

# The impact of medication reviews conducted in primary care on hospital admissions and mortality - an observational follow up of a randomized controlled trial

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## Research article

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# Abstract

**Background** Drug-related problems among the elderly population are common and increasing. Multi-professional medication reviews (MRs) have arisen as a method to optimize drug therapy for frail elderly patients. Research has not yet been able to show conclusive evidence of the effect of MRs on mortality or hospital admissions. **Aim** The aim of this study was to assess the impact of MRs' on hospital admissions and mortality after six and 12 months in a frail population of 369 patients in primary care in a randomized controlled study. **Methods** Patients were blindly randomized to an intervention group (receiving MRs) and a control group (receiving usual care). Descriptive data on mortality and hospital admissions at six and 12 months were collected. Survival analysis was performed for time to death and time to the first hospital admission within 12 months. **Results** Of the total number of 369 included patients, 182 were randomized to the intervention group and 187 to the control group. Most of the patients (75%) were females and lived in nursing homes. At six months, 50 patients of the baseline population (27%) in the control group had been admitted to hospital at least once, compared to 40 patients (21%) in the intervention group. At 12 months, the percentage had increased to 70 (37%) in the control group compared to 53 (29%) in the intervention group. Compared to usual care, we found that MRs reduced the risk of hospital admissions within 12 months by 42% (HR = 0.58, 95% CI 0.37-0.92), but found no difference on mortality (HR = 1.12, 95% CI 0.78-1.61) between the groups. **Conclusion** We suggest that MRs should be recommended in the care of frail elderly patients with expected benefits on delayed hospital admissions.

## 1. Background

Drug-related problems (DRPs) and polypharmacy in the elderly population are common and increasing (1). Lack of a consistent drug list and low physician continuity are some of the reasons for drug-related problems and the need for emergency hospital contacts (1). Most of these problems can be avoided and multi-professional medication reviews have arisen as a method to optimize drug therapy for frail elderly patients (2, 3). Meanwhile, medication reviews (MR) have been questioned for their efficacy (4). Even if studies have shown that pharmacists recommendations are clinically relevant (5, 6), recent research has not been able to demonstrate an effect on mortality or hospital admissions, due to a heterogeneity of intervention designs, patient characteristics as well as follow up time in different studies; hence the conclusion is that more evidence is needed (7, 8). Furthermore, a systematic review and meta-analysis of randomized controlled trials with short follow-up time (<3 months) showed no effect on clinical outcomes or quality of life and the researchers argued that it should be considered to stop performing medication reviews as standard care (9). This conclusion is controversial as medication reviews are well-established in Sweden and Swedish authorities grant incentives and other funding on an annual basis to healthcare providers in order to optimize drug therapy among elderly patients (10). It is important to mention that the heterogeneity of study designs and outcomes is a problem while performing meta-analyses.

The model for MRs in primary care in southern Sweden with focus on elderly patients with multiple illnesses is well-studied and has been shown to reduce potentially inappropriate medications (PIMs) and consequences due to medication errors (11). The integrated approach with a pharmacist helping in the clinical routine had initially been developed in hospital care but has been adapted to primary care. MRs using this approach have now been conducted for the last 15 years, both in nursing homes and community-dwelling elderly patients in Sweden, with the goal to improve quality of medication use and patient safety.

This randomized controlled trial (RCT), which was performed in 2012, included 369 patients and showed that multi-professional medication reviews conducted by teams including nurses, pharmacists and GPs had a lowering effect on the number of drugs as well as on the number of PIMs among elderly patients (12). The first results showed that two months after the medication reviews, both the number of patients with at least one PIM and the number of intervention patients having 10 or more drugs had decreased in the intervention group, while there were no statistically significant differences in the control patients. DRPs were identified in 93% of the 182 patients in the intervention group and 16% of the DRPs were related to PIMs. The two most common intervention recommendations the pharmacist presented to the physician were withdrawal of drug therapy (30%) and reduced dosage (28%) (12). A significant proportion (87%) of the studied population had treatment with fall risk increasing drugs (FRIDs) (12).

