

Environmental Regulations, Competition and Divestment of Foreign Direct Investment: Evidence from China

Lufeng Tai

Shanghai University

Linnan Yan (✉ yanlinnan123@163.com)

Shanghai University <https://orcid.org/0000-0001-8219-343X>

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Environmental Regulations, Competition and Divestment of Foreign Direct Investment: Evidence from China

Lufeng Tai^a, Linnan Yan^{a,*}

a. School of Economics, Shanghai University, 200444 Shanghai, China.

tailufeng123@163.com (L.T.), yanlinnan123@163.com (L.Y.)

*Correspondence: yanlinnan123@163.com; Tel.: +86 18800379686

Abstract: This paper is concerned with whether the Chinese increasingly stringent environmental regulations (ER) lead to the divestment of foreign direct investment (FDI). Based on industrial firm-level panel data from 2003 to 2010, our results show that the stricter ER do not induce the divestment of FDI but rather reduce the probability of foreign firms' withdrawal from China. Moreover, in cities with a higher degree of marketization, the ER have greatly reduced the exit probability of foreign firms. The mechanism analysis shows that due to the scale and technological advantages of foreign companies, the ER have stimulated innovation, increased the market share and profits of foreign companies. However, the ER have a greater negative impact on domestic firms' performance. This research has theoretical and empirical significance for the economic development and environmental protection of developing countries.

Keywords: Divestment of FDI; firm exit; environmental regulations; pollution haven hypothesis; Porter hypothesis

JEL classification: F21; O13

1. Introduction

The pollution haven hypothesis (PHH) suggests that environmental protection policies inevitably increase enterprises' costs, and countries with lower environmental costs are more attractive to multinational enterprises (Copeland and Taylor, 2004). However, the researches on whether the looser ER attracted more FDI have not reached a consistent

30 conclusion (Cai et al., 2016; Millimet and Roy, 2016; Muhammad and Khan, 2019; Kim and
31 Rhee, 2019; Santos and Forte, 2020; Dong et al., 2021). Kellenberg (2009) states that “the
32 empirical validity of pollution haven effects continues to be one of the most contentious
33 issues in the debate regarding international trade, foreign investment, and the environ-
34 ment”.

35 As the largest developing country in the world, China provides an excellent example
36 for testing the PHH. Since the 21st century, the Chinese government has continuously
37 enhanced environmental protection policies. Besides, the largest number of FDI indus-
38 trial firms in China was 77,847 in 2008, and the latest data was 43,588 in 2019. This
39 number has decreased by 44% compared to 2008. Some scholars worry about that the
40 environmental control policies result in the divestment of FDI in China (Greaney et al.,
41 2017). However, most of the existing studies testing the PHH are conducted from the
42 perspective of foreign capital inflows; whether the stricter ER lead to foreign firms’ exit
43 behavior is unknown.

44 Therefore, our work uses Chinese industrial firm-level data to detect whether the ER
45 induce foreign firms’ withdrawal in China. The contributions to the existing literature
46 in this paper are as follows.

- 47 ● First, most of the existing literature have studied the impact of the ER on the influx
48 of FDI from the macro-level (Cai et al., 2016), and yet we use firm-level data to
49 investigate the PHH from the aspect of the divestment of FDI. A firm’s exit behav-
50 ior is regarded as a decision that relies on balancing benefits and costs in this paper.
51 Our results show that the ER significantly reduce the probability of foreign firms’
52 exit behavior, which provides a new enlightenment for understanding the relation-
53 ship between ER and FDI.
- 54 ● Second, we shed light on the moderating effect of market competition. We find that
55 the market competition strengthens the role of the ER in reducing the probability
56 of foreign capital withdrawal. In a particular region with a higher degree of mar-
57 ketization, foreign firms can gain a better performance. The market competition
58 mechanism reveals the decision-making process of enterprises.

59 The rest of the paper is organized as follows: Section 2 provides a literature review

60 of the studies on PHH and develops hypotheses according to the theory and fact. Sec-
61 tion 3 establishes the estimation methodology and provides an explanation of variables.
62 Section 4 is the empirical test results of the hypotheses. Section 5 is the discussion of
63 the results, and section 6 is the conclusion.

64

65 **2. Literature review and hypothesis**

66 **2.1 Literature review**

67

68 The topic of PHH is the relationship between ER and FDI. The early theoretical literature pre-
69 dicts that international capital will move to countries with lower ER (Chichilnisky, 1994; Motta
70 and Thisse, 1994). But the early empirical studies fail to find conclusive evidence of PHH.
71 Some researchers' conclusions support PHH (Becker and Henderson, 2000; Henderson, 1996;
72 Keller and Levinson, 2002; Chung, 2014). However, there are also some studies that do not
73 support PHH (Eskeland and Harrison, 2003; Javorcik and Wei, 2003; Levinson, 1996). There
74 might be two reasons for the inconsistent empirical conclusions. On the one hand, the samples
75 selected in various studies are different, and PHH might not apply to all industries or regions
76 (Eskeland and Harrison, 2003). On the other hand, the potential endogeneity makes it difficult
77 to accurately reveal the causal relationship between ER and FDI (Levinson and Taylor, 2008).

78 To explain why the empirical researches do not support PHH, theoretical researchers study
79 the problem from more aspects, such as regional heterogeneity (Zeng and Zhao, 2009; Sanna-
80 Randaccio and Sestini, 2012) and firm heterogeneity (Dijkstra et al., 2011; Elliott and Zhou,
81 2013). Zeng and Zhao (2009) introduce agglomeration economies into a two-state model with
82 high transportation costs and imperfect competition. They show that agglomeration economies'
83 benefits might be greater than the reduction of pollution control costs due to relocation.

84 More recently, empirical researches began to deal with the unobserved heterogeneity and
85 potential endogeneity. However, there is also no consistent conclusion (Cai et al., 2016; Milli-
86 met and Roy, 2016; Dong et al., 2021; Muhammad and Khan, 2019; Kim and Rhee, 2019). Cai
87 et al. (2016) use the implementation of the TCZ (implemented by the Chinese government to
88 control SO₂ and acid rain) policy as a quasi-natural experiment. They conclude that pollution
89 control reduces the inflow of FDI significantly. Using Chinese FDI outflow data during 2008-

90 2018 and applying DID estimation, Dong et al. (2021) find that host countries with more re-
91 laxed ER attract FDI from polluting industries. However, Muhammad and Khan (2019) and
92 Kim and Rhee (2019), both using country-level panel data, suggest that stricter ER promote the
93 development of new industries, increase productivity, and attract FDI.

