

1 **Title: Global Arthropod beta-diversity is spatially and temporally structured by latitude**

2 **Short running title: Latitudinal arthropod diversity gradient**

3 **Abstract:**

4 Aim: Global gradients in species biodiversity may or may not be associated with greater  
5 species replacement closer to the equator. Yet, empirical validation of these patterns has so far  
6 focused on less diverse taxa, with comparable assessments of mega-diverse groups  
7 historically constrained by the taxonomic impediment.

8 Location: Global

9 Time period: 2010-2017

10 Major taxa studied: Terrestrial arthropods

11 Methods: Here we assess the temporal and spatial dynamics of arthropod communities using a  
12 beta-diversity framework, drawing on samples across 129 globally distributed monitoring  
13 sites. Overall, we encountered more than 150,000 unique barcode index numbers (BINs) (i.e.  
14 species proxies). We assessed between site differences in community diversity using beta-  
15 diversity and the partitioned beta-diversity components of species replacement (i.e. turnover)  
16 and richness difference (i.e. a measure of nestedness).

17 Results: We show that global differences in community compositional change are linked to  
18 latitudinal, spatial, and temporal gradients. General global beta-diversity trends remained  
19 consistent across biogeographic regions, with beta-diversity (dissimilarity) increasing with  
20 decreasing latitude, greater spatial distance and greater temporal distance. By contrast, species  
21 replacement and richness difference patterns varied across biogeographic regions, suggesting  
22 different underlying processes are shaping regional biodiversity patterns. Latitudinal effects  
23 on species replacement or richness difference were significant for 3 out of 5 regions.

24 Comparably we found significant spatial distance relationships with species replacement or  
25 richness difference for 3 out of 5 regions. Temporal distance was significantly associated with  
26 species replacement or richness difference for all 5 regions.

27 Main conclusion: The general expectations of the latitudinal diversity gradient (LDG) are  
28 supported using a large, extensive global sampling effort. Our findings generally support the  
29 global LDG expectations, however, partitioned assessment of beta-diversity indicates the  
30 underlying processes driving the general global pattern, here species replacement and richness

31 difference, may be regionally linked to differences in seasonality effects or variation in spatial  
32 distribution of environmental factors.

33 **Key words: barcoding, COI, biogeography, biodiversity, metacommunity,**  
34 **spatiotemporal**

## 35 **Introduction**

36 Biodiversity is influenced by many factors, including environmental, evolutionary, biotic  
37 interactions, and stochastic, resulting in a global distribution of over 2 million species (IUCN,  
38 2022). The amazing level of global biodiversity is essential for life, including ours, in the  
39 form of various environmental services, including nutrient cycling, food security, and waste  
40 management, etc. (Felipe-Lucia et al., 2020). As such, understanding what factors shape  
41 biodiversity across time and space, particularly at the global scale, is of interest to a wide  
42 range of researchers in ecology, evolutionary biology, conservation and invasive species  
43 management, agriculture, medical science and many others. In the first instance it is generally  
44 accepted that biodiversity is expected to scale with latitude, increasing toward the tropics, a  
45 phenomenon referred to as the latitudinal diversity gradient (LDG) (Hillebrand, 2004). The  
46 underlying mechanism(s) for the LDG expectation are not definitive, however, with over 30  
47 current hypotheses built on varying degrees of ecological, evolution and environmental  
48 complexity (Willig, Kaufman, & Stevens, 2003). More recent efforts have also discovered  
49 exceptions to the LDG, which are primarily regional (Dowle, Morgan-Richards, & Trewick,  
50 2013; Mannion, Upchurch, Benson, & Goswami, 2014; Qian & Ricklefs, 2007; Rahbek et al.,  
51 2019). The many hypotheses advanced to explain global biodiversity patterns are often  
52 difficult to test, particularly across taxonomic groups or at large spatial and temporal scales  
53 (Kinlock et al., 2018; Pontarp et al., 2019). However, we can gain insights into biodiversity  
54 patterns using pairwise site assessment of total beta-diversity and its associated decomposition  
55 components.

56 The beta-diversity (i.e.  $\beta$ -diversity) framework provides a robust means to assess differences  
57 in biodiversity between communities, which can, in turn be used to determine spatio-temporal  
58 or environmental response (Borcard, Gillet, & Legendre, 2018). Beta-diversity is the  
59 compositional difference (i.e. dissimilarity) between two communities (Borcard et al., 2018).  
60 This compositional dissimilarity between communities is expected to arise from two key  
61 processes: (i) species replacement (i.e. turnover), quantifies the change in community  
62 composition due to non-overlapping species and (ii) richness difference, the species gain or  
63 loss between two communities. Communities with high species replacement can result from  
64 strong environmental forcing, competition or historical disturbances (Victorero, Robert,  
65 Robinson, Taylor, & Huvenne, 2018). Richness difference (i.e. nestedness) can be caused by  
66 local species disappearing from a location (i.e. localized extinction), differing niche diversity  
67 or other processes resulting in the gain or loss of species (Lazarina et al., 2023; Schmera,

68 Podani, & Legendre, 2020). Following the beta-diversity framework from Podani *et al.* (2013)  
69 total beta-diversity itself can be separated into the components of (i) species replacement and  
70 (ii) richness difference, which together sum to the total beta-diversity measure. Subsequently  
71 the beta-diversity partitioning framework provides a means to investigate the potential links  
72 between global biodiversity patterns and the underlying processes associated with their  
73 formation across different species groups and ecological dimensions.