As we could show statistically significant benefits in the intervention group receiving MRs in the short follow-up time, we considered it important to study this further by focusing on mortality data and hospital admissions during a longer follow-up time of six and 12 months.

## 2. Aim

The aim of this study was to assess the impact of multi-professional MRs on hard outcomes, such as hospital admissions and mortality, at six and 12 months in elderly patients in primary care.

## 3. Methods

### *Study setting and design*

Primary care in Skåne county, southern Sweden is provided by approximately 150 primary health care centers (PHCCs), both public and private. All public PHCCs were invited to participate in the study. Four pharmacists with extensive prior experience of performing MRs were recruited to the project and assigned to four different areas in Skåne county. Patients eligible for inclusion were users of the multi-dose drug dispensing system aged 75 years or older, living in nursing homes or community dwelling with municipally provided home care. Patients were included in the study after they provided written consent directly or through relatives in cases of severe cognitive impairment. A total number of 391 were assessed for eligibility and 22 were excluded for following reasons: not meeting inclusion criteria (n=5), declined to participate (n=1), deceased before data collection at base line (n=5) or other reasons (n=11).

The study design is observational with a follow up of the randomized controlled trial with 369 included patients, 182 in the intervention group and 187 in the control group.

### *Intervention*

The pharmacists randomized patients blindly to the intervention or the control group using non-transparent closed envelopes and with a random number generator, stratified only for geographic area (12).

PIMs were identified following the national guidelines of the Swedish Board of Health and Welfare on drug therapy in the elderly (15).

The pharmacists' recommendations were documented in patients' EMRs and communicated to the physician in different ways, depending on the PHCC's routines and organization and consisted of team rounds, written contact, personal contact and telephone contact. The physician decided about medication changes and this was documented in the patient's EMR.

### *Data collection*

We followed the intervention and control groups from the initial baseline population of 369 patients for 12 months. Primary outcome variables were differences in survival time and time to the first hospital admission between the groups at six and 12 months. Secondary outcomes were difference in number of deaths and difference in number of hospital admissions between the groups at six and 12 months. Data from the EMR and the regional patient database were collected using Swedish personal numbers thus assuring 100% data coverage.

### *Data analysis*

Descriptive data for baseline characteristics as well as mortality and hospital admissions were performed. Kaplan-Meier analysis was used to estimate survival function with time to the first hospital admission and time to death as outcome events. Non-parametric Kaplan-Meier was chosen as the exact survival time and time to the first hospital admission for each patient were known. Because the MRs continued as a part of the routine care when yearly renewal of medication was needed, some of the patients in the control group received a MR during the follow up period. We defined this event as a major protocol deviation and chose to perform a per protocol analysis. The patients in the control group receiving MRs during the follow-up period were censored in the survival analysis. Furthermore, we used the log-rank test for equality across strata (i.e. control and intervention group). The proportionality assumption was tested ( $p$ -value < 0.05), which is the main assumption of the Cox proportional hazard model to determine the relationship between covariate and survival time or time to hospital admission and to obtain hazard ratio (HR).

Data were analysed using SPSS version 25 (16).

## 4. Results

Of the 369 included patients, 182 were randomized to the intervention group and 187 to the control group (Figure 1). Most of the patients (75%) were females and lived in nursing homes. Patient characteristics of the population at baseline are described in Table 1, with no significant observed differences in age, sex, living form and number of continuous drugs between the intervention and control group. At six months, 33 patients in the control group had received at least one MR. At 12 months, another 22 patients (meaning a total number of 55 patients) in the control group had received at least one MR. Patients were censored from the survival analysis at the time for the first MR during the follow up time of 6 and 12 months.

There was no difference in number of deaths between the groups at six and 12 months (Table 2). The survival analysis showed no differences in mortality between groups at 12 months ( $p=0.551$ ) with a HR = 1.12, (95% CI 0.78-1.61).

