

Bivariate Genomic Prediction of Phenotypes by Selecting Epistatic Interactions Across Years Based on Haplotype Blocks and Pruned Sets of SNPs

Elaheh Vojgani (✉ vojgani@gwdg.de)

University of Goettingen, Center for Integrated Breeding Research, Animal Breeding and Genetics Group, Goettingen, Germany <https://orcid.org/0000-0003-4375-3531>

Torsten Pook

University of Goettingen, Center for integrated Breeding Research, Animal Breeding and Genetics Group, Goettingen

Armin C. Hölker

Technical University of Munich, School of Life Sciences Weihenstephan, Plant Breeding, Freising

Manfred Mayer

Technical University of Munich, TUM School of Life Sciences Weihenstephan, Plant Breeding, Freising

Chris-Carolin Schön

Technical University of Munich, TUM School of Life Sciences Weihenstephan, Plant Breeding, Freising

Henner Simianer

University of Goettingen, Center for integrated Breeding Research, Animal Breeding and Genetics Group, Goettingen

Research Article

Keywords: Epistasis, Bivariate GBLUP, Prediction across years, Genomic correlation, Haplotype blocks

Posted Date: May 20th, 2021

DOI: <https://doi.org/10.21203/rs.3.rs-519981/v1>

License:   This work is licensed under a Creative Commons Attribution 4.0 International License.

[Read Full License](#)

1 **Bivariate genomic prediction of phenotypes by selecting epistatic interactions**
2 **across years based on haplotype blocks and pruned sets of SNPs**

3 **Elaheh Vojgani^{1*}, Torsten Pook¹, Armin C. Hölker², Manfred Mayer², Chris-Carolin Schön²,**
4 **Henner Simianer¹**

5 ¹ University of Goettingen, Center for Integrated Breeding Research, Animal Breeding and
6 Genetics Group, Goettingen, Germany; ² Plant Breeding, TUM School of Life Sciences
7 Weihenstephan, Technical University of Munich, Freising, Germany

8

9 **Key Message**

10 A bivariate epistasis model increases the prediction accuracy similarly for selected subsets of
11 epistatic interactions when utilizing haplotype blocks and pruned sets of SNPs by incorporating
12 phenotypic data across years.

13 **Abstract**

14 The importance of accurate genomic prediction of phenotypes in plant breeding is undeniable,
15 as higher prediction accuracy can increase selection responses. In this study, we investigated the
16 ability of three models to improve prediction accuracy by including phenotypic information from
17 the last growing season. This was done by considering a single biological trait in two growing
18 seasons (2017 and 2018) as separate traits in a multi-trait model. Thus, bivariate variants of the
19 Genomic Best Linear Unbiased Prediction (GBLUP) as an additive model, Epistatic Random
20 Regression BLUP (ERRBLUP) and selective Epistatic Random Regression BLUP (sERRBLUP) as
21 epistasis models were compared with respect to their prediction accuracies for the second year.
22 The results indicate that bivariate ERRBLUP is almost identical to bivariate GBLUP in prediction
23 accuracy, while bivariate sERRBLUP has the highest prediction accuracy in most cases. The
24 obtained prediction accuracies were similar when utilizing pruned sets of SNPs and haplotype
25 blocks, while utilizing haplotype blocks reduces the computational load significantly compared to
26 utilizing pruned sets of SNPs. The prediction accuracies of bivariate GBLUP, ERRBLUP and
27 sERRBLUP have been assessed across eight phenotypic traits and studied datasets from 471/402
28 doubled haploid lines in the European maize landrace Kemater Landmais Gelb/Petkuser
29 Ferdinand Rot. We further investigated the genomic correlation, phenotypic correlation and trait
30 heritability as factors affecting the bivariate models' prediction accuracy, with genetic correlation
31 between growing seasons being the most important one. For all three considered model
32 architectures results were far worse when using a univariate version of the model.

33 **Keywords:**

34 Epistasis, Bivariate GBLUP, Prediction across years, Genomic correlation, Haplotype blocks

35

36 **Declaration**

37 **Funding**

38 This work was funded by German Federal Ministry of Education and Research (BMBF) within the
39 scope of the funding initiative “Plant Breeding Research for the Bioeconomy” (MAZE – “Accessing
40 the genomic and functional diversity of maize to improve quantitative traits”; Funding ID:
41 031B0195)

42 **Conflict of interest**

43 On behalf of all authors, the corresponding author states that there is no conflict of interest.

44 **Ethics approval**

45 The authors declare that this study complies with the current laws of the countries in which the
46 experiments were performed.

47 **Consent to participate**

48 Not applicable

49 **Consent for publication**

50 Not applicable

51 **Availability of data and materials**

52 All data and material are available through material transfer agreements upon request.

53 **Code availability**

54 Not applicable

55 **Authors' contributions**

56 EV proposed epistasis models based on haplotype blocks, analyzed the data, wrote the
57 manuscript; TP proposed epistasis relationship matrices; ACH, MM and CCS prepared the
58 material; ACH proposed cross validation strategy in bivariate model; HS proposed the original
59 research question, guided the structure of the research. TP ACM MM CCS HS read, revised and
60 approved the manuscript.

61 **Corresponding author**

62 Correspondence to Elaheh Vojgani

63 Email: vojgani@gwdg.de

64 ORCID: 0000-0003-4375-3531

65

66 **Acknowledgements**

67 We are thankful to KWS SAAT SE, Misión Biológica de Galicia, Spanish National Research Council
68 (CSIC), Technical University of Munich, and University of Hohenheim for providing the extensive
69 phenotypic evaluation. We are grateful to the German Federal Ministry of Education and
70 Research (BMBF) for the funding of our project within the scope of the funding initiative “Plant
71 Breeding Research for the Bioeconomy” (MAZE – “Accessing the genomic and functional diversity
72 of maize to improve quantitative traits”; Funding ID: 031B0195).

73

74 Introduction

75 In plant breeding, genomic prediction has become a regular tool (Bernal-Vasquez *et al.* 2014;
76 Stich and Ingheland 2018) which enables the optimization of phenotyping costs of breeding
77 programs (Akdemir and Isidro-Sánchez 2019). The importance of genomic prediction of
78 phenotypes is not restricted to plants. Livestock (Daetwyler *et al.* 2013) and human research (de
79 los Campos *et al.* 2013) also have been widely developed in this regard. In the context of plant
80 and animal breeding, accurately predicting phenotypic traits is of special importance, since
81 raising all animals and growing all crops to measure their performances requires a considerable
82 amount of money under limited resources (Martini *et al.* 2016).

83 Several statistical models have been compared over the last decades in the term of prediction
84 accuracy. In this context, genomic best linear unbiased prediction (GBLUP) (Meuwissen *et al.*
85 2001; VanRaden 2007) as an additive linear mixed model has been widely used due to its high
86 robustness, computing speed and superiority in predictive ability to alternative prediction
87 models like Bayesian methods, especially in small reference populations (Da *et al.* 2014;
88 Rönnegård and Shen 2016; Covarrubias-Pazarán *et al.* 2018; Wang *et al.* 2018). Furthermore, the
89 inclusion of genotype \times environment interaction into additive genomic prediction models can
90 result in an increase in prediction accuracy (Hallauer *et al.* 2010; Bajgain *et al.* 2020). Such
91 approaches allow borrowing information across environments which potentially leads to higher
92 accuracy in phenotype prediction in multi environment models (Burgueño *et al.* 2012). In fact,
93 multivariate mixed models have been originally proposed in the context of animal breeding
94 (Henderson and Quaas 1976) with the purpose of modeling the genomic correlation among traits,
95 longitudinal data, and modeling genotype by environment interactions across multiple years or
96 environments (Mrode 2014; Lee and van der Werf 2016; Covarrubias-Pazarán *et al.* 2018). A
97 multivariate GBLUP model was reported to have higher prediction accuracy than univariate
98 GBLUP (Jia and Jannink 2012) when the genetic correlations were medium (0.6) or high (0.9)
99 (Covarrubias-Pazarán *et al.* 2018). It was also shown that aggregating the phenotypic data over
100 years to train the model and predict the performance of lines in the following years is a possible
101 approach which can improve prediction accuracy (Auinger *et al.* 2016; Schrag *et al.* 2019a).