94 There are also some studies that have reached complex conclusions (Yang et al., 2019; Yu
95 and Li, 2020). Yang et al. (2019) used the spatial Durbin model to test the influence of Chinese
96 urban ER on the inflow of FDI. Their results implicate that the validity of the PHH has a sig-
97 nificant regional difference; the level of ER in the eastern (central and western) region of China
98 is positively (negatively) correlated with the FDI introduction. Using Chinese province-level
99 panel data from 2009 to 2018 to explore the PHH, Yu and Li (2020) find that ER do not inhibit
100 the inflow of FDI but improve the quality of FDI.

101 There are many theoretical and empirical studies on PHH but remain some limitations. First,
102 most researches are carried out from the nation and industry levels, but the inflow and outflow
103 of FDI is actually a kind of corporate behavior. Existing empirical investigations ignore the
104 decision-making behavior of enterprises. Second, the inflow and outflow of FDI are two sides
105 of the same coin. Most of the researches focus on the relationship between ER and inflow of
106 foreign capital, while few on the withdrawal of foreign firms. Therefore, for the sake of envi-
107 ronmental protection and stable development of foreign investment, it is of great significance
108 to study whether ER lead to foreign firms' withdrawal. Our work has improved the two limita-
109 tions and has enriched the literature on PHH. We use microdata to study the PHH from the
110 perspective of firm exit behavior, and we also consider endogeneity.

111

112 **2.2 Hypothesis development**

113

114 Many studies suggest that the poor performance and the low profits of overseas subsidiaries are
115 the most fundamental reasons for divestment (Jagersma and Van Gorp, 2003; Hryckiewicz and
116 Kowalewski, 2011). Therefore, whether the stricter ER lead to the divestment of FDI depends
117 on the degree of ER policies' impact on their business performance.

118 Porter hypothesis suggests that ER can force firms to engage in innovation, and the compen-
119 sation effect of innovation may eventually exceed the costs of pollution control (Porter and

120 vander Linde, 1995). Firms may reduce their investment due to ER's additional expenses in the
121 short term, but in the long run, technological progress can decrease the costs (Jaffe et al., 2002).
122 Many studies supporting Porter hypothesis show that ER positively impact firm performance
123 (Dasgupta et al., 2006; Ramanathan et al., 2010; Kneller and Manderson, 2012). Lee (2020)
124 uses Chinese listed company data to study, and he find that both command and control ER and
125 voluntary instruments positively affect the firm financial performance.

126 Besides, the costs increase caused by the stricter ER might vary from company to company.
127 Dijkstra et al. (2011), using a Cournot duopoly three-stage model, show that ER could increase
128 all companies' costs, including domestic companies. But if a foreign firm has more efficient
129 technology than a domestic firm, the ER could give the foreign firm a greater cost advantage.

130 According to Schumpete's theory, large companies have more competitive and innovative
131 advantages than small companies. Figure 1 is the kernel density estimation of domestic and
132 foreign industrial firms' gross output and assets in China. The graphs indicate that the scale of
133 foreign firms is larger than that of domestic firms. Most Chinese foreign firms come from de-
134 veloped countries, and their technological level is relatively higher than the domestic firms.
135 The foreign firms are insensitive to China's ER (Dean et al., 2009). Therefore, the foreign firms
136 are less affected by the stricter ER due to their large scale, advanced technology, and more
137 competitiveness, which have no reasons to leave China. Thus, we develop the following hy-
138 pothesis:

139 H1: Stricter ER can reduce the probability of foreign firms leaving China, i.e. ER have a
140 negative effect on firms' exit behavior.

141

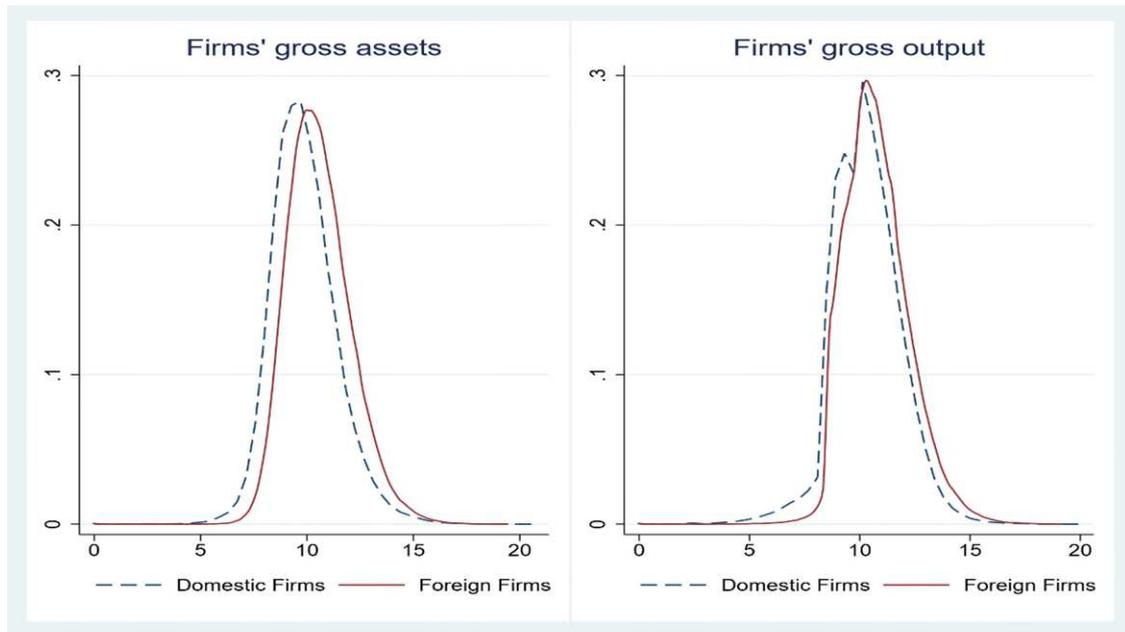


Fig 1 The kernel density estimation of the scale of foreign firms and domestic firms.