At six months, 50 patients of the baseline population (27%) in the control group had been admitted to hospital at least once, compared to 40 patients (21%) in the intervention group. At 12 months the percentage had increased to 70 (37%) in the control group compared to 53 (29%) in the intervention group. There was a significant difference ( $p=0.038$ ) between mean number of hospital admissions at 12 months in the intervention group and the control group (Table 3). A comparison between number of hospital admissions between groups at 12 months is presented in Figure 2. The mean survival time (time to death and time to the first hospital admission) in the groups is presented in Table 4. A survival analysis with Log Rank (Mantel Cox) showed a significant decrease ( $p=0.012$ ) in risk for the first hospital admission at 12 months between the intervention and control group (Figure 2). Patients were censored in the analysis at the exact time for event for death in the intervention group and for death or when receiving a MR in the control group, depending on which event occurred first. A Cox regression analysis was performed in order to estimate the risk for hospital admissions within 12 months, adjusted for sex, age, number of drugs and number of MRs, with patients censored at the first hospital admission. The Cox regression analysis showed that MRs reduced the risk of hospital admissions by 42% (HR = 0.58, 95% CI 0.37-0.92) compared to usual care ( $p=0.021$ ). The proportionality assumption was fulfilled.

## 5. Discussion

This study found that the patients in the intervention group had a lower risk for first hospital admission within 12 months after the MR compared to the patients in the control group. We found no difference in mortality rate between the groups. The study design with per protocol analysis instead of intention to treat analysis might have contributed to biased survival estimates, as patients in the control group receiving MR during the follow up time were excluded from the analysis. However, MRs in the control group were performed both as a part of the routine care and due to concerns about medication use that may have been related to condition deterioration. We believe therefore that the risk for bias is lower than expected.

The studied population was elderly individuals that had multimorbidity and polypharmacy. The intervention and the control group were similar concerning age, living form and number of drugs at baseline.

Adverse drug reactions account for a substantial amount of all hospital admissions (17-19), and many of these are preventable (18, 19). Previously published data from this RCT (12) showed no difference in number of drugs or PIMs between the groups at baseline but a significantly lower number of drugs and PIMs in the intervention group at two months follow-up, which indicates that MRs might reduce the occurrence of drug-related problems (DRPs). If DRPs can be prevented by reducing the number of drugs and PIMs through MRs, some hospital admissions could be avoided or delayed, and thereby could save both burden and costs for the individual patient as well as for the society.

Many studies of MRs are based on a patient population already admitted to the hospital and have shown that MRs in a hospital setting will extend the time to readmission (20-22). A strength of this study was that the MRs were performed in primary care, therefore adding this perspective to the body of evidence. We estimate the effect of the MR on direct patient outcomes such as hospital admissions, which is reliable data thanks to Swedish personal numbers.

Another strength of this study is the long follow-up time (12 months), which is rare in this kind of study (9), due to the high mortality rate among this frail group of elderly patients. The two most common intervention recommendations the pharmacist presented to the physician were withdrawal of drug therapy and reduced dosage. Many PIMs, as long-acting benzodiazepines and tramadol need a longer withdrawal period with an initial stepwise lowered dosage. The long follow-up period in the present study allowed a clinical effect of the medication withdrawal; this is a strength of this study. However, a major limitation is that additional data about medications was not collected after the initial follow up at two months from base line, to confirm the assumption of maintained or increased withdrawal of PIMs.

Censoring the individuals in the control group, which had received a MR, when the survival analysis was performed is also a strength of the study.

A major limitation of the study is the lack of information about patients' base line comorbidities, the reason for hospital admissions or if the hospital admission was due to a DRP. Some of the hospital admissions could be assumed to be related to DRPs as the previous paper on the same population showed a high incidence of reported falls (29%) and severe falls leading to hospital admissions (17%) prior to baseline (12). Thus, further analysis on the possible association between DRPs and hospital admissions needs to be performed.

We have not analyzed data on morbidity such as number of chronic conditions and/or cognitive impairment, and this is a limitation of the study. Given the similarities in the number of drugs both between groups and between patients in different living forms, we can only assume that there were similar levels of multimorbidity in the groups.