102 In addition, the inclusion of epistasis, defined as the interaction between loci (Falconer and
103 Mackay 1996; Lynch and Walsh 1998), into the genomic prediction model results in more
104 accurate phenotype prediction (Hu *et al.* 2011; Wang *et al.* 2012; Mackay 2014; Martini *et al.*
105 2016; Vojgani *et al.* 2021) due to the considerable contribution of epistasis in genetic variation of
106 quantitative traits (Mackay 2014). In this context, several statistical models have been proposed.
107 Extended genomic best linear unbiased prediction (EG-BLUP, Jiang and Reif 2015) and categorical
108 epistasis (CE, Martini *et al.* 2017) models are using a marker-based epistatic relationship matrix
109 that is constructed in a highly efficient manner. It has been shown that the CE model is as good
110 as or better than EG-BLUP and does not possess undesirable features of EG-BLUP such as coding-
111 dependency (Martini *et al.* 2017).

112 Moreover, it was shown that the accuracy of the epistasis genomic prediction model can be
113 increased in one environment by variable selection in another environment (Martini *et al.* 2016).
114 In this approach, the full epistasis model was reduced to a model with a subset of the largest
115 epistatic interaction effects, resulting in an increase in predictive ability (Martini *et al.* 2016),
116 through borrowing information across environments. Vojgani *et al.* (2021) showed that the
117 prediction accuracy can be increased even further by selecting the interactions with the highest
118 absolute effect sizes / variances in the epistasis model. The resulting higher computational needs
119 were offset by the development of a highly efficient software package “EpiGP” (Vojgani *et al.*
120 2019) to perform computations in a bit-wise manner (Schlather 2020). Thus, enabling to conduct
121 such predictions with data sets of practically relevant size across environments in the same year,
122 both with respect to sample size and number of markers (Vojgani *et al.* 2021). As the number of
123 interactions to account for increases quadratically with the number of included variables, the
124 computational load of methods like EpiGP can quickly go out of control as a model with 600.000
125 SNPs, as present in high density arrays (Kranis *et al.* 2013; Unterseer *et al.* 2014), would result in
126 more than a hundred billion interactions to account for. The most common methods for variable
127 reduction applied here is LD pruning (Purcell *et al.* 2007), but new linkage-based haplotyping
128 methods (Pook *et al.* 2019) have recently been proposed to even further reduce the
129 dimensionality of genomic data without much information loss.

130 The aim of this study is to assess the bivariate genomic prediction models which incorporate
131 epistatic interactions with the target of borrowing information across years to maximize the
132 predictive ability based on both, a pruned set of SNPs and haplotype blocks. Since the accuracy
133 of genomic prediction of phenotypes was shown to be increased by both borrowing information
134 across environments and years (Covarrubias-Pazarán *et al.* 2018; Schrag *et al.* 2019b) and
135 inclusion of epistasis into the prediction model (Martini *et al.* 2016; Vojgani *et al.* 2020), we
136 combine these two approaches to make the best use of the available information. We further
137 aim to assess the optimum proportion of epistatic interactions to be kept in the model in the
138 variable selection step across years and compare the obtained predictive ability when utilizing
139 pruned sets of SNPs and haplotype blocks. The data used for this purpose were generated in
140 multi-location trials of doubled haploid (DH) lines generated from two European maize landraces
141 in 2017 and 2018.

142 **Materials and Methods**

143 **Data used for analysis**

144 A set of 948 doubled haploid lines of the European maize landraces Kemater Landmais Gelb (KE,
145 Austria, 516 lines) and Petkuser Ferdinand Rot (PE, Germany, 432 lines) were genotyped with the
146 600 k Affymetrix® Axiom® Maize Array (Unterseer *et al.* 2014).

147 After quality filtering and imputation, 910 DH lines remained (501 lines in KE and 409 lines in PE)
148 and the panel of markers reduced to 501,124 markers (Hölker *et al.* 2019). Additionally, loci that
149 were in high level of pairwise linkage disequilibrium (LD) were removed (Calus and Vandenplas

150 2018) through linkage disequilibrium based SNP pruning with PLINK v1.07 (Purcell *et al.* 2007;
151 Chang *et al.* 2015). LD pruning was done by the parameters of 50, 5 and 2 which considered as
152 the SNPs window size, the number of SNPs at which the SNP window shifts and the variance
153 inflation factor, respectively. This resulted in a data panel containing 25'437 SNPs for KE and
154 30'212 SNPs for PE (Vojgani *et al.* 2020). Note that even a panel of 25'000 SNPs results in more
155 than 1 billion SNP interactions to account for. Therefore, in order to have further variable
156 reduction, haplotype blocks as a combination of closely linked markers, which has shown to be
157 an alternative approach for genomic prediction improving the prediction accuracy (Meuwissen
158 *et al.* 2014; Jiang *et al.* 2018), have been generated from the full panel of markers with the
159 software HaploBlocker (Pook *et al.* 2019) using default settings. This resulted in a data panel
160 containing 2'972 haplotype blocks in KE and 3'330 haplotype blocks in PE.

161 Out of 910 genotyped lines only 873 DH lines were phenotyped (471 lines in KE and 402 lines in
162 PE). Einbeck (EIN, Germany), Roggenstein (ROG, Germany), Golada (GOL, Spain) and Tomeza
163 (TOM, Spain) were the four locations that these lines were phenotyped for a series of traits in
164 both 2017 and 2018.

165 The means, standard deviations, maximum and minimum values of studied phenotypic traits in
166 2017 and 2018 in each landrace are compared in Table 1 which were derived from the Best Linear
167 Unbiased Estimations (BLUEs) of the genotype mean for each phenotypic trait by Hölker *et al.*
168 (2019). The comparison of the respective detailed values for each trait in each environment and
169 landrace in 2017 and 2018 are illustrated in the supplementary (Table S1). V_i in phenotypic traits
170 represents the vegetative growth stage when i leaf collars are visible based on the leaf collar
171 method of the corn growth (Abendroth *et al.* 2011). Early vigour at V3 stage (EV_V3), female
172 flowering (FF) and root lodging (RL) were not phenotyped in all four environments for both years.
173 EV_V3 was not phenotyped in EIN in 2018, FF was not phenotyped in GOL in 2017 and RL was not
174 phenotyped in TOM and GOL in both 2017 and 2018.

175 The number of phenotyped lines per year and environment for trait PH_V4, as the main trait in
176 this study, are summarized in Table 2. For EIN and ROG a higher number of phenotyped lines
177 were generated in 2017. On the contrary, more lines were phenotypes in GOL and TOM in 2018.

178 **Statistical models for phenotype prediction**

179 We used the bivariate statistical framework as the basis of the genomic prediction models which
180 has been proposed in the recent work by Vojgani *et al.* (2020). In this regard, GBLUP, ERRBLUP
181 and sERRBLUP as three different methods described in Vojgani *et al.* (2021) were used for
182 genomic prediction of phenotypes which differ in dispersion matrices representing their
183 covariance structure of the genetic effects. GBLUP as an additive model is based on a genomic
184 relationship matrix calculated according to VanRaden (2008). ERRBLUP (Epistatic Random
185 Regression BLUP) as a full epistasis model is based on all pairwise SNP interactions which
186 generates a new marker matrix considered as a marker combination matrix. The marker
187 combination matrix is a 0, 1 matrix indicating the absence (0) or presence (1) of each marker

188 combination for each individual. sERRBLUP (selective Epistatic Random Regression BLUP) as a
189 selective epistasis model is based on a selected subset of SNP interactions (Vojgani *et al.* 2021).
190 Vojgani *et al.* (2020) proposed estimated effect variances in the training set as the selection
191 criterion of pairwise SNP interactions due to its robustness in predictive ability specifically when
192 only a small proportion of interactions are maintained in the model.

193 **Assessment of genomic prediction models**

194 GBLUP, ERRBLUP and sERRBLUP models have been assessed via 5-fold cross validation by
195 randomly partitioning the original sample into 5 equal size subsamples in which one subsample
196 was considered as the test set to validate the model, and the remaining 4 subsamples were
197 considered as a joint training set (Erbe *et al.* 2010). The 5-fold cross validation technique was
198 utilized with 5 replicates through which the Pearson correlation between the predicted genetic
199 values and the observed phenotypes in the test set was considered as the predictive ability in
200 each fold of each replicate, which then was averaged across 25 replicates. In this study, predictive
201 ability was separately assessed for KE and PE for a series of phenotypic traits in four different
202 environments. Besides, we calculated the traits' prediction accuracies by dividing their predictive
203 abilities by the square-root of the respective traits' heritabilities (Dekkers 2007) derived from all
204 environments in both 2017 and 2018 jointly (Table S2).