74 While several studies have investigated global biodiversity patterns, predominately using the  
75 LDG, they were constrained by three major limitations. First, they were mainly based on  
76 meta-analyses, as they combined data collected using different methodologies at different  
77 spatial and temporal resolutions (e.g. seasonality) (Hillebrand, 2004; Kinlock *et al.*, 2018).  
78 While such data have high heuristic value, they are often affected by biases emerging from  
79 the varied sampling techniques underlying the individual data points (Koricheva & Gurevitch,  
80 2014). Second, the few studies which have sampled communities using standardized methods  
81 to estimate differences in biological communities across broad latitudinal ranges have  
82 generally ignored the effects of temporal variability (i.e. seasonality) within or between  
83 sampled communities (Rivadeneira *et al.*, 2015; Valdés *et al.*, 2015; Zhang *et al.*, 2020). If  
84 differences in biological communities are only assessed across space, estimates of site-  
85 specific diversity ignore the well-established importance of local temporal variation in  
86 biodiversity (Bista *et al.*, 2018; Fisher, Frank, & Leggett, 2010; Seymour *et al.*, 2021). If site-  
87 specific diversity comparisons are made across different time points, the potential estimates of  
88 patterns of beta-diversity (i.e. diversity difference between sites) in space will alter patterns of  
89 beta-diversity in time, i.e. the scope for spatiotemporal interactions (Gaston, Blackburn, &  
90 Spicer, 1998). Third, prior studies have either examined less diverse taxa (Field *et al.*, 2009;  
91 Gaston *et al.*, 1998; Hillebrand, 2004) or have generalized patterns emergent from local  
92 studies to the global scale (Novotny *et al.*, 2007; Qian & Ricklefs, 2007). Thus far, efforts to  
93 assess global patterns with standardized sampling methods have not been undertaken for  
94 taxonomic groups that comprise the bulk of global biodiversity.

95 With regards to the LDG, as the most prominent ecological assumption of global biodiversity  
96 distribution, species replacement is expected to increase at lower latitudes, reflecting greater  
97 habitat specialization and smaller ranges in more seasonally stable environments (Borcard *et al.*,  
98 2018; Qian, 2009; Soyninen, Heino, & Wang, 2018). By contrast, richness difference is  
99 expected to increase with latitude, reflecting recent recolonization from a shared species pool  
100 following deglaciation (Dobrovolski, Melo, Cassemiro, & Diniz-Filho, 2012). Alternatively,

101 species replacement may increase with latitude, which could reflect historical selection for  
102 species adapted for colder periods or to stronger changes in seasonality (Mateo et al., 2016).  
103 Different latitudinal patterns in richness difference, be it decreasing with latitude or unimodal,  
104 could indicate spatio-temporal disturbance patterns linked to regular or historical extinction  
105 events (Socolar, Gilroy, Kunin, & Edwards, 2016). A lack of general patterns across multiple  
106 regions or continents may also indicate inconsistent patterns of global biodiversity which may  
107 suggest that stochastic processes predominate over expected environmental gradient filtering  
108 of community assembly. Hence, comparisons between species replacement and richness  
109 difference can provide insights into the processes influencing global biodiversity patterns.

110 In this study, we adopt DNA-based methods to characterize beta-diversity for a highly diverse  
111 lineage of animals: terrestrial arthropods (Mora, Tittensor, Adl, Simpson, & Worm, 2011).  
112 We comprehensively sample 129 sites across the globe for an average of 22 sequential weeks  
113 each, encountering more than 150,000 different Barcode Index Numbers (BINs), a species  
114 proxy (Ratnasingham & Hebert, 2013). We calculated and partitioned beta-diversity into its  
115 species replacement and richness difference components in both space and time (e.g.  
116 seasonality) to determine how global biodiversity patterns relate to latitude, distance, and  
117 time.

## 118 **Methods**

119 The Global Malaise Trap Program (GMTP) was initiated in 2007 with the goal of observing  
120 global-scale spatiotemporal arthropod biodiversity dynamics (Figures 1 & 2). Between 2010  
121 and 2016, one or more Malaise traps were deployed at 129 sampling sites in 28 countries,  
122 with repeated weekly sampling ranging from 2-104 weeks (Figure 2). In total, institutions in  
123 28 counties contributed to the Global Malaise Trap Program (Figure 1; Table S1), which  
124 jointly produced the first set of global biodiversity data for terrestrial arthropods based on a  
125 uniform barcode sampling method. Details on GMTP standardized sampling protocols (e.g.  
126 trap type, sampling method, data curation) used for this study are outlined in detail in  
127 deWaard *et al.* (2019), but we provide a short summary here. A standard Townes-style  
128 Malaise trap was deployed at each sampling location (hereafter site) and captured arthropods  
129 from each trap were harvested weekly (hereafter trapping event). Arthropod specimens  
130 captured from each trapping event were sorted, photographed, and processed individually.  
131 Analysis began with each specimen identified morphologically to a taxonomic order and  
132 registered on the Barcode of Life Data Systems (BOLD). DNA from each specimen was then  
133 extracted and used to amplify and Sanger sequence the standard COI barcode region (Hebert,

134 Cywinska, Ball, & deWaard, 2003). The resulting COI sequence data was uploaded to the  
135 BOLD database, linking each specimen's morphological identification to its COI barcode  
136 sequence. For each trapping event, all specimens were sequenced, except when a particular  
137 morphospecies was represented by more than 50 individuals, in which case a subset of the  
138 individuals were sequenced to confirm that the specimens did indeed represent a single unique  
139 BIN (deWaard et al., 2019). The final GMTP dataset includes 1.2 million barcode records and  
140 155,185 unique barcode index numbers (BINs) (Ratnasingham & Hebert, 2013). Prior studies  
141 have established a strong correspondence between BINs and species identification in insect  
142 groups with well-established taxonomy, thereby justifying the recognition of BINs as species  
143 proxies (Ratnasingham & Hebert, 2013).

