on SRH_2 , $\beta = 0.086$, 95%CI: 0.020, 0.151; the effect of
247 LS_1 on HI_2 , $\beta = 0.073$, 95%CI: 0.020, 0.151). For the relationship between SS on LS , while the causal
248 influence of SS on LS was found to be significant ($\beta = 0.083$, 95%CI: 0.014, 0.152), its respective
249 reverse causal effect was however not detected. Overall, R^2 suggested the model accounted for 45%, of
250 the variance of the variables of LS_2 .

251

252 **Fig. 1** The causal and reverse causal effects of household income, self-rated health, and
253 social support on life satisfaction

254

255 **The implicit WTPs of SRH and SS on LS**

256 Based on the results of the cross-lagged analysis (Figure 1), the marginal rate of substitution of SRH
257 and SS on LS were estimated to be 1.28 (95%CI: 0.43, 2.15) and 1.36 (95%CI: 0.23, 2.49), respectively.
258 This is interpreted as respondents' increase (decrease) in LS owing to their improvement (deterioration)
259 in SRH and SS is equivalent to the effect of increasing (reducing) their HI by 1.28 and 1.36 unit of
260 standard deviation respectively. Substituting it into the formulae of compensating surplus, the implicit
261 WTPs of improving the sample's SRH from "poor health" to "excellent health" and enhancing their SS
262 from "little support" to "a lot of support" are thus equivalent to an additional increase in their
263 equivalised monthly household income by HK\$56,000 (95%CI: HK\$26,900, HK\$68,600) and
264 HK\$39,400 (95%CI: HK\$10,800, HK\$48,300), respectively.¹

265

266 **Discussion**

267 Drawing data from a representative panel survey in Hong Kong, this study examined the causal effects
268 of change in self-rated health and social support on individuals' overall satisfaction with life. Adopting
269 the wellbeing valuation method [13, 16, 17], this study expressed these causal effects in an economic

¹ Estimation of the WTPs of SRH and SS were based on the formulae of compensating surplus. In this estimation, \bar{w} is the sample's equivalised monthly HI : HK\$14.2k.

270 term (i.e., WTPs). To the best of the author's knowledge, this is the first to derive the shadow price of
271 self-rated health status and social support in a predominantly Chinese society. It extends the wellbeing
272 valuation literature on health and social outcomes as it is one of the few conducted in an Asian context.

273 This research has implications for the practice of social impact evaluation. Specifically, the
274 derived shadow price of self-rated health and social support provide ways to tie social impact evaluation
275 with some forms of cost-benefit analysis (e.g., social return on investment), which is a tool where
276 policymakers commonly relied on for decision making [5]. Currently, in the absence of culturally-
277 attuned economic proxies, practice among social impact evaluators sometimes apply economic proxies
278 derived from other socio-cultural context (e.g., Western settings) to monetise social non-market
279 outcomes, despite the notable cultural difference on individuals' evaluation on domains of life [50, 51].
280 This study offers a way to make socio-cultural embedded cost-impact assessment plausible in the
281 Chinese context.

282 Findings of the cross-lagged analysis illustrated that the relationships among the examined
283 variables are complex. First, the model reflected the existence of significant autoregressive correlations
284 within each main variable of interest (i.e., SRH, SS, HI, and LS). This is fairly intuitive given the
285 timeframe between two data collection timepoints was not long (i.e., approximately twelve months). In
286 addition, cross-sectional correlations among the variables of interest were also detected (e.g., between
287 SRH and SS). The finding is in line with the domains-of-life literature highlighting the intertwining
288 nature among individuals' various aspects of life [31]. Furthermore, the results not only showed the
289 existence of the causal influences of self-rated health, social support, and equivalised monthly
290 household income on samples' satisfaction with life, it also suggested that the reverse causal
291 mechanisms among these variables were significant. Particularly, the sample's baseline life satisfaction
292 was positively and casually associated with their equivalised monthly household income and self-rated
293 health, respectively. Previous research has illustrated that individuals with greater life satisfaction are
294 more productive and are more likely to be promoted and hence may earn a higher income [20, 52, 53].
295 Positive affect has also been shown to improve the immune system and reduce a person's susceptibility
296 to illness [54, 55]. Taken together, results of this study, on the one hand, add to a large pool of literature
297 illustrating individual's wellbeing as constitutes of multiple areas of life, it nevertheless reflects the

298 plausibility of the situational mechanism in which one's overall wellbeing affects their different spheres
299 of life [30].