205 Univariate GBLUP within 2018 was assessed by training the model in the same year (2018) as the
206 test set was sampled from. However, bivariate GBLUP, ERRBLUP and sERRBLUP were assessed by
207 training the model with both the training set of the target environment in 2018 and the full
208 dataset of the respective environment in 2017. The interaction selection step in bivariate
209 sERRBLUP is done by first using the complete dataset of target environment in 2017 to estimate
210 all pairwise SNP interaction effect variances. Then, an epistatic relationship matrix for all lines is
211 constructed based on the subset of top ranked interaction effect variances, which is finally used
212 to predict phenotypes of the target environment test set in 2018 (Vojgani *et al.* 2020).

213 **Variance component estimation**

214 Variance component estimation in univariate GBLUP was done by EMMREML (Akdemir and
215 Godfrey 2015) based on the training set in each run of 5-fold cross validation with 5 replicates.
216 In bivariate models this was done by ASReml-R (Butler *et al.* 2018) with the approach specified
217 by Vojgani *et al.* (2020) for pre estimating the variance components from the full dataset to derive
218 the initial values for the variance components in ASReml models in 100 iterations for each
219 combination. If the variance estimation based on the full set did not converge after 100
220 iterations, then the estimated variance components at the 100th iteration were extracted as
221 initial values of the bivariate model in the cross validation step. Afterwards, the model used these
222 values to re-estimate the variance components based on the training set in each run of 5-fold
223 cross validation in 50 iterations. The estimated variance components in the converged models
224 based on the full set deviated only slightly from the estimated variance components based on
225 the training set (Fig. S1). However, the variance component estimations did not converge in all

226 folds of 5-fold cross validation with 5 replicates. In such cases, the initial values were set as the
227 fixed values for the model to predict the breeding values. This approach appears justifiable in the
228 case of non-convergence of the bivariate model, since we have shown in Fig. S2 that the
229 difference in mean predictive ability of all folds and only the converged folds is not critical. This
230 difference can get higher as the number of non-converged folds increases. The number of not
231 converged folds in all studied material is shown in the supplementary (Table S3).

232 **Genomic correlation estimation**

233 Genomic correlations were estimated from the genetic variances and covariance derived from
234 the ASReml bivariate model based on the full dataset of each location in both 2017 and 2018.

235 **Results**

236 **Comparison of univariate GBLUP, bivariate GBLUP, bivariate ERRBLUP and bivariate sERRBLUP** 237 **based on pruned set of SNPs and haplotype blocks in PH_V4**

238 Our results confirm that bivariate models outperform the univariate models (Vojgani *et al.* 2020)
239 as illustrated by the comparison in predictive ability of bivariate GBLUP and univariate GBLUP for
240 the trait PH-V4 in both landraces indicating the superiority of bivariate GBLUP to univariate
241 GBLUP in most cases (see Fig. 1 and Fig. 2). Among the bivariate genomic prediction models, the
242 predictive ability obtained from bivariate ERRBLUP is almost identical to bivariate GBLUP. This
243 predictive ability increases in bivariate sERRBLUP and the highest gain in accuracy is generally
244 obtained when the top 10 or 5 percent of pairwise SNP interactions are kept in the model. A too
245 strict selection like using only the top 0.001 percent interactions, results in a decrease in
246 predictive ability (see Fig. 1 and Fig. 2).. The number of interactions maintained in the model for
247 each proportion of interactions are tabulated in the supplementary (Table S4). Robustness of the
248 predictive ability depending on the share of selected markers was higher in PE than KE. Moreover,
249 Fig. 1 and Fig. 2 illustrate the comparison between the predictive abilities obtained from the
250 respective genomic prediction models in KE and PE when utilizing pruned set of SNPs and
251 haplotype blocks. It is shown that the GBLUP, ERRBLUP and sERRBLUP (for the optimum
252 proportions of interactions) predictive abilities are almost identical in both pruned set of SNPs
253 and haplotype blocks. It should be noted that the robustness of sERRBLUP when a very small
254 proportions of interactions maintained in the model is higher when utilizing the pruned set of
255 SNPs compared to haplotype blocks. This should not be surprising as the total number of
256 interactions in the HaploBlocker panel is much smaller, thus, leading to a dataset with an
257 extremely low number of explanatory variables (Table S4). Similar patterns are observed across
258 a series of other traits for bivariate models which are shown in the supplementary (Fig. S3-S16).
259 Additionally, the predictive ability of univariate GBLUP by training the model on the average
260 phenotypic values of both 2017 and 2018, when utilizing pruned set of SNPs, was evaluated for
261 a series of phenotypic traits, which yielded quite similar predictive ability as obtained with
262 univariate GBLUP within year 2018 or worse in some cases (Table S5 (KE) and S6 (PE)).

263 **Correlation between prediction accuracy and the genomic correlation in bivariate models**

264 The absolute gain in predictive ability from univariate GBLUP to maximum bivariate sERRBLUP,
265 when utilizing pruned set of SNPs, was regressed on the respective sERRBLUP genomic
266 correlation between the two respective environments and across the series of studied traits (Fig.
267 3). Regression coefficients range between 0.14 and 0.48 and thus show a clear association
268 between the absolute gain in prediction accuracy and the genomic correlation between
269 environments. When combining all traits and environments, this correlation is 0.65 (p-value =
270 0.00018) in KE and 0.69 (p-value = 4.393e-05) in PE. This correlation is also significant for most of
271 the environments when utilizing haplotype blocks (Fig. S17).

272 **Interplay of GBLUP and sERRBLUP prediction accuracy and genomic correlations**

273 The genomic correlations across years estimated with GBLUP and sERRBLUP based on pruned set
274 of SNPs for the trait PH_V4 are illustrated in Table 3, indicating that the proportion of interactions
275 in bivariate sERRBLUP which maximized the predictive ability are not necessarily linked to the
276 highest genomic correlation. In contrast, the best sERRBLUP for trait PH_V4 is linked to the lowest
277 genomic correlation in most cases. However, this is not the general pattern observed for a series
278 of other traits and the best sERRBLUP for some traits and environments combinations are linked
279 to the highest genomic correlation (Table S7-S13). In fact, there is a significant correlation
280 between the absolute increase in predictive ability from bivariate GBLUP to maximum bivariate
281 sERRBLUP and the difference between genetic correlations estimated with GBLUP and maximum
282 sERRBLUP in both KE and PE when utilizing pruned set of SNPs (Fig. S18).

283 **Correlation between prediction accuracy and the phenotypic correlation in bivariate models**

284 There might be some tendency that including phenotypes of the previous year into prediction
285 becomes more efficient when the phenotypic correlation between years is high. In this context,
286 the correlation between the absolute gain in predictive ability from univariate GBLUP to
287 maximum bivariate sERRBLUP and the phenotypic correlation among the years (see Table S14)
288 over all studied traits in all four environments and in both landraces was studied. Fig. 4
289 demonstrates that the maximum correlation between the absolute gain in the respective
290 predictive ability based on the pruned set of SNPs and the phenotypic correlation is obtained in
291 EIN for KE (0.42) and in ROG for PE (0.62). Across all studied traits and environments, there is a
292 significant correlation of 0.55 in KE (p-value= 0.003) and 0.50 in PE (p-value= 0.007). This
293 correlation is also significant in most of the environments when utilizing haplotype blocks (Fig.
294 S19).