144 Temporal differences (i.e. distance in time) between each pair of sampling events was  
145 calculated using circular statistics by first determining the Julian day of the two sampling  
146 events and taking two measures (1) the absolute difference between the two Julian days  
147 divided by 0.986 (0.986 degrees = 1 day) and (2) 360 minus the absolute difference between  
148 the two Julian days divided by 0.986. The minimum value between measure (1) and (2) was  
149 then used as the distance in time between the two sampling events. Here we refer to difference  
150 in time and seasonality since the study spans multiple seasons (Figure 2) and since seasonality  
151 differs drastically between different global locations. Distance between each pair of sampling  
152 locations (i.e. distance in space) was calculated as the geographic distance between site pairs  
153 using the function `distHaversine` in the R package `geosphere` (Hijmans, Williams, Vennes, &  
154 Hijmans, 2017). Mean absolute latitude was calculated between each pair of sites along the  
155 LDG. To understand how this metric behaves, consider a site pair in which both members are  
156 at the Equator. In this case, their mean absolute latitude is  $0^\circ$  - which also applies to two  
157 samples from the same trap at the Equator. For a trap pair at the North Pole, the mean  
158 absolute latitude is  $90^\circ\text{N}$ ; for a trap pair with its members on the North vs South Pole, it will  
159 be  $90^\circ$ , and for a trap pair of which one member sits on the North Pole and the other at the  
160 Equator, mean absolute latitude will be  $45^\circ\text{N}$ .

#### 161 Statistical analyses

162 Community data were converted to presence absence data for calculations and analyses of  
163 diversity. Beta-diversity and its components were calculated as Jaccard dissimilarity using the  
164 Podani family of indices, which is a "true" beta-diversity estimate that is unaffected by the  
165 species pool (i.e., gamma-diversity) (Schmera et al., 2020). Total beta-diversity (here Jaccard  
166 dissimilarity) and the associated components of species replacement and richness difference

167 were calculated for each site pair using the function `beta.div.comp` in the R-package  
168 *adespatial* (Borcard et al., 2018). We do note that there are alternative beta-diversity  
169 partitioning methods (Baselga, 2010b; Borcard et al., 2018; Schmera et al., 2020). The Podani  
170 family was utilized here as it does not overestimate diversity differences and provides a  
171 “true” diversity estimate that is unaffected by the total species pool (Borcard et al., 2018;  
172 Schmera et al., 2020), but see also alternative true-diversity based partition approaches (Engel  
173 et al., 2021; McGlinn et al., 2019). Utilizing a “true-diversity” allows for independent  
174 measures of alpha, beta, and gamma diversity. Whereas alpha-diversity reflects within site  
175 variation, beta-diversity may either reflect between site variation independently or  
176 dependently (i.e. scaling with alpha-diversity) depending on the measure used (Koleff,  
177 Gaston, & Lennon, 2003). Using an independent (i.e. “true”) measure of beta-diversity  
178 becomes more important when comparing beta-diversity measures, dependence of alpha and  
179 beta diversities, to avoid compromising interpretation that may actually reflect within-site  
180 instead of between site observations (Baselga, 2010a).

181 Pairwise values of total beta-diversity, species replacement, and richness difference were  
182 calculated for all trapping event pairs by taking the lower triangle values from the associated  
183 distance matrix. As there were 2,412 trapping events in total,  $N = 2,907,666$  pairs of trapping  
184 events were included in our analyses.

185 Using linear regression, we modelled each pairwise beta-diversity component as a separate,  
186 univariate function of distance in space, distance in time, mean latitude, and the interactions  
187 mean latitude  $\times$  distance in space and mean latitude  $\times$  distance in time. Here, the two  
188 interaction terms are of key interest in explicitly testing whether the rate of beta-diversity,  
189 species replacement or richness difference in space or time, respectively, varies detectably  
190 with latitude.

191 Our data are not fully balanced as the number of data points per site, and hence pairs of sites,  
192 varies. In the analyses, we wished to give each site, and pair of sites, an equal weight in the  
193 analyses. If  $n_{s_1s_2}$  is the number of pairs of trapping events for which one trapping event  
194 belongs to site  $s_1$  and the other trapping event belongs to site  $s_2$ , in an unweighted regression  
195 this pair of sites would achieve the total weight of  $n_{s_1s_2}$ , and thus sites with more data would  
196 contribute disproportionately to our analyses. To account for the unbalanced sampling effort in  
197 our models, we applied a weighted linear regression, where the weight for each data point was  
198 set to  $1/n_{s_1s_2}$ , so that the total weight was equal among all pairs of sites.

199 We note that data points are not independent of each other, because each data point in the  
200 linear model involves a pair of samples that are correlated in time and space. For this reason,  
201 we did not perform significance tests based on output from the linear model but instead  
202 employed the following permutation approach to determine significance for each of the  
203 explanatory variables in our models.

204 Given the unequal temporal sampling and spatial sampling design across the multiple GMTP  
205 project datasets we used a series of permutation test to assess the significance of each  
206 explanatory variable (Chihara & Hesterberg, 2018). For each permutation test we assessed the  
207 significance of each explanatory variable individually, including distance in space, distance in  
208 time, mean latitude, and the interactions mean latitude  $\times$  distance in space and mean latitude  $\times$   
209 distance in time, by doing the following. We first calculated the log-likelihood ratio between  
210 the model where the focal explanatory variable was included (the full model), and the model  
211 where the focal explanatory variable was excluded (the reduced model). We compared the  
212 observed log-likelihood ratio to its null distribution which we computed by permuting the data  
213  $N=1,000$  times, with the permutation scheme detailed below for each specific test. In general,  
214 if the log-likelihood ratio for the full vs. reduced model fitted to the actual data was greater  
215 than the log-likelihood ratio for the full vs. reduced model fitted to the permuted data for at  
216 least 95% of the permutation outcomes, the explanatory variable was deemed significant  
217 (Chihara & Hesterberg, 2018).