Another issue sheds light on the role of market competition in regulating the relationship between ER and FDI. Porter hypothesis suggests that firms can gain a competitive advantage through first-mover advantage or differentiation strategy by adopting ER (Porter, 1991). Sanna-Randaccio and Sestini (2012) take advantage of an international duopoly model to show that stricter regulation does not induce firms to relocate in a country with larger market size. Foreign firms could obtain greater competitive advantages due to their scale and technological advantages. However, China began to implement market-oriented reforms in 1992, and the degree of marketization varies widely in different regions. ‘To support the SOEs’ (state-owned enterprises) role, government intervention has ensured the allocation of resources to them, and positioned SOEs at the higher tiers of the value chain’ (Lee, 2020). If the market is highly controlled and monopolized by SOEs, the foreign firms cannot use their competitive advantages and might be squeezed out of the market.

On the contrary, if a particular region has a higher degree of marketization, even if the ER become stricter, the foreign companies can use their scale and technological advantages to obtain competitive advantages and exhibit better financial performance. Javeed et al. (2020) show that the high degree of competition has a significant and positive moderating effect on the relationship between ER and firm performance. Therefore, we develop the following hypothesis:

H2: There is a moderating effect of competition on the relationship between ER and foreign

163 firms' exit behavior, i.e. in cities that with a higher degree of marketization, ER have a more
164 negative impact on firms' exit behavior.

165 **3. Methodology and variables**

166 **3.1 Methodology**

167
168
169 To investigate the impact of ER on the foreign firms' exit, considering the potential problem of
170 missing variables, enterprise heterogeneity and data availability, we develop the benchmark
171 model as follows:

$$172 \text{EXIT}_{it} = \alpha_0 + \alpha_1 \text{ER}_{ct} + \alpha_2 X_{it} + \alpha_3 Z_{ct} + \gamma_c + \lambda_j + \theta_t + \varepsilon_{it} \quad (1)$$

173
174
175 Where i , c , j and t represent the firm, city, industry and year, respectively. EXIT_{it} is the de-
176 pendent variable denotes the exit behavior of the firm i , and ER_{ct} is the independent variable
177 which denotes the ER. X_{it} and Z_{ct} represent the firm-level and city-level control variables.
178 γ_c , λ_j , θ_t represent city, industry and time effects to control unobserved factors and heter-
179 ogeneity. ε_{it} is the random error.

180 The dependent variable is the exit behavior of the enterprise; it is a dummy variable which
181 equals to 1 if firm i exits in year t , and 0 otherwise. Following Greenaway et al. (2008) and
182 Eslava et al. (2013), we mainly use the probit model to estimate Eq (1). The probit model is
183 non-linear, and the linear model is also unbiased in estimating this problem; therefore, the linear
184 probability model (LPM) is used as a robustness test. From our framework, we expect that ER
185 will reduce the exit probability of foreign companies, that is, we expect $\alpha_1 < 0$.

186 Second, in order to verify the moderating role of market competition in Hypothesis 2, we
187 introduce the market competition variable (and the interaction term with ER) into Eq (1). The
188 following model is set:

$$189 \text{EXIT}_{it} = \beta_0 + \beta_1 \text{ER}_{ct} + \beta_2 \text{MC}_{ct} + \beta_3 \text{ER}_{ct} \cdot \text{MC}_{ct} + \beta_4 X_{it} + \beta_5 Z_{ct} + \gamma_c + \lambda_j + \theta_t +$$

190
191 $\varepsilon_{it} \quad (2)$

192
193 MC_{ct} represents market competition which is the moderating variable, $\text{ER}_{ct} \cdot \text{MC}_{ct}$ is the

194 interaction term of ER and market competition, and other settings are the same as Eq (1). We
195 estimate Eq (2) using the same method as Eq (1). The hypothesis 2 implicates that in cities with
196 a high degree of market competition, ER greatly reduce the probability of foreign firms exit;
197 i.e. we expect $\beta_3 < 0$.

198 In addition, to explore the mechanism of ER on firm exit, we further study the impact of ER
199 on corporate performance. Model 3 is set as follows:

200

$$201 \quad Per_{it} = \gamma_0 + \gamma_1 ER_{et} + \gamma_2 X_{it} + \gamma_3 Z_{ct} + \zeta_i + \theta_t + \varepsilon_{it} \quad (3)$$

202

203 Where Per_{it} is the performance variable of firm i in period t , we use the fixed effects model
204 to estimate Eq (3). ζ_i is firm fixed effects which controls those non-time variables of firm level,
205 and other variables are the same as Eq (1). The potential reason why ER will not cause foreign
206 companies to leave China is that ER improve the performance of foreign companies, so we
207 expect $\gamma_1 > 0$.

208

209 **3.2 DATA**

210

211 In this paper, we draw on both city-level and firm-level data. City-level data are 2003-2010
212 yearly data from the Chinese City Statistical Yearbook, which includes 290 Chinese cities. This
213 data contains indicators of the city's economic development and environmental governance.
214 The firm-level data are sourced from the Chinese Industrial Enterprise Database (CIED), which
215 is developed and maintained by the Chinese National Bureau of Statistics. This data records
216 basic corporate information and various financial indicators from 2003 to 2010, spanning 37
217 industries in the two-digit manufacturing industry (for more details, please see Cai and Liu,
218 2009).

219

220 **3.3 Variables description**

221

222 Chinese foreign firms are mainly divided into two types: one is the exclusively foreign-owned
223 enterprise, and the other is the equity joint ventures. The equity joint ventures in China refer to
224 an enterprise whose foreign capital accounts for more than 25% of its registered capital. The

225 foreign firms in this article include these two types. Our variables and their explanations are as
226 follows:

227 (1) Dependent variable (EXIT): Following Weintraub et al. (2008), $EXIT_{it}$ is a dummy var-
228 iable, which equals to 1 if firm i exits in year t , and 0 otherwise. If a firm appears in year t but
229 does not appear in year $t+1$, it means that the firm has exited from the market in year t . If a firm
230 appeared in the market in both years t and $t+1$, it means that the firm did not withdraw from the
231 market in year t . Besides, we delete samples of companies that re-enter after exiting.