295 **Relative increase in prediction accuracy across all traits in all environments and landraces with** 296 **bivariate models**

297 Overall, the percentage of relative increase in prediction accuracy from the bivariate GBLUP to
298 the maximum bivariate sERRBLUP based on pruned set of SNPs in both landraces are illustrated

299 in Fig. 5 with the average increase of 7.61 percent in KE and 3.47 percent in PE over all studied
300 traits. Among all traits, the maximum relative increase in prediction accuracy for KE is 22.63
301 percent which was obtained in EV_V6 in EIN, and for PE is 34.59 percent which was obtained in
302 EV_V4 in EIN. However, Fig. 5 shows some slight decreases in prediction accuracy from bivariate
303 GBLUP to maximum bivariate sERRBLUP for some combinations of traits and environment in both
304 landraces. This is more often observed in PE than KE, where the maximum decrease was found
305 in EV_V6 in TOM for both PE (-3.198 percent) and KE (-2.795 percent). Overall, the average
306 relative increase from bivariate GBLUP to maximum bivariate sERRBLUP was over 3 percent in
307 most cases. The absolute increase in prediction accuracy is also illustrated in the supplementary
308 (Fig. S20) indicating the average increase of 0.046 in KE and 0.015 in PE over all combinations of
309 traits and environments. In addition, the absolute increase in prediction accuracy from the
310 bivariate GBLUP to the maximum bivariate sERRBLUP based on haplotype blocks is shown in the
311 supplementary indicating the average absolute increase of 0.034 in KE and 0.013 in PE (Fig. S21).
312 Overall, the increase in prediction accuracy from bivariate GBLUP to maximum bivariate
313 sERRBLUP is significantly higher in KE than PE in both cases of utilizing pruned set of SNPs and
314 haplotype blocks (Fig. S22).

315 Discussion

316 In this study, bivariate ERRBLUP as a full epistasis model incorporating all pairwise SNP
317 interactions is almost identical to bivariate GBLUP. This was expected, since ERRBLUP
318 incorporates a high number of interactions by which a large number of unimportant variables are
319 introduced into the model (Martini *et al.* 2016), thus introducing potential ‘noise’ which can
320 prevent gains in predictive ability. In contrast, bivariate sERRBLUP substantially increases the
321 predictive ability compared to bivariate GBLUP which is only caused by inclusion of relevant
322 pairwise SNP interactions. Note that all bivariate models substantially outperformed univariate
323 GBLUP, as phenotypic data of the respective environment in the previous year was used.
324 ERRBLUP and sERRBLUP models have shown to display similar behaviors in the context of
325 univariate statistical setting in the maize dataset for prediction across environments (Vojgani *et*
326 *al.* 2020). ERRBLUP, which had been introduced as categorical epistasis (CE) model by Martini *et*
327 *al.* (2017), performs as good as the best EG-BLUP which is EG-BLUP with symmetric marker
328 coding. Similarly, selection of optimum proportions of interactions in EG-BLUP has shown to
329 increase the predictive ability compared to the EG-BLUP which includes all pairwise SNP
330 interactions (Martini *et al.* 2016) for a wheat dataset (Pérez and de los Campos 2014).
331 Consequently, ERRBLUP and sERRBLUP leads to higher predictive ability than EG-BLUP and
332 reduced EG-BLUP with non-symmetric coded markers, respectively. This was shown by Vojgani
333 *et al.* (2021) in the wheat dataset (Pérez and de los Campos 2014).

334 Furthermore, in this study we have found that GBLUP, ERRBLUP and maximum sERRBLUP
335 predictive abilities when utilizing haplotype blocks are very similar to the respective models’
336 predictive abilities when utilizing pruned sets of SNPs. This finding is of high relevance in practice,
337 since it helps to overcome the high computational load of epistatic models. The required

338 computational time for sERRBLUP based on 3'330 haplotype blocks indicating 5'546'115
339 interactions was 9 minutes out of which 4 minutes were needed to estimate the pairwise SNP
340 interaction effect variances and 5 minutes were needed to generate the sERRBLUP relationship
341 matrix for a selected proportion of interactions by utilizing the R-package miraculix with 15 cores
342 on a server cluster with Intel E5-2650 (2X12 core 2.2GHz) processors in the released EpiGP R-
343 package (Vojgani *et al.* 2019). As the computing time is increasing approximately quadratically in
344 the number of included markers, the computing time for the respective SNP-based model with
345 30'212 SNPs took 90 times as long (Vojgani *et al.* 2020), while it resulted in similar predictive
346 abilities with the absolute difference being less than 0.01 in most cases across all traits in all
347 environments and both landraces (Fig. S23). Although this difference in predictive abilities is
348 statistically significant based on paired t-test, it is not of practical relevance.

349 Although, sEERBLUP is a method that is using multiple environments, it is not a GxE model (de
350 Leon *et al.* 2016) in the traditional sense. While GxE models typically assign effect to combination
351 of specific genotypes depending on the environment, the second environment in sEERBLUP is
352 “only” used to detect which markers affect a given trait and use the information to put more
353 focus on these marker in the actual prediction step. The estimation of marker effects itself is then
354 executed only based on the environment itself (or in the case of the bivariate model with some
355 contributions from the second environment but still not in the sense of a traditional GxE model).
356 As a different model is used for each environment, sEERBLUP will of course still assign different
357 marker effects in different environments. As the set of selected marker interactions will however
358 be different between different models, a direct comparison of the effects assigned to specific
359 marker interactions is not statistically sound. Similar to most GxE models (Shin and Lee 2020),
360 the computational load of sEERBLUP using markers is extremely high. However, our suggested
361 use of haplotype blocks massively reduced this problem, while the predictive ability is almost as
362 good as sERRBLUP based on pruned set of SNPs.

363 It was shown that multivariate GBLUP is superior in predictive ability compared to univariate
364 GBLUP with of medium (~ 0.6) to high (~ 0.9) genomic correlations, and that low genomic
365 correlations results in no increase in multivariate GBLUP compared to univariate GBLUP
366 (Covarrubias-Pazarán *et al.* 2018). Calus *et al.* (2011) also found an increase of 3 to 14 percent in
367 predictive ability of multi-trait SNP-based models in a simulation study when genetic correlations
368 ranged from 0.25 to 0.75. In our study, we also found a significant correlation between the
369 absolute gain in prediction accuracy from univariate GBLUP to maximum bivariate sERRBLUP and
370 the respective genomic correlation based on both pruned sets of SNPs ($r_{KE} = 0.65$, $r_{PE} = 0.69$)
371 and haplotype blocks ($r_{KE} = 0.47$, $r_{PE} = 0.49$) across all traits and environments combinations.

372 Moreover, Martini *et al.* (2016) showed that the predictive ability in one environment can be
373 increased by variable selection in the other environment under the assumption of positive
374 phenotypic correlation between environments. It was shown in a wheat dataset (Pérez and de
375 los Campos 2014), where environments 2 and 3 had the highest phenotypic correlation (0.661),
376 that the predictive ability for phenotype prediction in environment 2 was maximized by variable

377 selection in environment 3 and vice versa (Martini *et al.* 2016). Therefore, the increase in
378 prediction accuracy is expected to be influenced by the phenotypic correlations between the
379 environments or between the years in the same environment in bivariate models. In our study,
380 although 2017 and 2018 were climatically quite different, since 2018 suffered from a major heat
381 stress compared to 2017 (Table 1), we see a significant correlation between the absolute gain in
382 predictive ability from univariate GBLUP to maximum predictive ability of bivariate sERRBLUP and
383 the phenotypic correlation between years in each environment based on both pruned sets of
384 SNPs ($r_{KE} = 0.55$, $r_{PE} = 0.50$) and haplotype blocks ($r_{KE} = 0.55$, $r_{PE} = 0.56$).

385 In addition to the genomic and phenotypic correlations between the years, the trait heritability
386 is another factor which is expected to be influential for such an increase in bivariate sERRBLUP
387 predictive ability as well. Therefore, the traits with lower heritability are expected to obtain less
388 gain in sERRBLUP predictive ability than the traits with higher heritability. This was confirmed in
389 our study, as traits with low heritability (e.g. 0.59 for RL in PE) showed only a small increase in
390 prediction accuracy from univariate GBLUP to maximum bivariate sERRBLUP. However, not all
391 traits with higher heritabilities did necessarily show a higher gain in predictive ability for all traits.
392 It should be noted that the trait heritabilities were calculated on an entry-mean basis (Hallauer
393 *et al.* 2010) within each KE and PE landraces by Hölker *et al.* (2019) over all four environments in
394 both years 2017 and 2018 jointly. The trait heritabilities obtained only from 2017 are significantly
395 higher than the trait heritabilities obtained only from 2018 in both KE and PE based on a paired
396 t-test (Table S2). This also results in an increase in predictive ability from univariate GBLUP to
397 maximum bivariate sERRBLUP in KE and PE, since multi-trait models have the potential of
398 increasing the predictive ability when traits with low heritability are joined with traits with higher
399 heritability, given they are genomically correlated (Thompson and Meyer 1986).