218 When testing for the interaction between mean absolute latitude and distance in space, we  
219 permuted the sampling sites, keeping all trapping events that belonged to the same original  
220 site in the same group. When testing for the interaction between mean absolute latitude and  
221 distance in time, we first permuted the sampling sites as described above and permuted the  
222 sampling dates within each group of trapping events. When testing for the main (non-  
223 interactive) effects of the explanatory variables, we reduced the full model to exclude the  
224 respective interaction associated with the explanatory variable being tested. When testing for  
225 the main effect of distance in space or for the main effect of mean absolute latitude, we  
226 permuted the sampling sites. When testing for the main effect of distance in time, we  
227 permuted the dates within sampling locations.

228

## 229 **Results**

230 Overall the project collected 155,185 unique BINs across five geographic regions,



231 representing regional variation in latitudinal, temporal, and spatial profiles (Figures 2 & S1).  
232 BIN diversity (Figure 2) captured a large diversity of terrestrial diversity across 50 orders  
233 dominated by Diptera (51% of total BINs; 77,046 unique BINs), Hymenoptera (22% of total  
234 BINs; 33,265 unique BINs), Coleoptera (7% of total BINs; 12,550 unique BINs), Lepidoptera  
235 (7% of total BINs; 11,899 unique BINs), and Hemiptera (5% of total BINs; 7,783 unique  
236 BINs) (Figure 2).

237 Global beta-diversity (similarity) increased significantly with absolute latitude, which  
238 supports the general expectations of the LDG (Table 1; Figure S3 & S4). More generally, site  
239 comparisons between communities at higher latitudes were more similar to each other than  
240 site comparisons between communities closer to the Equator, with this observation also  
241 extending over longer temporal and spatial scales. Total beta-diversity spatio-temporal  
242 patterns were generally consistent across biogeographic regions (Table 1; Figure S4). There  
243 were some deviations in patterns across the regions, particularly when assessing the  
244 partitioned components of species replacement and richness difference (Figure S2). Oceania  
245 did not have significant total beta-diversity or species replacement association with latitude,  
246 spatial distance or temporal distance, but richness difference was found to increase  
247 significantly with increase temporal distance ( $p < 0.01$ ) (Table 1). Eurasia did not have  
248 significant latitudinal relationships with total beta-diversity or its partitioned components but  
249 did show significant positive association between total beta-diversity dissimilarity and spatial  
250 distance ( $p < 0.01$ ) along with significant positive associations between species replacement  
251 and spatial distance ( $p = 0.4$ ), temporal distance ( $p < 0.01$ ), and temporal distance x latitude  
252 ( $p = 0.05$ ), as well as significant positive associations between richness difference and temporal  
253 distance ( $p < 0.01$ ) (Table 1, Figure 3, 4). More generally, latitude, latitude x distance, or  
254 latitude x time were significantly associated with species replacement for two of five regions  
255 (North America and Eurasia) and for one of five regions for richness difference (South  
256 America). Spatial distance or space x latitude was significantly associated with species  
257 replacement for Africa and Eurasia, with richness replacement associated with North America  
258 and Africa. Temporal distance or time x latitude was significantly associated with species  
259 replacement for four of five regions (except Oceania), with richness difference being  
260 significant for North America, Oceania, and Eurasia (Table 1; Figures 3 & 4).

## 261 **Discussion**

262 The dataset generated by the Global Malaise Trap Program offers a unique opportunity to  
263 assess the underpinnings of global latitudinal biodiversity patterns using a highly diverse and

264 dominant group of terrestrial organisms. Our main finding is that total pairwise beta-diversity  
265 dissimilarity increases with decreasing latitude, increasing spatial distance, and increasing  
266 distance in time. We did not find strong indications of generalized partitioned beta-diversity  
267 patterns (i.e. species replacement or richness difference) at a global scale. However, we did  
268 find partitioned beta-diversity patterns at the regional scale which differed in the influence of  
269 latitude, spatial distance, and time indicating regional factors play a key role in overall beta  
270 diversity patterns (Figures 3, 4 & S4).

271 To evaluate community dissimilarity in space and time, we partitioned overall beta-diversity  
272 into its components: richness difference and species replacement. This partitioning framework  
273 provides insights into what factors are driving differences in biological diversity between  
274 sites, and subsequently overall diversity gradients (Borcard et al., 2018; Lennon, Koleff,  
275 Greenwood, & Gaston, 2001). We expected richness difference to increase with increasing  
276 latitude, reflecting the recent recolonization of high latitude environments following  
277 deglaciation (Dobrovolski et al., 2012). If species expanded their distributions from a joint  
278 source pool in a glacial refugium, communities established along the expansion front should  
279 form nested subsets of the source species pool. Consistent with this scenario, we generally  
280 found increased richness difference with increasing latitude (Table 1, Figure 3 & 4). We also  
281 expected species replacement to decrease with increasing latitude as a direct effect of limiting  
282 factors or eco-evolutionary processes, leading to smaller ranges and increased specialization  
283 in areas with less seasonal variability (Qian, 2009; Soininen et al., 2018). However, while we  
284 found clear support for latitudinal changes in overall beta-diversity, the underlying trends in  
285 species replacement and richness difference were inconsistent across regions, suggesting  
286 different, regional, ecological or evolutionary process are influencing biodiversity patterns  
287 (Table 1, Figures 3 & 4).