232 (2) Independent variable (ER_{ct}): Many studies measure ER in annual fee on waste discharge
233 (Smarzynska, 2001; Fredriksson, 2002; Lee, 2020), but it does not indicate the ultimate pollu-
234 tion control effect. Based on data availability, we collected the sulfur dioxide removal rate,
235 industrial smoke removal rate, and industrial wastewater purification rate from 2003 to 2010 at
236 the city level in China. Following to Liu and Lin (2018), we calculate the weighted index of
237 three pollution reduction indicators to replace the degree of ER. The calculation method is as
238 follows:

239 Since different pollutant indicators cannot be directly compared, the first step is to standard-
240 ize these indicators:

$$241 \quad US_{ij}^S = (UE_{ij} - \min UE_j) / (\max UE_j - \min UE_j)$$

242 Where UE_{ij} represents the emission reduction indicator of j pollutants in city i , $\max UE_j$
243 and $\min UE_j$ are the maximum and minimum values of pollutant j in all cities, and US_{ij}^S rep-
244 resents the standardized reduction indicator of j pollutant in city i .

245 Each city has a different level of pollutant emissions, and the proportion of various pollutants
246 in the city is also different. It is necessary to calculate the weighted coefficient for each city.
247 The calculation method of the weighted sum coefficient S_{ij} is as follows:

$$248 \quad S_{ij} = (q_{ij} / \sum_i q_{ij}) / (Ind_i / \sum_i Ind_i)$$

249 where q_{ij} represents the emission of pollutant j in city i , and Ind_i represents the gross
250 industrial output of city i .

251 Finally, the index ER_{ct} of the city i in year t is calculated as follows:

$$252 \quad ER = \sum_j S_{ij} \times US_{ij}^S$$

253 ER_{ct} is the city-level panel data from 2003 to 2010. This index represents the pollution re-
254 duction performance and the intensity of ER in different cities.

255 (3) Moderating variables (MC): We use two indicators to measure market competition. The
 256 first one (MC1) is the marketization index. It indicates that the degree of liberalization of the
 257 regional market. The second one is the proportion of the gross industrial output of non-SOEs
 258 to the gross industrial output of the region (MC2). The government directly supervises and
 259 regulates SOEs (Zhang et al, 2017), therefore the larger the indicator, the higher the degree of
 260 marketization.

261 (4) Instrumental variables: Given the potential endogeneity, we adopt the instrumental vari-
 262 able approach to solve this problem. Following Hering et al. (2014), we choose the ventilation
 263 coefficient (VC) as the instrumental variable of ER. In the standard Box model of atmospheric
 264 pollution, the ventilation coefficient is considered of the determinant of air pollution's disper-
 265 sion speed. It determines the height at which pollutants disperse in the atmosphere (Jacobsen,
 266 2002). This data comes from the European Centre for Medium-Term Weather Forecasting
 267 (ECMWF) ERA-Interim data set (for more details, please see Hering et al., 2014).

268 Furthermore, we also use city standard coal consumption (SC) data in 1985 as another in-
 269 strumental variable for ER. Because the ventilation coefficient and the consumption of standard
 270 coal in 1985 belong to natural and historical factors. They are related to pollution emissions but
 271 have little to do with current firm behavior. Therefore, we think that the two instrumental vari-
 272 ables are exogenous and associated with the independent variables.

273 (5) Control variables: We add characteristic variables at the city level and the firm level as
 274 the control variables, such as regional GDP per capita (PGDP), regional salary per capita (CS),
 275 firm salary per capita (FS). Table 1 shows the explanations of all variables.

276

277 **Table 1. Descriptive of variables**

Variables	Definition	Observations	Min	Max	Mean
EXIT	Equal to 1 if firm exits in year t, and 0 otherwise.	293,499	0	1	0.122
ER	Weighted sum indicators of three pollutant emission reduction data	293,499	0.145	8.023	1.436
PGDP	Logarithm of regional actual GDP per capita	293,499	3.286	7.820	6.035
PIND	Logarithm of regional actual industrial output per capita	293,499	1.254	8.700	5.457
FDI	The proportion of the gross industrial output of foreign firms of the region	293,499	0.0007	0.9400	0.453
CS	Logarithm of regional actual salary per capita	293,499	2.283	11.828	10.124

FS	Logarithm of firm actual salary per capita	293,499	1.130	4.767	2.798
FAGE	Logarithm of firm age	293,499	0	3.135	1.799
FLOW	The ratio of firm prime operating revenue to circulating assets	293,499	0.281	29.95	3.367
FSIZE	Logarithm of firm gross assets	293,499	7.636	14.38	10.38
MC1	Logarithm of marketization	279,616	0.9555	2.4681	2.224
MC2	The proportion of the gross industrial output of non-SOEs to the region	293,499	0.2841	0.9635	0.7397
VC	Logarithm of regional ventilation coefficient	293,499	6.560	8.211	7.471
SC	Logarithm of city standard coal consumption	159,676	0	7.191	3.960

278

279

4. Empirical results

280

4.1 Causality between ER and foreign firms' exit behavior

281

282

283 Based on Eq (1), we estimate the impact of ER on the foreign firms' exit behavior. Table 2
 284 presents the results of all regressions. Column (1) reports the result without control variables.
 285 We can find that the ER's coefficient is significantly negative at 1% level. After controlling the
 286 city, industry, and year fixed effects and other variables in column (2), ER's coefficient is still
 287 significantly negative. Column (3) is the average marginal effect of the probit model. The result
 288 indicates the intensity of ER increases by 1%, and the probability of withdrawal of foreign
 289 enterprises decreases by 0.01%. As expected, the strengthening of ER reduce the exit probabili-
 290 ty of foreign firms rather than inducing foreign firms withdrawal. Our hypothesis 1 is con-
 291 firmed.