400 It should be noted that the increase in predictive ability from univariate GBLUP to maximum
401 bivariate sERRBLUP is caused by both borrowing information across years and capitalizing on
402 epistasis, while the increase in predictive ability from bivariate GBLUP to maximum bivariate
403 sERRBLUP is caused by accounting for epistasis alone. Overall, the traits behave differently
404 among different environments and landraces due to their genomic correlations, phenotypic
405 correlations and heritabilities. To shed light on this, the maximum increase in prediction accuracy
406 from bivariate GBLUP to bivariate sERRBLUP based on pruned set of SNPs in KE was observed for
407 the trait EV_V6 (0.112) in EIN where the corresponding sERRBLUP genomic correlation (0.809) is
408 higher than the GBLUP genomic correlation (0.768). This trait has a high heritability (0.90) and
409 high phenotypic correlation (0.551) as well. In contrast, the respective prediction accuracy
410 decreases (-0.018) for EV_V6 in TOM for KE indicating the lower sERRBLUP genomic correlation
411 (0.458) than GBLUP genomic correlation (0.703) and the particularly low phenotypic correlation
412 (0.383). It should be noted that the phenotypic correlation does not play a major role for the
413 increase in prediction accuracy from bivariate GBLUP to bivariate sERRBLUP, since both models
414 are bivariate and benefit from the same phenotypic correlations. Therefore, EV_V6 obtaining the
415 maximum and minimum increase in the respective prediction accuracy for KE indicates the
416 significant role of genomic correlation among the possible causes. In general, bivariate sERRBLUP

417 improves the prediction accuracy compared to bivariate GBLUP more in KE than PE which is
418 potentially due to significantly higher sERRBLUP genomic correlation and heritability in KE
419 compared to PE, based on paired t-test.

420 Overall, our results indicate that incorporating a suitable subset of epistatic interactions besides
421 utilizing information across years can substantially increase the predictive ability. The amount of
422 this increase is affected by the genomic and phenotypic correlations between the years and the
423 heritability of the phenotypic trait. Moreover, utilizing haplotype blocks instead of pruned sets
424 of SNPs in epistasis model is proposed, since the obtained predictive abilities of the best epistasis
425 model are quite similar, while the required computational time when utilizing haplotype blocks
426 is significantly lower than the required computational time when utilizing pruned set of SNPs.
427 Therefore, this computationally efficient approach is potentially beneficial for genomic
428 prediction of phenotypes under the assumption of sufficient genomic and phenotypic correlation
429 between years for highly heritable traits. This may allow to reduce the number of lines which
430 have to be phenotyped over several years and thus reduce phenotyping costs which and thus be
431 of high interest in practical plant breeding.

432

433 **Figures and Tables Captions**

434

435 **Fig. 1** Predictive ability for univariate GBLUP within 2018 (orange and red dashed horizontal line),
436 bivariate GBLUP (green and blue dashed horizontal line), bivariate ERRBLUP (open circle) and
437 bivariate sERRBLUP (filled circles and solid line) for trait PH-V4 in KE based on Pruned set of SNPs
438 (left) and haplotype blocks (right). In each plot, the sERRBLUP maximum indicates the maximum
439 predictive ability obtained from bivariate sERRBLUP.

440 **Fig. 2** Predictive ability for univariate GBLUP within 2018 (orange and red dashed horizontal line),
441 bivariate GBLUP (green and blue dashed horizontal line), bivariate ERRBLUP (open circle) and
442 bivariate sERRBLUP (filled circles and solid line) for trait PH-V4 in PE based on pruned sets of SNPs
443 (left) and haplotype blocks (right). In each plot, the sERRBLUP maximum indicates the maximum
444 predictive ability obtained from bivariate sERRBLUP.

445 **Fig. 3** Regression of the absolute increase in predictive ability from univariate GBLUP to maximum
446 bivariate sERRBLUP on the respective sERRBLUP genomic correlation between 2017 and 2018 in
447 KE (left) and in PE (right) for all studied traits. In each panel, the overall linear regression line (gray
448 solid line) with the regression coefficient (***b***) and R-squared (***R*²**) are shown.

449 **Fig. 4** Regression of the absolute increase in predictive ability from univariate GBLUP to maximum
450 bivariate sERRBLUP on the phenotypic correlation between 2017 and 2018 in KE (left) and in PE
451 (right) for all studied traits. In each panel, the overall linear regression line (gray solid line) with
452 the regression coefficient (***b***) and R-squared (***R*²**) are shown.

453 **Fig. 5** Percentage of change in prediction accuracy from bivariate GBLUP to the maximum
454 prediction accuracy of bivariate sERRBLUP based on pruned set of SNPs in KE (left side plot) and
455 in PE (right side plot). The average percentage of change in prediction accuracy for each trait and
456 environment is displayed in all rows and columns, respectively.

457 **Table 1** Phenotypic trait description and the mean, minimum, maximum and standard deviation
458 of the BLUEs for each phenotypic trait in KE and PE landraces in the years 2017 and 2018.

459 **Table 2** Number of KE and PE lines phenotyped in each location for the years 2017 (blue numbers)
460 and 2018 (red numbers) for trait PH_V4.

461 **Table 3** Genomic correlation between 2017 and 2018 in each environment for trait PH_V4 for KE
462 (blue numbers) and PE (red numbers). The blue and red bold numbers with stars indicate which
463 proportion of interactions in bivariate sERRBLUP maximized the predictive ability based on
464 pruned set of SNPs in each environment for KE and PE, respectively.

465 **References**

- 466 Abendroth LJ, Elmore RW, Boyer MJ, and Marlay SK (2011) Corn Growth and Development.
467 *PMR 1009. Iowa State University of Science and Technology, Cooperative Extension Service,*
468 *Ames, Iowa.*
- 469 Akdemir D and Godfrey OU (2015) EMMREML: Fitting Mixed Models with Known Covariance
470 Structures. Available at: <https://cran.r-project.org/package=EMMREML>
- 471 Akdemir D and Isidro-Sánchez J (2019) Design of training populations for selective phenotyping
472 in genomic prediction. *Scientific Reports* 9(1446).
473 <https://doi.org/https://doi.org/10.1038/s41598-018-38081-6>
- 474 Auinger H-J, Schönleben M, Lehermeier C, Schmidt M, Korzun V, Geiger HH, Piepho H-P,
475 Gordillo A, Wilde P, Bauer E, and Schön C-C (2016) Model training across multiple breeding
476 cycles significantly improves genomic prediction accuracy in rye (*Secale cereale* L.). *Theoretical*
477 *and Applied Genetics* 129(11): 2043–2053. <https://doi.org/10.1007/s00122-016-2756-5>
- 478 Bajgain P, Zhang X, and Anderson JA (2020) Dominance and G×E interaction effects improve
479 genomic prediction and genetic gain in intermediate wheatgrass (*Thinopyrum intermedium*).
480 *The Plant Genome*. John Wiley & Sons, Ltd 13(1): e20012.
481 <https://doi.org/https://doi.org/10.1002/tpg2.20012>
- 482 Bernal-Vasquez A-M, Möhring J, Schmidt M, Schönleben M, Schön C-C, and Piepho H-P (2014)
483 The importance of phenotypic data analysis for genomic prediction - a case study comparing
484 different spatial models in rye. *BMC Genomics* 15(1): 646. [https://doi.org/10.1186/1471-2164-](https://doi.org/10.1186/1471-2164-15-646)
485 [15-646](https://doi.org/10.1186/1471-2164-15-646)
- 486 Burgueño J, Campos G de los, Weigel K, and Crossa J (2012) Genomic Prediction of Breeding
487 Values when Modeling Genotype × Environment Interaction using Pedigree and Dense
488 Molecular Markers. *Crop Science* 52(2): 707–719. <https://doi.org/10.2135/cropsci2011.06.0299>
- 489 Butler DG, Cullis BR, Gilmour AR, Gogel BJ, and Thompson R (2018) ASReml-R Reference
490 Manual Version 4. VSN International Ltd., Hemel Hempstead
- 491 Calus MPL and Vandenplas J (2018) SNPPrune: an efficient algorithm to prune large SNP array
492 and sequence datasets based on high linkage disequilibrium. *Genetics Selection Evolution* 50(1):
493 34. <https://doi.org/10.1186/s12711-018-0404-z>
- 494 Calus MPL and Veerkamp RF (2011) Accuracy of multi-trait genomic selection using different
495 methods. *Genetics Selection Evolution* 43(1): 26. <https://doi.org/10.1186/1297-9686-43-26>
- 496 Chang CC, Chow CC, Tellier LC, Vattikuti S, Purcell SM, and Lee JJ (2015) Second-generation
497 PLINK: rising to the challenge of larger and richer datasets. *Gigascience* 4(7).
498 <https://doi.org/10.1186/s13742-015-0047-8>
- 499 Covarrubias-Pazarán G, Schlautman B, Diaz-García L, Grygleski E, Polashock J, Johnson-Cicalese
500 J, Vorsa N, Iorizzo M, and Zalapa J (2018) Multivariate GBLUP Improves Accuracy of Genomic
501 Selection for Yield and Fruit Weight in Biparental Populations of *Vaccinium macrocarpon* Ait.