We found a positive effect of MRs on the risk for hospital admissions. As it is not solely about avoiding hospital admission, the impact of MRs on quality of life would be interesting to assess as only a few studies have raised this important question (7, 23).

Our results showed that MRs had no effect on mortality or on time to death when the intervention group and control group were compared, which is in-line with other research (4, 24).

It has been argued that MRs are part of the physicians' professional responsibilities. However, the multi-professional approach with pharmacist and nurse involvement might bring different dimensions upon patients' medication list, giving additional attention to frail elderly in an over-burdened primary care.

About 90,000 individuals over the age of 65 live in nursing homes in Sweden. This number represents less than 5% of the population aged 65 years and older (25) and hence is the part of the elderly population needing most daily care. We suggest that frail elderly individuals should be treated with increased attention, including by conducting MRs, in order to avoid hospital admissions.

## 6. Conclusion

The results show that MRs can have a significant impact on the risk for hospital admissions but not on mortality rate. MRs should be considered in the care of frail elderly patients with expected benefits on avoiding hospital admissions.

## Abbreviations

DRP	Drug Related Problem
EMR	Electronic Medical Record
FRIDs	Fall Risk Increasing Drugs
GP	General Practitioner
MR	Medication Review
PHCC	Primary Health Care Center

PIM Potentially Inappropriate Medication

RCT Randomized Controlled Trial

## Declarations

### Ethics and consent to participate

The study has been approved by the Regional Ethical Review Board in Lund (Reg. no. 2018/8). Patients were included in the study after they provided written consent directly or through relatives in cases of severe cognitive impairment. The study adheres to CONSORT guidelines.

### Consent to publish

We have obtained informed and written consent to publish from all study participants, as required by the Regional Ethical regulations in Sweden.

### Availability of data and materials

The datasets used during the current study are available from the corresponding author on reasonable request.

### Competing Interest

The authors declare that they have no competing interests.

### Funding

This study was funded by Lions Forskningsfond Skåne.

### Authors contributions

VMN, CL and BBB participated in the study design and coordination. VMN and CL participated in data collection. VMN and BBB participated in statistical analysis. All authors drafted, reviewed and approved the final manuscript.

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## Tables

**Table 1.** Population characteristics at baseline.

All patients n = 369	<b>Baseline characteristics</b>	<b>Intervention (n=182)</b>	<b>Control (n=187)</b>	<b>p-value</b>
	Female sex, n (%)	138 (75.8)	142 (75.9)	0.980
	Institutionalized, n (%)	139 (76.4)	140 (74.9)	0.809
	Age, mean (SD)	87.0 (5.8)	87.7 (5.5)	0.662*
	No of continuous drugs (min-max)	9.2 (1-20)	9.7 (3-25)	0.528*

\* Chi<sup>2</sup> test

\*\* Student's t test.

**Table 2.** Patient characteristics and mortality numbers at six and twelve months, when patients in the control group receiving MRs were excluded.

At 6 m (33 excluded) n = 336	<b>Patient characteristics</b>	<b>Intervention (n=182)</b>	<b>Control (n=154)</b>	<b>p-value</b>
	Female sex, n (%)	138 (75.8)	115 (74.4)	0.899*
	Deceased, n (%)	28 (15.4)	34 (22.0)	0.472*
	Institutionalized, n (%)	138 (76.4)	115 (71.4)	0.319*
	Age, mean	87.0	87.9	0.303**
At 12 m (55 excluded) n = 314	<b>Base line characteristics</b>	<b>Intervention (n=182)</b>	<b>Control (n=132)</b>	<b>p-value</b>
	Female sex, n (%)	133 (75.1)	103 (75.2)	1.000*
	Deceased, n (%)	60 (32)	60 (45.4)	0.771*
	Institutionalized, n (%)	137 (76.3)	98 (71.5)	0.364*
	Age (mean)	87.0	87.8	0.226**

\* Chi<sup>2</sup> test

\*\* Student's t-test.