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To test the results' robustness, we conduct the three robustness tests: (1) Considering the financial crisis. The 2008 financial crisis caused a large number of companies to withdraw from the market. In order to eliminate the impact of this factor, we only use the data from 2003 to 2007 for estimation. (2) Considering the error of statistical threshold. The CIED includes all SOEs and non-SOEs whose annual business income is greater than 5 million Chinese Yuan (CNY). If a company's sales fall below 5 million yuan, it may only withdraw from the database because of the shrinking of business income, but not really withdraw from the market. To remove this error, we use samples with incomes more than 20 million Yuan for estimation. (3) Using linear probability model (LPM). When the dependent variable is 0-1, OLS estimation is still unbiased, although it has heteroscedasticity. Therefore, this paper uses the LPM to explore

302 the robustness of the results.

303 Columns (4)-(6) are the results of three robustness tests respectively. The coefficients of ER
304 are all still significantly negative, which are consistent with the results of the basic regression.

305 In addition, we have also investigated whether there is a U-shaped effect between ER and for-
306 eign capital withdrawal, but the results show that there is no such effect.

307

308 **Table 2. Regression of ER's Impact on foreign firms' exit**

Variables	(1)	(2)	(3)	(4)	(5)	(6)
	No control variables	Baseline model	Marginal ef- fect	Year before 2008	Sales > 20 million yuan	LPM
ER	-0.0563*** (0.0202)	-0.0547** (0.0269)	-0.0100** (0.0049)	-0.0897*** (0.0266)	-0.0934*** (0.0331)	-0.0089* (0.0048)
PGDP		0.0604 (0.0907)	0.0110 (0.0165)	0.0419 (0.0923)	-0.1116 (0.1330)	0.0054 (0.0182)
FSIZE		-0.2247*** (0.0106)	-0.0418*** (0.0019)	-0.1932*** (0.0125)	-0.1005*** (0.0086)	-0.0383*** (0.0017)
PIND		0.0527 (0.0504)	0.0096 (0.0092)	0.0765 (0.0469)	0.0722 (0.0677)	0.0121 (0.0107)
FS		-0.0841*** (0.0161)	-0.0410*** (0.0029)	0.0009 (0.0155)	0.0320*** (0.0121)	-0.0206*** (0.0035)
CS		0.1618 (0.1035)	0.0295 (0.0189)	0.1945 (0.1200)	0.4586** (0.1873)	0.0245 (0.0180)
FAGE		0.0877*** (0.0111)	0.0160*** (0.0020)	0.0831*** (0.0114)	0.0577*** (0.0121)	0.0157*** (0.0019)
Flow		-0.0121*** (0.0024)	-0.0022*** (0.0004)	-0.0123*** (0.0027)	0.0079*** (0.0014)	-0.0022*** (0.0004)
FDI		-0.6571** (0.3233)	-0.1198** (0.0588)	-0.7601* (0.4224)	-0.3946 (0.3955)	-0.1252** (0.0588)
Obs	293 423	293 423	293 423	240,260	194,013	293 499
City FE	YES	YES	YES	YES	YES	YES
Industry FE	YES	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES	YES

309 Robust standard errors are in parentheses; asterisks ***, **, and * denote significance at the 1%, 5%, and 10% levels, respectively.

310

311 4.2 Endogeneity

312

313 We also pay attention to the endogenous problems such as missing variables and bidirectional
314 causality, which may lead to the bias of results. To address the problem, we introduce the ven-
315 tilation coefficient as an instrumental variable in the first column of Table 3 and estimate Eq (1)

316 by the two-stage least square method. The result implicates that the coefficient of ER becomes
 317 significantly negative at 1% level and larger. And the KPF value is greater than 10, which indi-
 318 cates that the instrumental variable has passed the weak identification test.

319 In the third column, we use both ventilation coefficient and city standard coal consumption
 320 in 1985 as the instrumental variables. The coefficient of ER remains significantly negative at
 321 1% level. We also conducted weak identification test (KPF) and overidentification test (Hansen
 322 J) on 2SLS. It is shown that there is no weak identification problem and the instrumental vari-
 323 ables are exogenous. In summary, after controlling endogeneity, the impact of ER on foreign
 324 firms' exit is in line with the basic regression's result.

325

326 **Table 3. IV Regression of ER's Impact on foreign firms' exit**

Variables	(1) Exit	(2) First-stage ER	(3) Exit	(4) First-stage ER
ERS	-0.1965*** (0.0277)		-0.1124*** (0.0219)	
VC		-0.4865*** (0.0196)		-0.8042*** (0.0326)
SC				0.0085*** (0.0023)
PGDP	-0.1542*** (0.0208)	-0.7754*** (0.0118)	-0.1453*** (0.0234)	-1.1119*** (0.0167)
FSIZE	-0.0722*** (0.0019)	0.0074*** (0.0023)	-0.0662*** (0.0025)	0.0054 (0.0036)
PIND	0.1205*** (0.0145)	0.5010*** (0.0067)	0.1013*** (0.0136)	0.5945*** (0.0093)
FS	-0.0330*** (0.0015)	0.0120*** (0.0018)	-0.0363*** (0.0019)	0.0105*** (0.0028)
CS	-0.0599*** (0.0135)	-0.4560*** (0.0240)	-0.0182* (0.0111)	-0.4630*** (0.0300)
FAGE	0.0360*** (0.0020)	0.0066* (0.0029)	0.0332*** (0.0025)	0.0040 (0.0044)
Flow	-0.0022*** (0.0003)	-0.0005 (0.0003)	-0.0016*** (0.0003)	-0.0013** (0.0005)
FDI	-0.0710*** (0.0187)	-0.2350*** (0.0221)	0.1022*** (0.0262)	0.1357*** (0.0409)
Obs	279,594	279,594	151,882	151,882
Firm FE	YES	YES	YES	YES
Year FE	YES	YES	YES	YES
Kleibergen-Paaprk Wald F		614.226		306.378

Robust standard errors are in parentheses; asterisks ***, **, and * denote significance at the 1%, 5%, and 10% levels, respectively.