300 The existence of the reverse causal influences existed in the relationship between sample's HI and
301 LS and between SRH and LS further affirms the need to attend the endogeneity problem in the
302 regression analysis, in order to avoid an overestimation of the implicit monetary value of health and
303 social outcomes [27, 28]. In addition, this study also shows collinearity likely exists between exogenous
304 variables (i.e., social and health conditions), it serves as a reminder for researchers to attend to this
305 problem in future wellbeing valuation studies. The cross-lagged approach simultaneously addresses the
306 cross-sectional associations between the exogenous variables and the reverse-causal influences could
307 be a plausible way to minimise the endogeneity issue in this line of valuation work.

308 This study is the first to derive the implicit monetary values of self-rated health status and social
309 support on individual's life satisfaction in predominantly Chinese society. Also, unlike some prior work,
310 this study employed statistical procedure in order to minimise potential bias in the estimation process
311 arising from the endogeneity issue. However, it is not without limitations. First, this study was
312 secondary data analysis, and hence the selection of variables for model estimation was restricted by the
313 original survey design. Second, although it is aware that the causal relationship between *SRH*, *SS* and
314 *LS* may vary across age and sex [56], this study, however, did not perform sub-demographic analysis
315 owing to a limited power due to the restricted sample size. While the WTPs estimated in this study may
316 reflect the implicit monetary value of self-rated health and social support among the general population,
317 caution should be exercised when generalising it to particular sub-demographic groups. Further research
318 is warranted to estimate the variation. Third, as in almost all wellbeing valuation analysis, it cannot rule
319 out the possibility that endogeneity still existed in the estimates, as the problem can arise from other
320 sources (e.g., measurement error) that may not be able to be controlled from a statistical standpoint.

321

322 **Conclusions**

323 This study offers a way to monetise the impact of increasing individuals' self-rated health status and
324 social support on their overall satisfaction with life. It has practical implications for the use of social
325 cost-benefit analysis in assessing wellbeing policy initiatives for the Chinese population.

326

327 **Supplementary Information**

328 **Additional file 1:** Implicit WTPs of self-rated health status and social support from wellbeing valuation
329 studies in an Asian context.

330

331 **Abbreviations**

332 SRH: Self-rated health; SS: Social Support; CI: Confidence Interval; WTP: willingness-to-pay; LS:
333 Life Satisfaction; HI: Household Income; SEM: Structural Equation Modelling; CFI: Comparative Fit
334 Index; TLI: Tucker-Lewis Index; RMSEA: Root Mean Square Error of Approximation; MRSR: Root
335 Mean Square Residual

336

337 **Declarations**

338 **Ethics approval and consent to participate**

339 The ethical considerations were approved by the Human Research Ethics Committee, the University of
340 Hong Kong (HREC Reference Number: EA1506006). The author declares that all methods were
341 performed in accordance with the relevant guidelines and regulations. Informed consent was obtained
342 from all individual participants included in the study.

343

344 **Consent for publication**

345 Not applicable.

346

347 **Availability of data and materials**

348 The datasets used and/or analysed during the current study available from the corresponding author on
349 reasonable request.