502 *Frontiers in Plant Science* 9(1310). <https://doi.org/10.3389/fpls.2018.01310>

503 Da Y, Wang C, Wang S, and Hu G (2014) Mixed Model Methods for Genomic Prediction and
504 Variance Component Estimation of Additive and Dominance Effects Using SNP Markers. *PLOS*
505 *ONE* 9(1). <https://doi.org/10.1371/journal.pone.0087666>

506 Daetwyler HD, Calus MPL, Pong-Wong R, Campos G de los, and Hickey JM (2013) Genomic
507 Prediction in Animals and Plants: Simulation of Data, Validation, Reporting, and Benchmarking.
508 *Genetics* 193: 347–365. <https://doi.org/10.1534/genetics.112.147983>

509 Dekkers JCM (2007) Prediction of response to marker-assisted and genomic selection using
510 selection index theory. *Journal of Animal Breeding and Genetics* 124: 331–341.
511 <https://doi.org/10.1111/j.1439-0388.2007.00701.x>

512 Erbe M, Pimentel E, Sharifi AR, and Simianer H (2010) Assessment of cross-validation strategies
513 for genomic prediction in cattle. *9th World Congress of Genetics Applied to Livestock Production*
514 129–132

515 Falconer DS and Mackay TFC (1996) *Introduction to Quantitative Genetics*. Longman. Essex Engl.

516 Hallauer AR, Carena MJ, and Miranda Filho JB (2010) *Quantitative genetics in maize breeding*.
517 Springer. Berlin

518 Henderson CR and Quaas RL (1976) Multiple Trait Evaluation Using Relatives' Records. *Journal*
519 *of Animal Science* 43(6): 1188–1197. <https://doi.org/10.2527/jas1976.4361188x>

520 Hölker AC, Mayer M, Presterl T, Bolduan T, Bauer E, Ordas B, Brauner PC, Ouzunova M,
521 Melchinger AE, and Schön C-C (2019) European maize landraces made accessible for plant
522 breeding and genome-based studies. *Theoretical and Applied Genetics* 132(12): 3333–3345.
523 <https://doi.org/10.1007/s00122-019-03428-8>

524 Hu Z, Li Y, Song X, Han Y, Cai X, Xu S, and Li W (2011) Genomic value prediction for quantitative
525 traits under the epistatic model. *BMC Genet* 12(15).
526 <https://doi.org/https://doi.org/10.1186/1471-2156-12-15>

527 Jia Y and Jannink J-L (2012) Multiple-Trait Genomic Selection Methods Increase Genetic Value
528 Prediction Accuracy. *Genetics* 192(4): 1513 LP – 1522.
529 <https://doi.org/10.1534/genetics.112.144246>

530 Jiang Y and Reif JC (2015) Modeling Epistasis in Genomic Selection. *Genetics* 201(2): 759–768.
531 <https://doi.org/10.1534/genetics.115.177907>

532 Jiang Y, Schmidt RH, and Reif JC (2018) Haplotype-Based Genome-Wide Prediction Models
533 Exploit Local Epistatic Interactions Among Markers. *G3: Genes/Genomes/Genetics* 8(5): 1687 LP
534 – 1699. <https://doi.org/10.1534/g3.117.300548>

535 Kranis A, Gheyas AA, Boschiero C, Turner F, Yu L, Smith S, Talbot R, Pirani A, Brew F, Kaiser P,
536 Hocking PM, Fife M, Salmon N, Fulton J, Strom TM, Haberer G, Weigend S, Preisinger R,
537 Gholami M, Qanbari S, Simianer H, Watson KA, Woolliams JA, and Burt DW (2013)
538 Development of a high density 600K SNP genotyping array for chicken. *BMC Genomics* 14(1):

539 59. <https://doi.org/10.1186/1471-2164-14-59>

540 Lee SH and van der Werf JHJ (2016) MTG2: an efficient algorithm for multivariate linear mixed
541 model analysis based on genomic information. *Bioinformatics* 32(9): 1420–1422.
542 <https://doi.org/10.1093/bioinformatics/btw012>

543 de Leon N, Jannink J-L, Edwards JW, and Kaeppler SM (2016) Introduction to a Special Issue on
544 Genotype by Environment Interaction. *Crop Science*. John Wiley & Sons, Ltd 56(5): 2081–2089.
545 <https://doi.org/10.2135/cropsci2016.07.0002in>

546 de los Campos G, Vazquez AI, Fernando R, Klimentidis YC, and Sorensen D (2013) Prediction of
547 Complex Human Traits Using the Genomic Best Linear Unbiased Predictor. *PLoS Genetics* 9(7).
548 <https://doi.org/10.1371/journal.pgen.1003608>

549 Lynch M and Walsh B (1998) *Genetics and Analysis of Quantitative Traits*. Sinauer Associates

550 Mackay TFC (2014) Epistasis and Quantitative Traits: Using Model Organisms to Study Gene-
551 Gene Interactions. *Nat Rev Genet*. 15(1): 22–33. <https://doi.org/10.1038/nrg3627>

552 Martini JWR, Wimmer V, Erbe M, and Simianer H (2016) Epistasis and covariance: how gene
553 interaction translates into genomic relationship. *Theoretical and Applied Genetics* 129(5): 963–
554 976. <https://doi.org/10.1007/s00122-016-2675-5>

555 Martini JWR, Gao N, Cardoso DF, Wimmer V, Erbe M, Cantet RJC, and Henner S (2017)
556 Genomic prediction with epistasis models: on the marker-coding-dependent performance of
557 the extended GBLUP and properties of the categorical epistasis model (CE). *BMC Bioinformatics*
558 18(3). <https://doi.org/10.1186/s12859-016-1439-1>

559 Meuwissen THE, Odegard J, Andersen-Ranberg I, and Grindflek E (2014) On the distance of
560 genetic relationships and the accuracy of genomic prediction in pig breeding. *Genetics Selection
561 Evolution* 46(1): 49. <https://doi.org/10.1186/1297-9686-46-49>

562 Meuwissen THE, Hayes BJ, and Goddard ME (2001) Prediction of total genetic value using
563 genome-wide dense marker maps. *Genetics* 157(4): 1819–1829

564 Mrode RA (2014) *Linear Models for the Prediction of Animal Breeding Values*. CABI.
565 <https://doi.org/10.1079/9781780643915.0000>

566 Pérez P and de los Campos G (2014) Genome-wide regression and prediction with the BGLR
567 statistical package. *Genetics*. 2014/07/09. Genetics Society of America 198(2): 483–495.
568 <https://doi.org/10.1534/genetics.114.164442>

569 Purcell S, Neale B, Todd-Brown K, Thomas L, Ferreira MAR, Bender D, Maller J, Sklar P, Bakker
570 PIW de, Daly MJ, and Sham PC (2007) PLINK: A Tool Set for Whole-Genome Association and
571 Population-Based Linkage Analyses. *American Journal of Human Genetics* 81(3): 559–575.
572 <https://doi.org/10.1086/519795>

573 Rönnegård L and Shen X (2016) Genomic prediction and estimation of marker interaction
574 effects. *bioRxiv* 38935. <https://doi.org/10.1101/038935>

575 Schlather M (2020) Efficient Calculation of the Genomic Relationship Matrix. *bioRxiv*.
576 <https://doi.org/10.1101/2020.01.12.903146>