4.3 Test for the moderating effect of market competition

Table 4 presents the results of Eq (2), which evaluates the moderating effect of market competition between ER and foreign enterprises' exit behavior. We first use the marketization index (MC1) as a proxy variable of market competition. To ensure our findings' validity, we then introduce MC2 as another proxy variable of market competition.

The coefficients of ER*MC1 and ER*MC2 are -0.1977 and -0.1627 respectively in column (1) and column (3), and both are significantly negative at the 1% level. The results indicate that market competition has a significantly moderating effect on the relationship between ER and foreign firms' exit behavior. In other words, ER are more likely to inhibit foreign firms' leaving China, in cities with a higher degree of marketization. Our results are consistent with Javeed et al. (2020). As we mentioned above, in the fierce market competition environment, when facing the negative impact, such as stricter ER, foreign firms can better use their technological and scale advantages to improve performance. Thus, the probability of foreign firms leaving China is reduced. Besides, we apply the LPM to estimate moderating effects of market competition, and the results remain robust.

Table 4. Test for moderating effect of market competition

Variables	(1)	(2)	(3)	(4)
	Market Competition 1		Market Competition 2	
	Probit	LPM	Probit	LPM
ER	-0.0661*** (0.0184)	-0.0120*** (0.0039)	-0.0608*** (0.0210)	-0.0101** (0.0041)
MC1	-0.3541 (0.6457)	-0.0758 (0.1174)		
ER*MC1	-0.1977*** (0.0521)	-0.0367*** (0.0118)		
MC2			0.4224 (0.2661)	-0.0005 (0.0132)
ER*MC2			-0.1627*** (0.0604)	-0.0081** (0.0038)
PGDP	0.1230 (0.1191)	0.0179 (0.0237)	0.0304 (0.0944)	0.0004 (0.0186)
FSIZE	-0.2093***	-0.0348***	-0.2247***	-0.0383***

	(0.0120)	(0.0021)	(0.0106)	(0.0017)
PIND	0.0403	0.0057	0.0796	0.0150
	(0.0593)	(0.0115)	(0.0546)	(0.0118)
FS	-0.0028	-0.0013	-0.0853***	-0.0208***
	(0.0109)	(0.0021)	(0.0161)	(0.0035)
CS	0.2591	0.0400	0.1362	0.0199
	(0.1628)	(0.0273)	(0.1024)	(0.0181)
FAGE	0.0854***	0.0148***	0.0879***	0.0158***
	(0.0112)	(0.0019)	(0.0110)	(0.0019)
Flow	-0.0109***	-0.0018***	-0.0122***	-0.0022***
	(0.0025)	(0.0004)	(0.0025)	(0.0004)
FDI	-0.3569	-0.0436	-0.9816***	-0.1782***
	(0.3771)	(0.0744)	(0.3805)	(0.0672)
Obs	279,544	279,616	293,423	293,499
City FE	YES	YES	YES	YES
Industry FE	YES	YES	YES	YES
Year FE	YES	YES	YES	YES

Robust standard errors are in parentheses; asterisks ***, **, and * denote significance at the 1%, 5%, and 10% levels, respectively.

347

348

349 **4.4 Heterogeneity test**

350

351 As ER have a great impact on high-pollution industries, the effect of ER in reducing foreign
 352 enterprises' withdrawal should be different. Besides, there is a noticeable gap in Chinese re-
 353 gional economic development and policy implementation. To examine the difference, we take
 354 on the industrial and regional heterogeneity test in Table 5, respectively.

355 First, according to the geographical location, we divide the companies into two types: the
 356 eastern regions and the non-eastern regions. As shown in the first and second columns, the
 357 impact of ER on the foreign firms' exit in the east of China is significantly negative, which is
 358 in line with our expectation. While the impact is not significant in the non-eastern areas. Dean
 359 et al. (2009) suggest that the Chinese non-eastern regions' economic activities are dominated
 360 by industries related to natural resources, which may be more sensitive to ER policies. The
 361 improvement of ER cannot lead to technological innovation and productivity improvement of
 362 enterprises in non-eastern regions. In addition, the degree of marketization in Chinese non-
 363 eastern regions is lower; therefore, the role of market competition is not prominent.

364 Second, we divide samples into three industrial types on basis of the degree of pollution

365 emissions: heavy pollution, medium pollution and light pollution, and then do grouping regres-
366 sion. The results are shown in columns 3-5. We find that all the regression coefficients of ER
367 are negative, but only the medium pollution industries are significant. On the one hand, the
368 costs of pollution control in the heavy pollution industries are higher, and ER have a more
369 negative impact on serious pollution firms. Therefore, more stringent ER cannot reduce the exit
370 probability of enterprises in heavy pollution industries. On the other hand, ER have less impact
371 on the costs and performance of companies in the cleaning industries, therefore the coefficient
372 is not significant. The results are consistent with Bu and Wagner (2016).

373

374

Table 5. Heterogeneity test

Variables	(1)	(2)	(3)	(4)	(5)
	Eastern	non-Eastern	Heavy pollution	Medium pollution	Light pollution
ER	-0.0634** (0.0287)	0.0221 (0.0362)	-0.0322 (0.0351)	-0.0835** (0.0327)	-0.0442 (0.0305)
PGDP	0.0411 (0.0975)	-0.1992 (0.3220)	-0.0001 (0.1203)	0.0034 (0.0964)	0.1432 (0.1046)
FSIZE	-0.2327*** (0.0105)	-0.1318*** (0.0145)	-0.1986*** (0.0111)	-0.2385*** (0.0111)	-0.2353*** (0.0120)
PIND	0.0453 (0.0543)	0.9670*** (0.2636)	0.0588 (0.0593)	0.0968 (0.0656)	0.0240 (0.0562)
FS	-0.0799*** (0.0170)	-0.1378*** (0.0233)	-0.1305*** (0.0195)	-0.0828*** (0.0155)	-0.0618*** (0.0191)
CS	0.1505 (0.1071)	-0.2969 (0.4124)	0.5737*** (0.1696)	0.0091 (0.0810)	0.1704 (0.1190)
FAGE	0.0874*** (0.0119)	0.1004*** (0.0176)	0.1416*** (0.0145)	0.1062*** (0.0137)	0.0520*** (0.0118)
Flow	-0.0133*** (0.0027)	-0.0016 (0.0027)	-0.0089*** (0.0025)	-0.0097*** (0.0031)	-0.0155*** (0.0027)
FDI	-0.6451* (0.3443)	-0.0837 (0.5971)	-0.2402 (0.5227)	-0.8055** (0.3571)	-0.7851** (0.3172)
Obs	274,559	18,864	71,937	90,687	130,499
City FE	YES	YES	YES	YES	YES
Industry FE	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES

375

Robust standard errors are in parentheses; asterisks ***, **, and * denote significance at the 1%, 5%, and 10% levels, respectively.