350

351 **Competing interests**

352 The author declares that he/she has no competing interests.

353

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356

357 **Author's contributions**

358 CHC conceived of the study, designed, executed the statistical analyses, and drafted the manuscript.

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362

363 **Author's information**

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365

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Figures

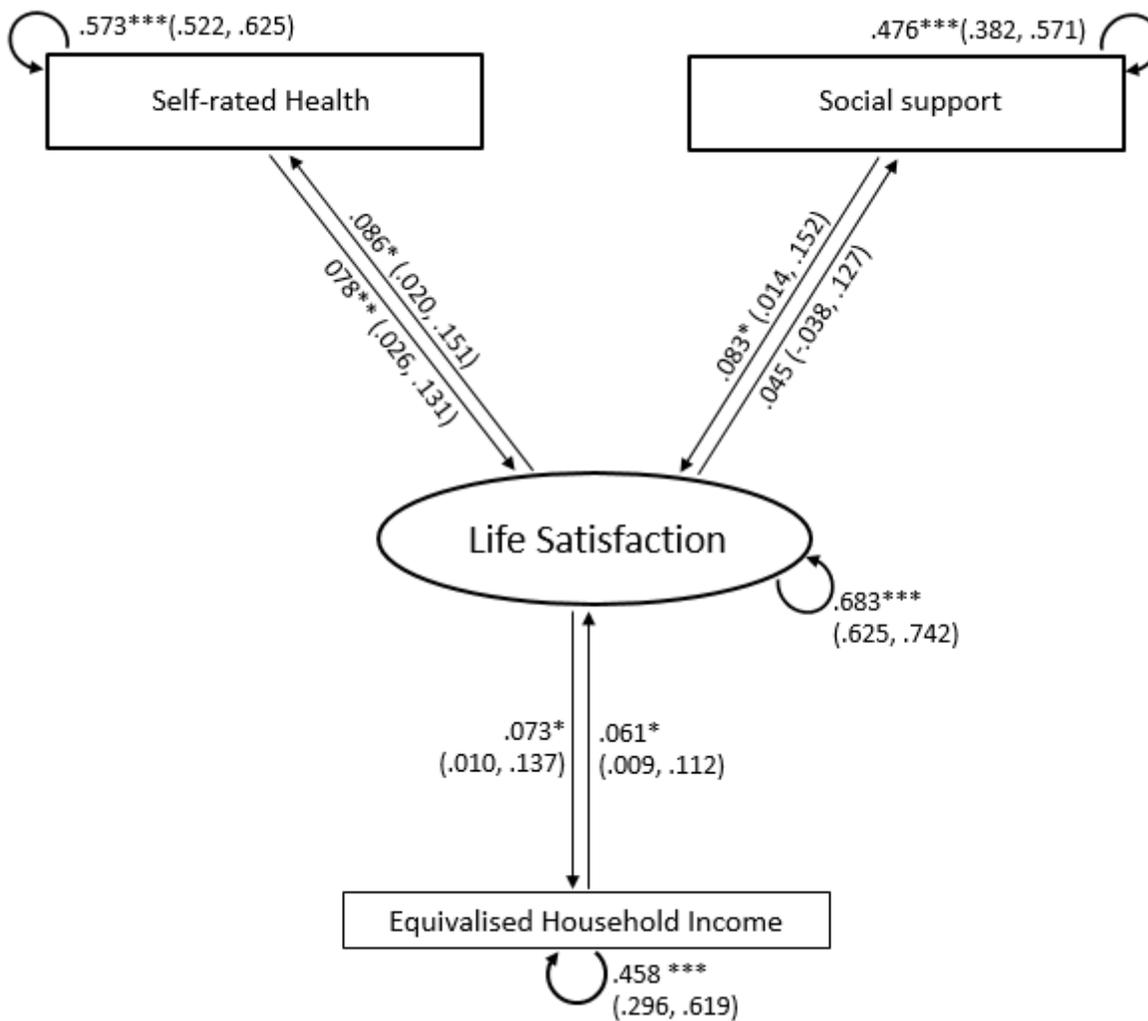


Figure 1

The causal and reverse causal effects of household income, self-rated health, and social support on life satisfaction

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