### 288 *Regional trends*

289 The appearance of generally consistent total beta-diversity patterns across geographic regions  
290 suggests alternatives to earlier observations pointing to regional rather than global factors  
291 influencing changes in community composition. However, the Oceania region seems to be an  
292 exception, as we found no association between changes in biodiversity community  
293 composition and spatial distance or latitude. This supports findings from Novotny *et al.*  
294 (2007), who proposed that regional differences in arthropod communities may actually be  
295 lower in Papua New Guinea than in temperate sites, indicating that beta-diversity is unlikely  
296 to explain the high diversity of tropical arthropod communities. However, their patterns and

297 ours may suffer from lower site coverage than is needed to confirm the lack of latitudinal  
298 gradient, but it is interesting none the less. Past studies have largely focused on temperate  
299 sites, drawing some concern that tropical sites may be underestimated in assessing LDG  
300 dynamics. The GMTP data includes 30 tropical sampling sites (23% of the sampled sites)  
301 with 986 trapping events (40.1% of the total trapping effort), providing a wider assessment.  
302 We did find partitioned components of beta-diversity differed across regions, however, which  
303 would suggest that although the general global spatial trends in beta-diversity are consistently  
304 observed the underlying environmental or biotic drivers of species replacement and richness  
305 difference may be regional rather than global (Pontarp et al., 2019). The more frequent  
306 association with spatial distance versus latitude, particularly outside North America might  
307 also support recent indications that the strength of the LDG may be greater in the western  
308 hemisphere (Kinlock et al., 2018).

#### 309 *Temporal beta-diversity trends*

310 In stark contrast to previous studies, especially in assessing global scale patterns, our analyses  
311 also considered temporal effects on latitudinal beta-diversity dynamics. Temporal species  
312 replacement or richness difference were significant for all five regions, including the global  
313 scale (Table 1). Temporal ecological dynamics are important for understanding seasonal shifts  
314 in habitat and home ranges which can influence spatial biodiversity patterns (Massol et al.,  
315 2011; Seymour et al., 2021). While pronounced temporal changes in environmental  
316 conditions (i.e. seasonality) in the temperate zone have been well documented, the same may  
317 be true for the tropics. Temporal species replacement was noted for several regions, indicating  
318 seasonal shifts in community composition likely due to competition and seasonal effects of  
319 environmental forcing (Seymour, Brown, et al., 2020; Tonkin, Bogan, Bonada, Rios-Touma,  
320 & Lytle, 2017). Here we note variation in rainfall, radiation, leaf flush etc. has been proposed  
321 to generate strong seasonality in the activity of arthropods (Basset et al., 2015; Grøtan, Lande,  
322 Chacon, & DeVries, 2014). Temporal richness replacement, which predominantly was co-  
323 associated with a significant effect of latitude likely attributed to species loss during key  
324 seasonal shifts, which may be more prominent at higher latitudes where seasonal shifts in  
325 environmental conditions are greater (Tonkin et al., 2017). The observed temporal patterns  
326 here attest to finer partitioning of community composition that should be accounted for in  
327 determining the mechanistic associations with larger spatial/latitudinal biodiversity patterns.

#### 328 *Spatial beta-diversity trends*

329 The prevalence of high beta-diversity values between sites, particularly the high number  
330 (88%) of pairwise global sites that were completely dissimilar (i.e. shared no species)  
331 highlights the extraordinary diversity of terrestrial arthropods. Communities become more  
332 similar at the regional level with 63% dissimilar sites for North America, 69% for Oceania,  
333 50% for Eurasia, 51% for Africa and 34% for South America. While most arthropods can  
334 disperse by flight, both individual home ranges and species distributions are commonly  
335 restricted (Wagner, Grames, Forister, Berenbaum, & Stopak, 2021). Range sizes have been  
336 proposed to shrink towards the tropics, following the so-called Rapoport's rule. Nonetheless,  
337 the evidence for this assertion is very limited as it has traditionally been derived from studies  
338 predominately conducted in the northern hemisphere (Gaston et al., 1998), but some recent  
339 support for southern hemisphere trends are available (Dyer, Redding, Cassey, Collen, &  
340 Blackburn, 2020; Pintor, Schwarzkopf, & Krockenberger, 2015). Our observation of greater  
341 species replacement may reflect greater niche partitioning and specialization allowed by  
342 higher productivity or stemming from greater levels of speciation (Hillebrand, 2004).

343 While previous assessments of beta-diversity have largely involved regional assessments  
344 (Seymour, Edwards, et al., 2020; Tonkin et al., 2017), which were then used to fuel meta-  
345 analyses (Graco-Roza et al., 2022; Hillebrand, 2004; Kinlock et al., 2018), this study  
346 represents a true global assessment of temporal-spatial dynamics of the most diverse lineage  
347 of terrestrial animals. The consistency in general global patterns which were decomposed at  
348 the regional scale enables a mechanistic assessment of the planetary biodiversity patterns.  
349 This synthesis was only made possible by our coupling of a standardized sampling method  
350 with DNA-based taxonomic assignments (Hebert et al., 2003). Importantly, convincing  
351 analyses of beta diversity require an efficient means for rigorously establishing the incidence  
352 of species shared across sites in massive sampling programs. Such methods are finally  
353 available for our use at a planetary scale.

354

355

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531 **Table 1. Statistical significance of patterns detected.** To assess the importance of each  
532 candidate variable (listed on the left as Covariates), we used a series of permutation tests. We  
533 first calculated the log-likelihood ratio between the model where the explanatory variable  
534 being tested was included (the full model), and the model with the explanatory variable being  
535 tested was excluded (the reduced model). We then compared the observed log-likelihood ratio  
536 to its null distribution, which we computed by permuting the data N=1,000 times (see  
537 Methods for the exact permutation schemes implemented). This table shows the proportion of  
538 permutation outcomes for which the log-likelihood ratio of the model fitted to the actual data  
539 was lower than the log-likelihood ratio for the models fitted to the permuted data. Values at or  
540 below 0.05 are deemed significant and are indicated in bold face.