376

377 4.5 Mechanism analysis

378

379 The literature indicates that ER can improve corporate performance (Porter, 1991; Javeed,

380 2020). The high degree of competition forces firms to be innovative and differentiated with
 381 high-quality products to gain maximum profit (Ammann et al. 2013; Byun et al. 2012). The
 382 foreign firms have no reason to leave China if their performance is improved. Therefore, per-
 383 formance can be viewed as an intermediate channel for the relationship between ER and firms'
 384 exit behavior.

385 In order to examine the Eq (3), we use three ways to measure Per_{it} . The first one is firm
 386 market share, which is the proportion of the firms' gross output to industry's gross output. The
 387 second one is firm profit margin, which is firm financial performance. The third is firm total
 388 factor productivity (TFP). We use the LP method to estimate the firm TFP, representing the
 389 enterprise's technological level.

390 The underlying assumption of this article is that the performance of foreign enterprises
 391 caused by ER is better than that of domestic enterprises. Therefore, we also analyze the impact
 392 of ER on domestic companies. As shown in Table 6, ER have a significant negative impact on
 393 market share and profits for the domestic firms but have no significant effect on TFP. However,
 394 ER have a significant positive impact on the three corporate performance indicators of foreign
 395 firms. These results have confirmed our hypothesis, on the one hand, the stricter ER stimulate
 396 innovation and improve the productivity of foreign firms. On the other hand, ER reduce the
 397 market share of domestic enterprises, and then make foreign enterprises obtain larger market
 398 share and higher profit margin through competition. Therefore, ER reduce the outflow of for-
 399 eign capital.

400

401 **Table 6. Mechanism analysis**

Variables	Domestic firms			Foreign firms		
	(1) Market Share	(2) Profit	(3) TFP	(1) Market Share	(2) Profit	(3) TFP
ER	-0.0021* (0.0012)	-0.0003* (0.0002)	-0.0025 (0.0024)	0.0139*** (0.0039)	0.0048*** (0.0005)	0.0382*** (0.0048)
FSIZE	0.4762*** (0.0023)	0.0110*** (0.0003)	0.2727*** (0.0037)	0.6400*** (0.0055)	0.0220*** (0.0006)	0.4490*** (0.0059)
FS	0.2304*** (0.0015)	0.0018*** (0.0002)	0.1926*** (0.0028)	0.1582*** (0.0032)	0.0057*** (0.0004)	0.1757*** (0.0042)
FAGE	0.0320*** (0.0022)	-0.0000 (0.0003)	0.0344*** (0.0040)	0.0661*** (0.0057)	0.0017*** (0.0007)	0.0558*** (0.0064)
FDI	-0.0647***	-0.0045***	-0.2435***	-0.0693	-0.0676***	-0.2910***

	(0.0057)	(0.0006)	(0.0103)	(0.0484)	(0.0053)	(0.0551)
Flow	0.0255***	-0.0003***	0.0214***	0.0644***	0.0017***	0.0549***
	(0.0002)	(0.0000)	(0.0002)	(0.0008)	(0.0001)	(0.0008)
PGDP	0.1771***	-0.0161***	0.3476***	0.0778***	0.0102***	0.0896***
	(0.0092)	(0.0012)	(0.0176)	(0.0139)	(0.0018)	(0.0177)
PIND	0.2205***	0.0177***	0.1368***	0.0228***	-0.0052***	0.0739***
	(0.0096)	(0.0012)	(0.0161)	(0.0077)	(0.0013)	(0.0100)
CS	-0.0753***	-0.0107***	-0.0225	-0.0387***	0.0024	0.1367***
	(0.0094)	(0.0013)	(0.0205)	(0.0134)	(0.0017)	(0.0177)
Obs	1,176,327	1,195,584	682,325	283,205	293,499	293,499
Firm FE	YES	YES	YES	YES	YES	YES
Year FE	YES	YES	YES	YES	YES	YES

402 Robust standard errors are in parentheses; asterisks ***, **, and * denote significance at the 1%, 5%, and 10% levels, respectively.

403

404 5. Discussion

405

406 The empirical results confirm our hypotheses about the impact of the ER on foreign firms, and
407 provid new evidence for PHH as well.

408 Our hypothesis 1 shows that the ER reduce the probability of foreign firms leaving China.
409 The results are consistent with Dijkstra et al. (2011). Many studies have shown that the ER
410 reduce the influx of FDI (Cai et al., 2016). However, our evidence indicates that the inflow of
411 FDI is different from its outflow. Multinational enterprises need to pay sunk costs and additional
412 switching costs to withdraw from the host country's market. The ER prompt foreign firms with
413 scale and technological advantages to gain greater cost and competitive advantages. There is
414 no reason for multinational enterprises to withdraw from China.

415 The second hypothesis implicates that the ER have a more significant effect on reducing
416 foreign firms' exit from regions with higher marketization degree. As long as the market is large
417 enough, the ER do not cause companies to relocate (Sanna-Randaccio and Sestini 2012). If a
418 region is controlled by the government and monopolized by the SOEs, the foreign firms cannot
419 use their competitiveness to improve performance. The free market competition environment
420 can alleviate the contradiction between environmental protection and FDI. This result is in line
421 with Javeed et al. (2020).