577 Schrag TA, Schipprack W, and Melchinger AE (2019a) Across-years prediction of hybrid
578 performance in maize using genomics. *Theoretical and Applied Genetics*. Springer Verlag 132(4):
579 933–946. <https://doi.org/10.1007/s00122-018-3249-5>

580 Schrag TA, Schipprack W, and Melchinger AE (2019b) Across-years prediction of hybrid
581 performance in maize using genomics. *Theoretical and Applied Genetics* 132: 933–946

582 Shin J and Lee SH (2020) GxEsum: genotype-by-environment interaction model based on
583 summary statistics. *bioRxiv* 2020.05.31.122549. <https://doi.org/10.1101/2020.05.31.122549>

584 Stich B and Ingheland D Van (2018) Prospects and Potential Uses of Genomic Prediction of Key
585 Performance Traits in Tetraploid Potato. *Frontiers in Plant Science* 9(159).
586 <https://doi.org/10.3389/fpls.2018.00159>

587 Thompson R and Meyer K (1986) A review of theoretical aspects in the estimation of breeding
588 values for multi-trait selection. *Livestock Production Science* 15(4): 299–313.
589 [https://doi.org/https://doi.org/10.1016/0301-6226\(86\)90071-0](https://doi.org/https://doi.org/10.1016/0301-6226(86)90071-0)

590 Unterseer S, Bauer E, Haberer G, Seidel M, Knaak C, Ouzunova M, Meitinger T, Strom TM, Fries
591 R, Pausch H, Bertani C, Davassi A, Mayer KF, and Schön C-C (2014) A powerful tool for genome
592 analysis in maize: 584 development and evaluation of the high density 600 k SNP genotyping
593 array. *BMC Genomics* 15(823). <https://doi.org/10.1186/1471-2164-15-823>

594 VanRaden P (2007) Efficient estimation of breeding values from dense genomic data. *Journal of*
595 *Dairy Science* 90: 374–375

596 VanRaden P (2008) Efficient methods to compute genomic predictions. *Journal of Dairy Science*
597 91(11): 4414–4423. <https://doi.org/10.3168/jds.2007-0980>

598 Vojgani E, Pook T, Martini JWR, Hoelker AC, Mayer M, Schoen C-C, and Simianer H (2020)
599 Accounting for epistasis improves genomic prediction of phenotypes with univariate and
600 bivariate models across environments. *bioRxiv* 2020.10.08.331074.
601 <https://doi.org/10.1101/2020.10.08.331074>

602 Vojgani E, Pook T, and Simianer H (2019) EpiGP: Epistatic relationship matrix based genomic
603 prediction of phenotypes. Available at: <https://github.com/evoigani/EpiGP>

604 Vojgani E, Pook T, and Simianer H (2021) Phenotype Prediction under Epistasis. in KC, W. (ed.)
605 *Epistasis: Methods and Protocols*. Springer. https://doi.org/10.1007/978-1-0716-0947-7_8

606 Wang D, El-Basyoni IS, Baenziger PS, Crossa J, Eskridge KM, and Dweikat I (2012) Prediction of
607 genetic values of quantitative traits with epistatic effects in plant breeding populations.
608 *Heredity* 109(5): 313–319. <https://doi.org/10.1038/hdy.2012.44>

609 Wang J, Zhou Z, Zhang Zhe, Li H, Liu D, Zhang Q, Bradbury PJ, Buckler ES, and Zhang Zhiwu
610 (2018) Expanding the BLUP alphabet for genomic prediction adaptable to the genetic
611 architectures of complex traits. *Heredity* 121(6): 648–662. <https://doi.org/10.1038/s41437-018->

612 0075-0

613

Figures

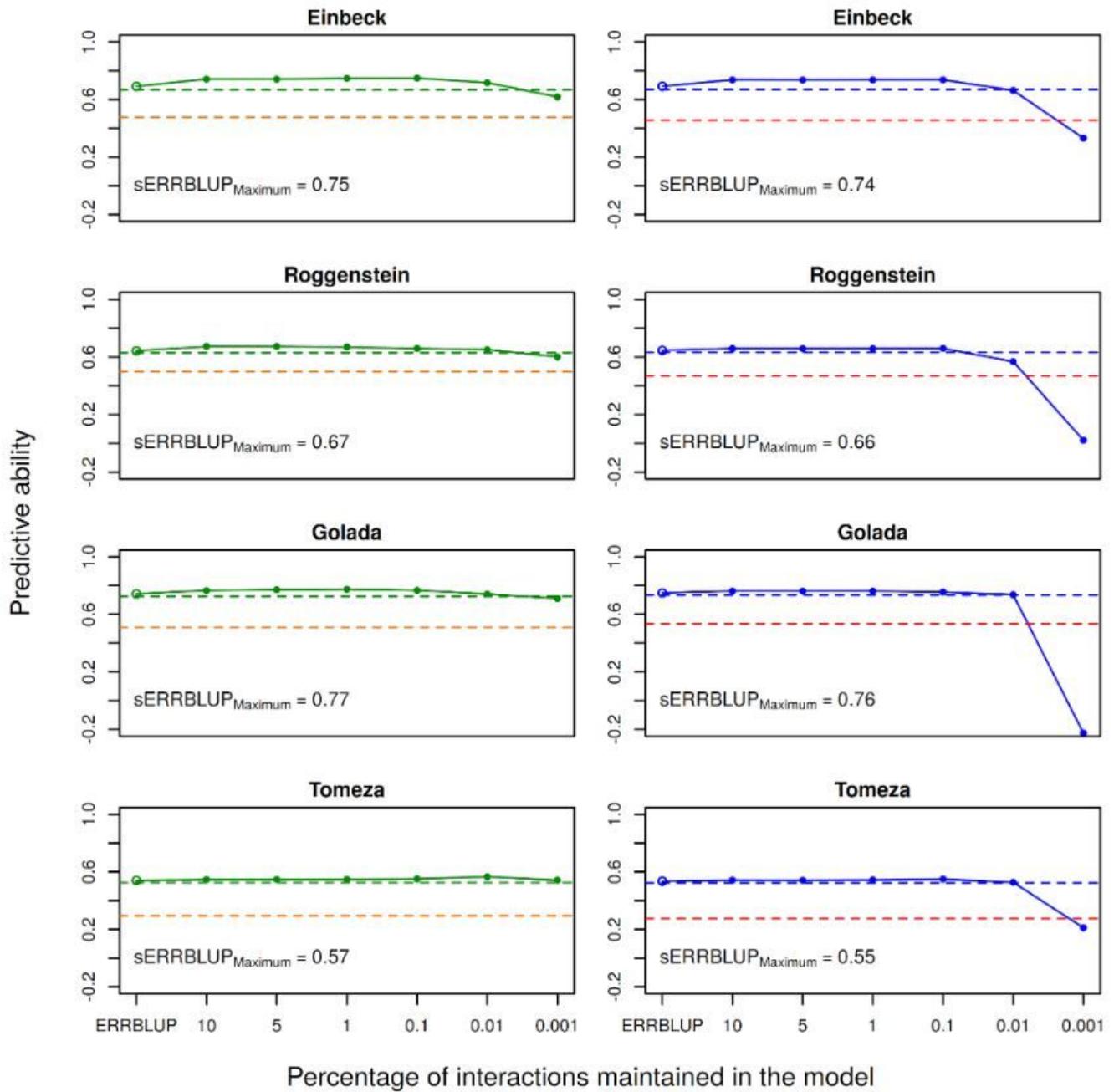


Figure 1

Predictive ability for univariate GBLUP within 2018 (orange and red dashed horizontal line), bivariate GBLUP (green and blue dashed horizontal line), bivariate ERRBLUP (open circle) and bivariate sERRBLUP (filled circles and solid line) for trait PH-V4 in KE based on Pruned set of SNPs (left) and haplotype blocks (right). In each plot, the sERRBLUP maximum indicates the maximum predictive ability obtained from bivariate sERRBLUP.