541

Covariate	Region	Beta	Richness Difference	Species Replacement
Latitude	Global	<b>&lt;0.01</b>	0.19	0.18
Distance in space ( $\Delta S$ )	Global	<b>&lt;0.01</b>	0.95	0.91
Distance in time ( $\Delta T$ )	Global	<b>&lt;0.01</b>	1.00	1.00
$\Delta T \times$ Latitude	Global	<b>&lt;0.01</b>	<b>0.04</b>	<b>0.05</b>
$\Delta S \times$ Latitude	Global	<b>&lt;0.01</b>	<b>0.03</b>	<b>0.03</b>
Latitude	North America	<b>&lt;0.01</b>	<b>0.02</b>	<b>0.01</b>
Distance in space ( $\Delta S$ )	North America	<b>&lt;0.01</b>	<b>0.01</b>	0.45
Distance in time ( $\Delta T$ )	North America	1.00	<b>&lt;0.01</b>	<b>&lt;0.01</b>
$\Delta T \times$ Latitude	North America	<b>&lt;0.01</b>	0.07	<b>&lt;0.01</b>
$\Delta S \times$ Latitude	North America	<b>&lt;0.01</b>	<b>0.01</b>	0.63
Latitude	South America	<b>&lt;0.01</b>	<b>&lt;0.01</b>	0.17
Distance in space ( $\Delta S$ )	South America	0.17	0.51	0.66
Distance in time ( $\Delta T$ )	South America	<b>&lt;0.01</b>	1.00	<b>&lt;0.01</b>
$\Delta T \times$ Latitude	South America	0.33	0.49	0.33
$\Delta S \times$ Latitude	South America	0.85	0.15	0.19
Latitude	Oceania	0.71	0.90	0.81
Distance in space ( $\Delta S$ )	Oceania	0.57	0.98	0.11
Distance in time ( $\Delta T$ )	Oceania	1.00	<b>&lt;0.01</b>	1.00
$\Delta T \times$ Latitude	Oceania	0.33	0.37	0.45
$\Delta S \times$ Latitude	Oceania	0.43	0.08	0.07
Latitude	Eurasia	0.13	0.29	0.37
Distance in space ( $\Delta S$ )	Eurasia	<b>&lt;0.01</b>	0.93	<b>0.04</b>
Distance in time ( $\Delta T$ )	Eurasia	1.00	<b>&lt;0.01</b>	<b>&lt;0.01</b>
$\Delta T \times$ Latitude	Eurasia	0.29	0.06	<b>0.05</b>
$\Delta S \times$ Latitude	Eurasia	0.43	0.71	0.85
Latitude	Africa	<b>0.03</b>	0.83	0.40
Distance in space ( $\Delta S$ )	Africa	<b>&lt;0.01</b>	<b>0.01</b>	<b>0.01</b>
Distance in time ( $\Delta T$ )	Africa	1.00	1.00	<b>&lt;0.01</b>
$\Delta T \times$ Latitude	Africa	0.52	0.60	0.88
$\Delta S \times$ Latitude	Africa	0.91	0.54	0.22

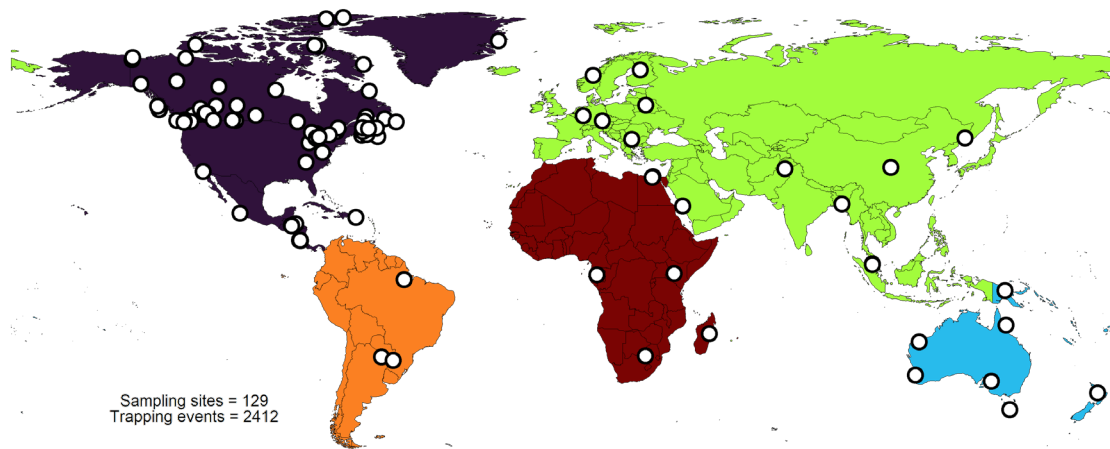
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544 Figures

545 **Figure 1. Sampling sites and five biogeographical regions considered.** Regions are  
546 differentiated by color. White points indicate sampling locations.