422 In addition, we document the mechanism of the results above. Our results reveal that the ER
423 improve the performance of foreign firms but reduce the performance of domestic ones. Most

424 of the foreign firms in China come from developed countries. They are more prominent in scale
425 and more advanced in technology. As a result, the foreign firms are less negatively affected by
426 the ER (Dijkstra et al. 2011), and the competitiveness of foreign enterprises becomes greater
427 compared with Chinese companies. That is why the ER reduce the probability of divestment of
428 FDI.

429

430

6. Conclusion

431

432 The purpose of this study is to investigate whether the ER lead to foreign firms leaving China.
433 By using firm-level data and considering endogeneity, the results propose that the ER have
434 reduced the probability of divestment of FDI in China. In addition, we introduce the moderating
435 effect of market competition. The results indicate that the ER's effect in reducing foreign firms'
436 exit probability is more significant in places with a higher marketization degree. The reason is
437 that the ER have improved the performance of foreign firms. The results implicate that the ER
438 and economic development are not opposed to each other, which provides a new evidence for
439 PHH and Porter hypothesis.

440 Based on this article, it is necessary to continue to carry out reasonable environmental pro-
441 tection policies to achieve the purpose of protecting the environment and attracting FDI. The
442 government should provide an orderly and good market competition environment, and let en-
443 terprises play their competitive advantage. At the same time, domestic enterprises in developing
444 countries should also improve their competitive advantages to meet the challenges brought by
445 the strengthening of ER.

446 It is worth noting that this research may suffer from some limitations, which need further
447 investigation. First, the companies in developing countries are less competitive than multina-
448 tional enterprises from developed countries. Although the environmental protection policies of
449 developing countries will not lead to the divestment of FDI, the ER may harm domestic enter-
450 prises. Therefore, future research should focus on how to protect the environment while ensur-
451 ing the development of domestic and foreign companies. Second, since the implementation of
452 ecological supervision is difficult to observe, more regional heterogeneous adjustment factors,
453 such as institutional quality and corruption, should be considered.

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461 **Availability of data and materials**

462 The datasets used and/or analyzed during the current study are available from the corresponding
463 author on reasonable request.

464 **Competing interests**

465 The authors declare that they have no known competing financial interests or personal relationships
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467 **Authors' contributions**

468 Lufeng Tai: conceptualization, formal analysis, methodology, software, writing. Linnan Yan: data
469 collection, review and editing, funding acquisition, investigation, supervision.

471 **References**

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Figures

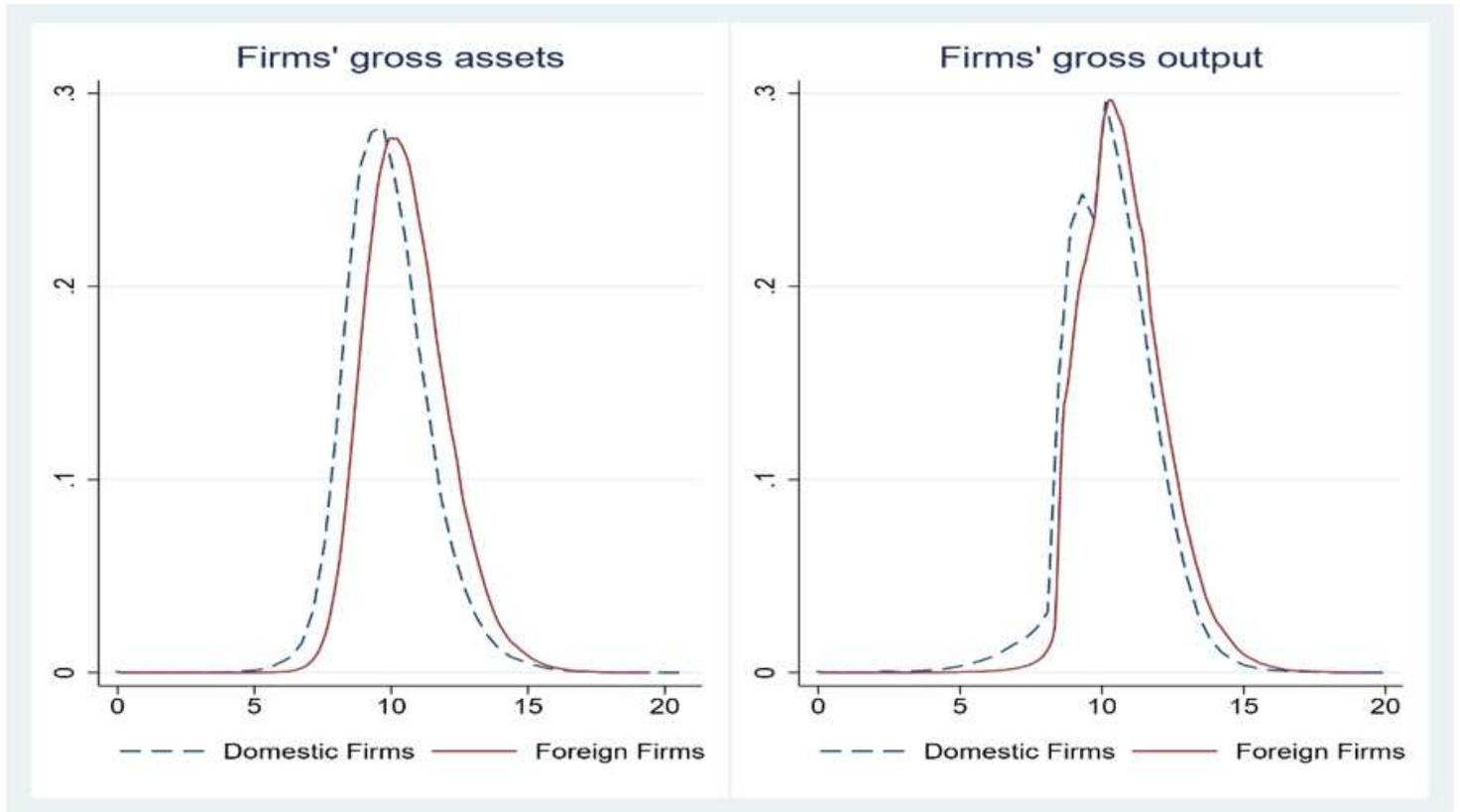


Figure 1

The kernel density estimation of the scale of foreign firms and domestic firms.