**Table 3.** Group comparison in hospital admissions for all patients and when patients from control group receiving a MR are excluded at 6 and 12 months.\*Chi<sup>2</sup> test.

Population	Hospital admissions at 6 months, mean (min, max)			Hospital admissions at 12 months, mean (min, max)		
	Intervention (n=182)	Control (n=187)	p-value	Intervention (n=182)	Control (n=187)	p-value
All patients n = 369	0.29 (0-4)	0.37 (0-5)	0.609	0.45 (0-4)	0.63 (0-7)	0.077
Within 6 m (33 excluded) n = 336	0.29 (0-4)	0.40 (0-5)	0.505			
Within 12 m (55 excluded) n = 314				0.46 (0-4)	0.69 (0-7)	<b>0.038</b>

**Table 3.** Group comparison in hospital admissions for all patients and when patients from control group receiving a MR are excluded at 6 and 12 months.\*Chi<sup>2</sup> test.

Population	Hospital admissions at 6 months, mean (min, max)			Hospital admissions at 12 months, mean (min, max)		
	Intervention (n=182)	Control (n=187)	p-value	Intervention (n=182)	Control (n=187)	p-value
All patients n = 369	0.29 (0-4)	0.37 (0-5)	0.609	0.45 (0-4)	0.63 (0-7)	0.077
Within 6 m (33 excluded) n = 336	0.29 (0-4)	0.40 (0-5)	0.505			
Within 12 m (55 excluded) n = 314				0.46 (0-4)	0.69 (0-7)	<b>0.038</b>

Table 4. Mean survival time in the control and intervention groups

Intervention/Control	N	Minimum	Maximum	Mean	Std. Deviation	
Control	Days to first hospital admission	187	0	361	160.61	69.7
	Days to death	187	5	363	175.23	58.1
	No of hospital admissions within 6 months	187	0	5	0.37	0.7
	No of hospital admissions within 12 months	187	0	7	0.63	1.1
	Valid N	187				
Intervention	Days to first hospital admission	182	4	356	167.04	54.4
	Days to death	182	2	341	175.40	56.1
	No of hospital admissions within 6 months	182	0	4	0.29	0.6
	No of hospital admissions within 12 months	182	0	4	0.45	0.8
	Valid N	182				

## Figures

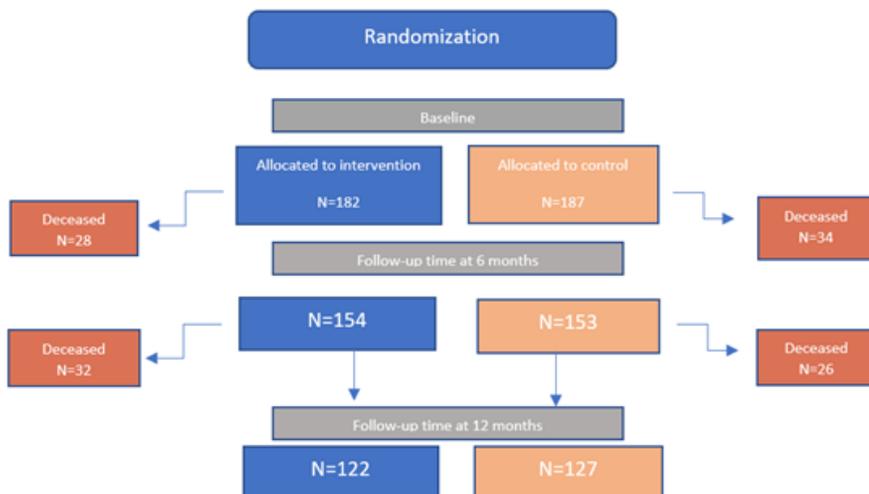
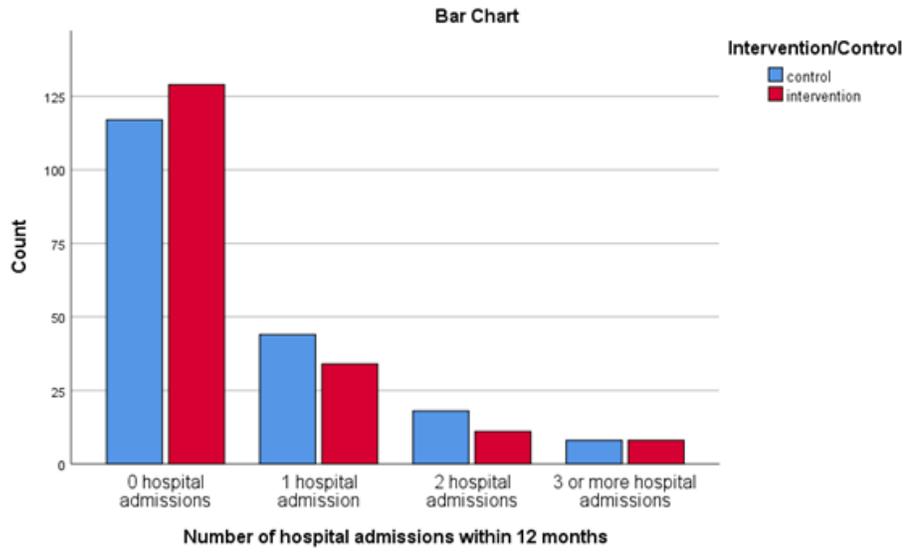