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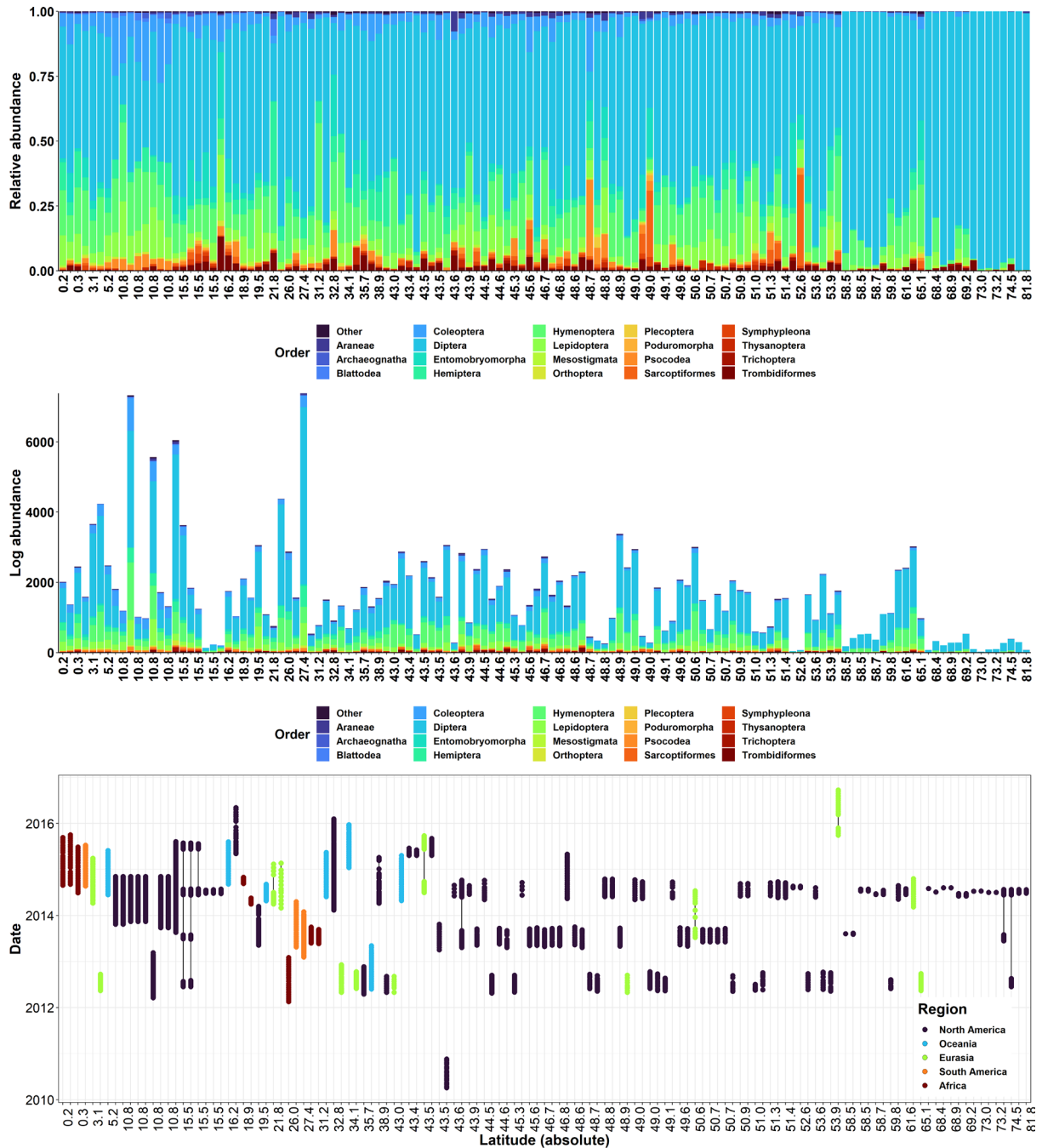
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552 **Figure 2. Diversity of BINs (i.e. species proxies) across the sampling sites.** top and middle  
 553 panels: (each vertical bar). Colors shown are unique terrestrial arthropod orders with each  
 554 height corresponding to the relative abundance (top) or log abundance (middle) of unique  
 555 BINs across the given site. The absolute latitude is provided on the x-axis with sites arranged  
 556 from low to high. Bottom panel, each circle represents a sampling unique event, with colors  
 557 corresponding to the region. Each horizontal line shows a unique sampling location with  
 558 latitude (absolute) indicated on the y-axis. Colors correspond to regional groups following the  
 559 same color scheme as Figure 1.

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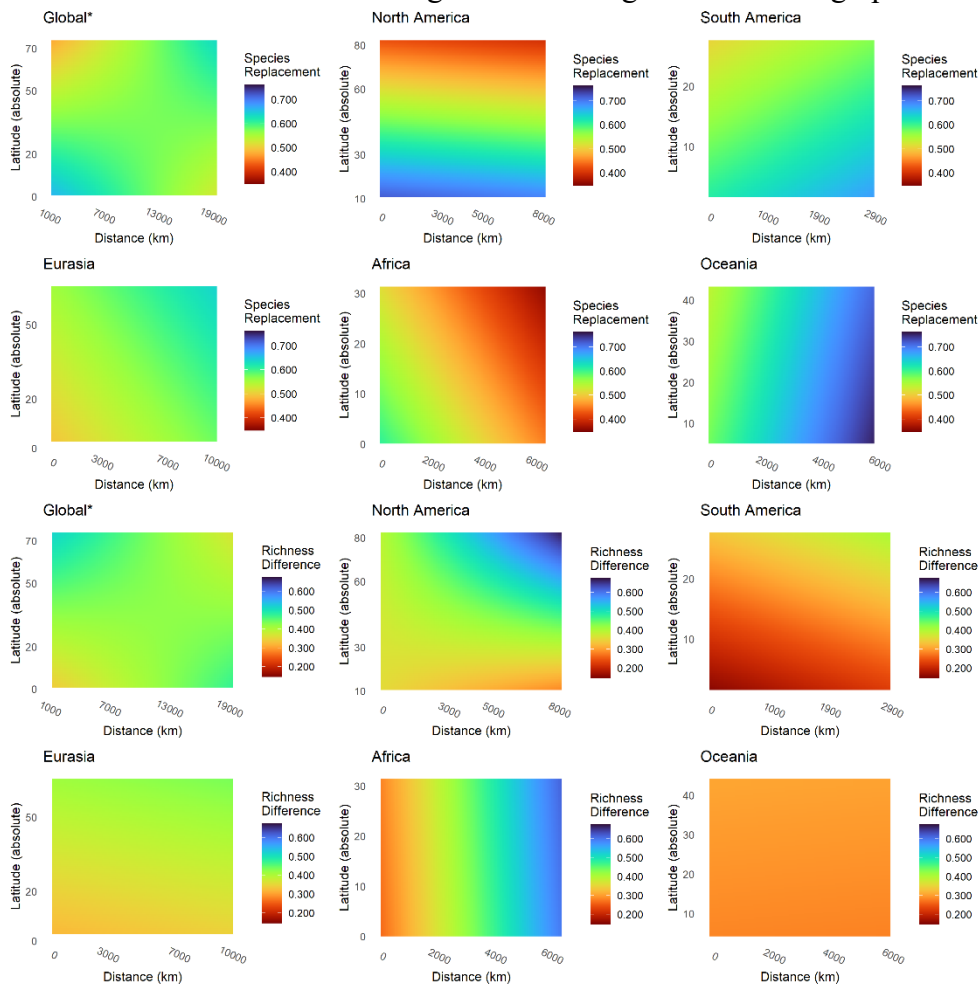


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563 **Figure 3. Impacts of pairwise distance in space and latitude on community dissimilarity.**

564 This figure shows the species replacement and richness difference components of beta-  
565 diversity, plotted from the fitted values of a linear model of the metric in question as a  
566 function of distance in space, distance in time (here set to zero), mean absolute latitude, and  
567 the interactions mean latitude  $\times$  distance in space and mean latitude  $\times$  distance in time (see  
568 Table 1 for statistical significances). In this figure, we explicitly test whether spatial patterns  
569 of community beta-diversity *in space* (in terms of overall beta-diversity, species replacement  
570 or richness difference) varies detectably with latitude. Regions with a significant interaction  
571 between pairwise difference in latitude and pairwise distance are indicated by an asterisk.  
572 Note the differences in the scaling of axes among the individual graphs.



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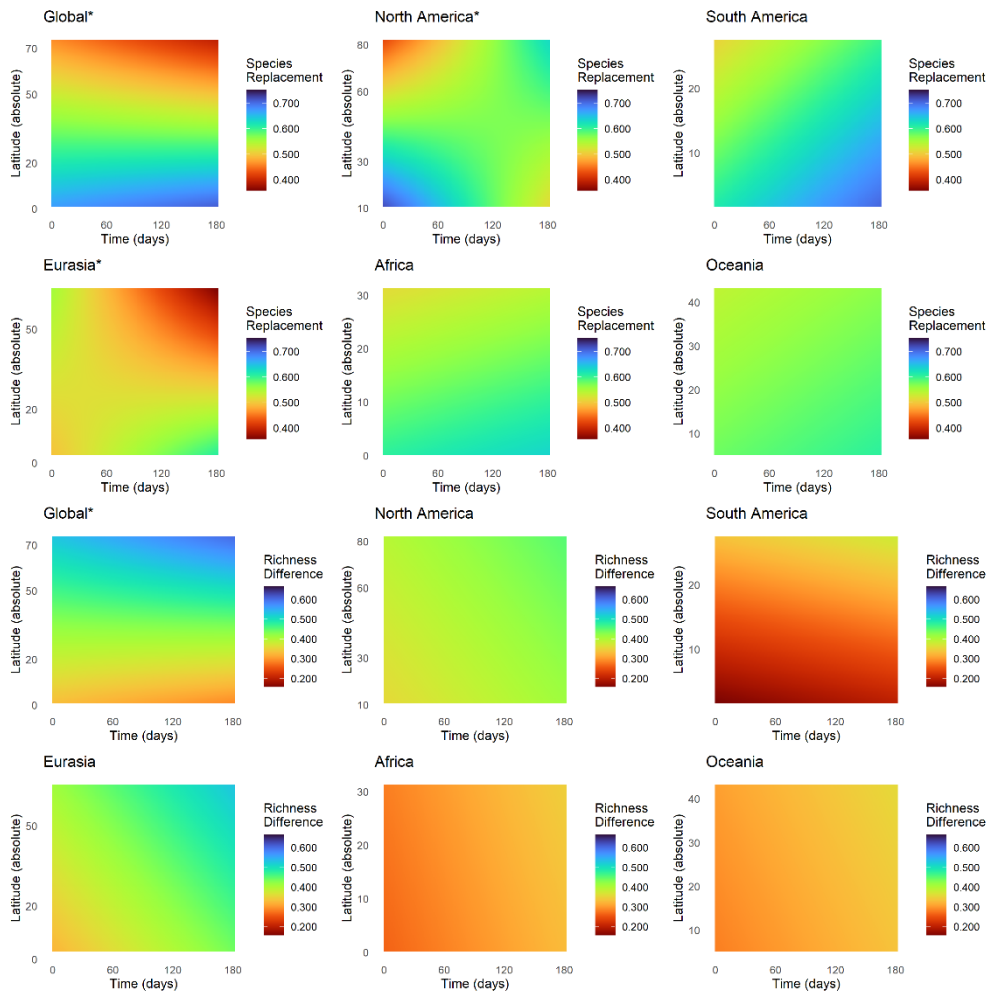
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577 **Figure 4. Impacts of pairwise distance in time and latitude on community dissimilarity.**

578 This figure shows the species replacement and richness difference components of beta-  
579 diversity, plotted from the fitted values of a linear model of the metric in question as a  
580 function of distance in space (here set to zero), distance in time, mean latitude, and  
581 interactions between mean latitude  $\times$  distance in space and mean latitude  $\times$  distance in time  
582 (see Table 1 for statistical significances). In this figure, we explicitly test whether temporal  
583 patterns of community beta-diversity *in time* (in terms of overall beta-diversity, species  
584 replacement and richness difference) varies detectably with latitude.

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