Figure 1

Flow chart for the studied population at baseline, six and 12 months.



**Figure 2**

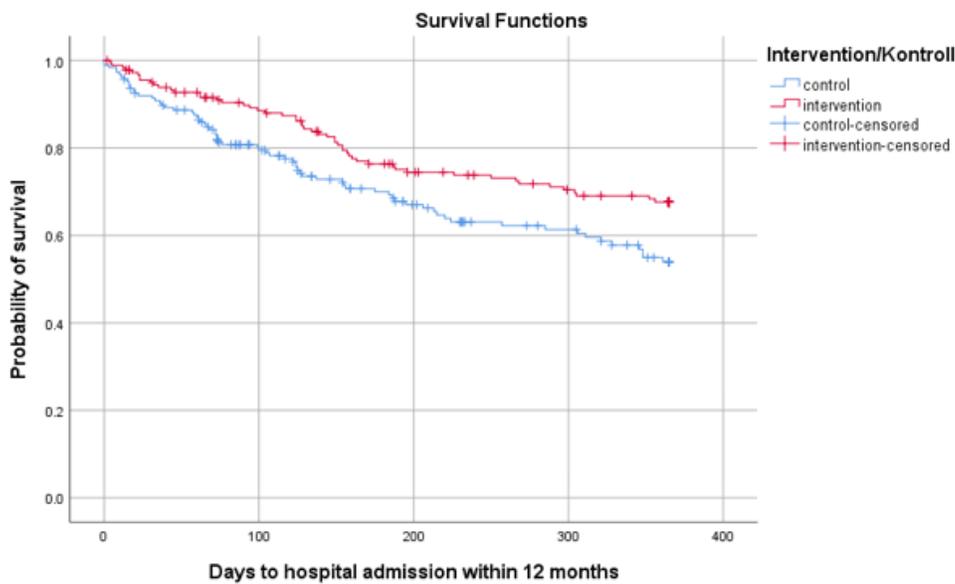
Comparison in number of hospital admissions within 12 months in the control group.

Intervention/Control	Total N	N of Events	Censored	
			N	Percent
Control	187	70	117	62.6%
Intervention	182	53	129	70.9%
Overall	369	123	246	66.7%

### Overall Comparisons

	Chi-Square	df	Sig.
Log Rank (Mantel-Cox)	6.316	1	.012

Test of equality of survival distributions for the different levels of Intervention/Control.



**Figure 3**

Output for Kaplan Meier estimates for days to first hospital admission in intervention and control group at 12 months and Log Rank test.

## Supplementary Files

This is a list of supplementary files associated with this preprint. Click to download.

- [CONSORT2010ChecklistMR.doc](#)