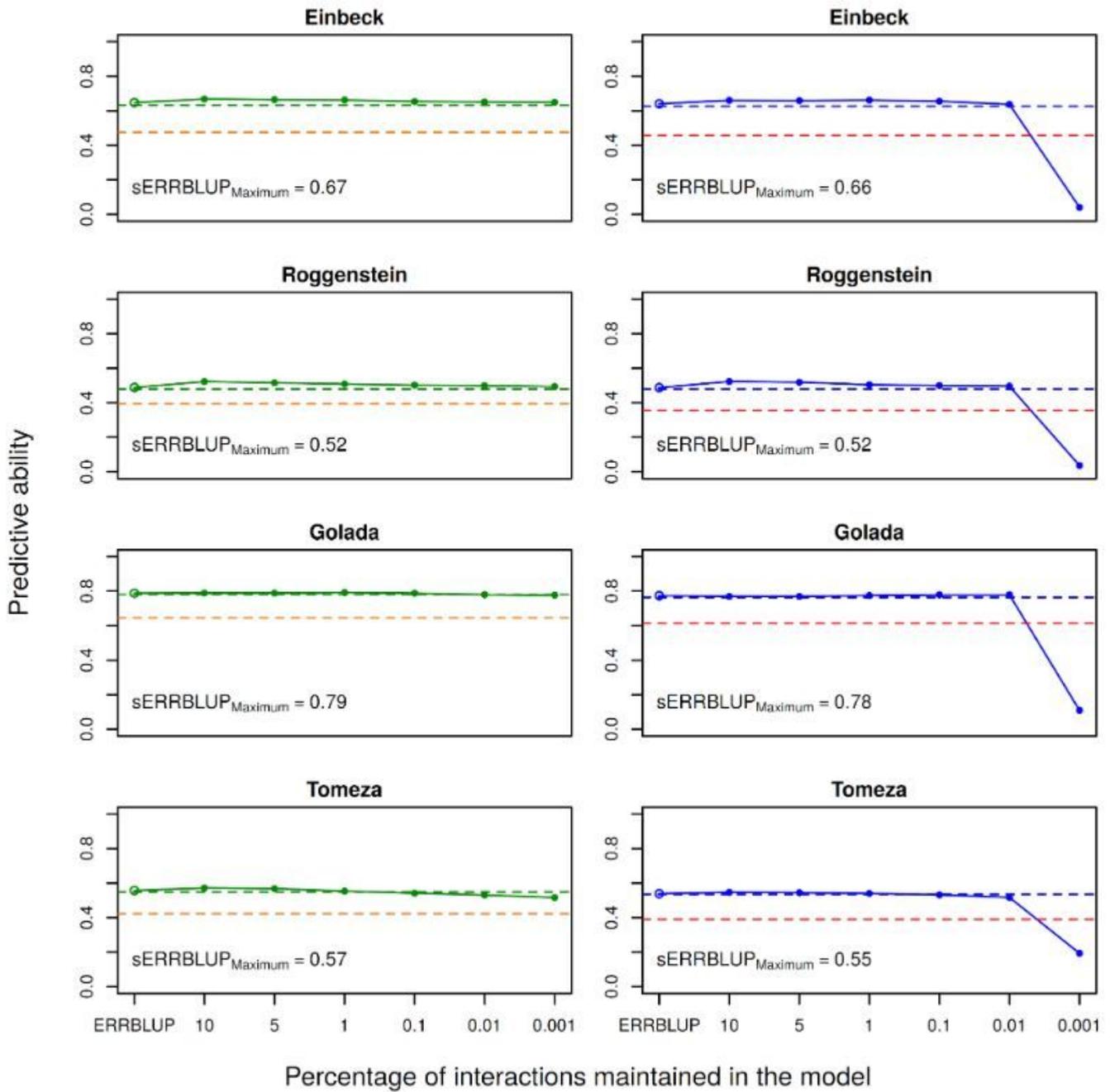


Figure 2

Predictive ability for univariate GBLUP within 2018 (orange and red dashed horizontal line), bivariate GBLUP (green and blue dashed horizontal line), bivariate ERRBLUP (open circle) and bivariate sERRBLUP (filled circles and solid line) for trait PH-V4 in PE based on pruned sets of SNPs (left) and haplotype blocks (right). In each plot, the sERRBLUP maximum indicates the maximum predictive ability obtained from bivariate sERRBLUP.

Increase in predictive ability from univariate GBLUP to maximum bivariate sERRBLUP based on pruned SNPs

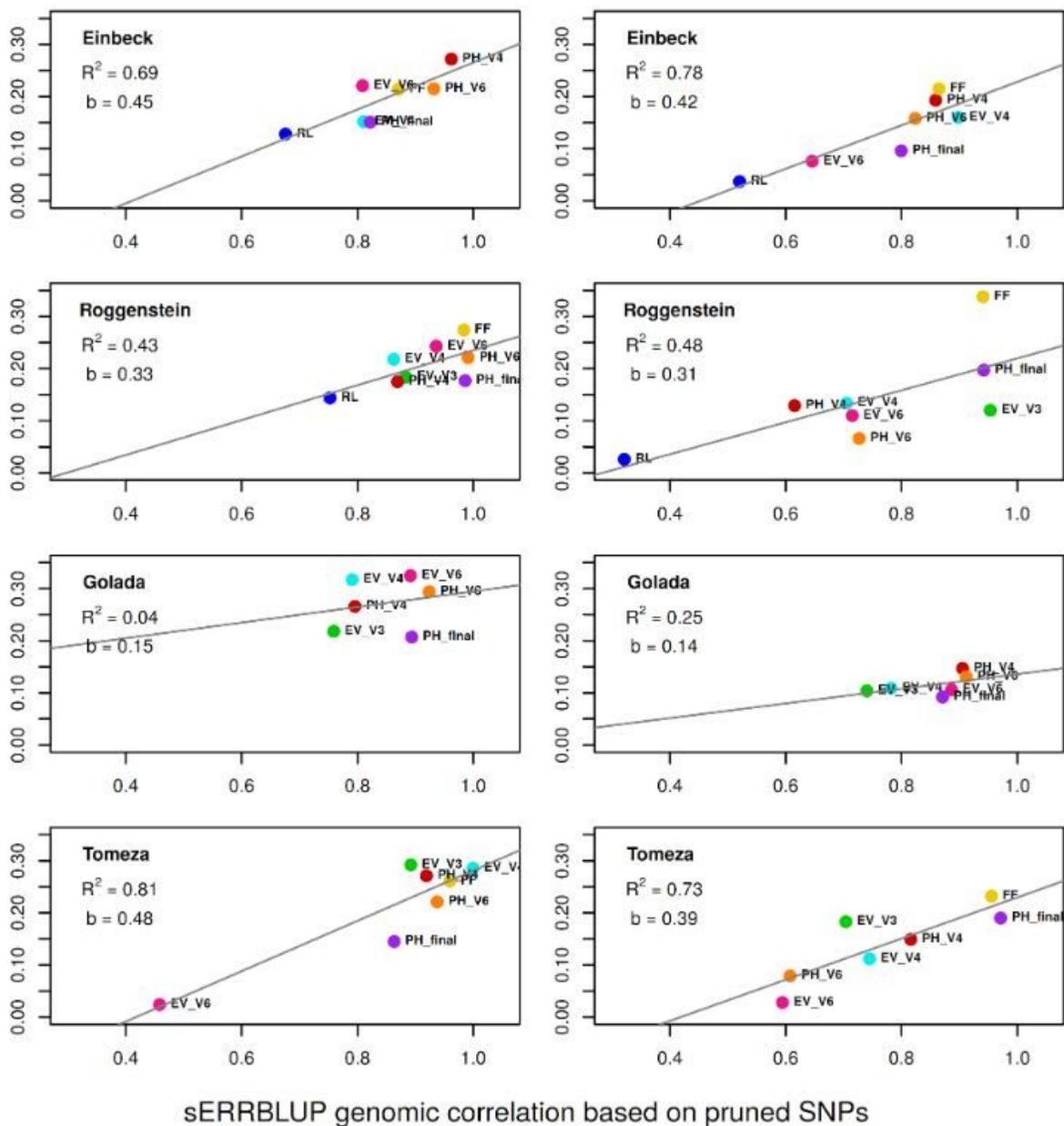


Figure 3

Regression of the absolute increase in predictive ability from univariate GBLUP to maximum bivariate sERRBLUP on the respective sERRBLUP genomic correlation between 2017 and 2018 in KE (left) and in PE (right) for all studied traits. In each panel, the overall linear regression line (gray solid line) with the regression coefficient (b) and R-squared (R²) are shown.

Increase in predictive ability from univariate GBLUP to maximum bivariate sERRBLUP based on pruned SNPs

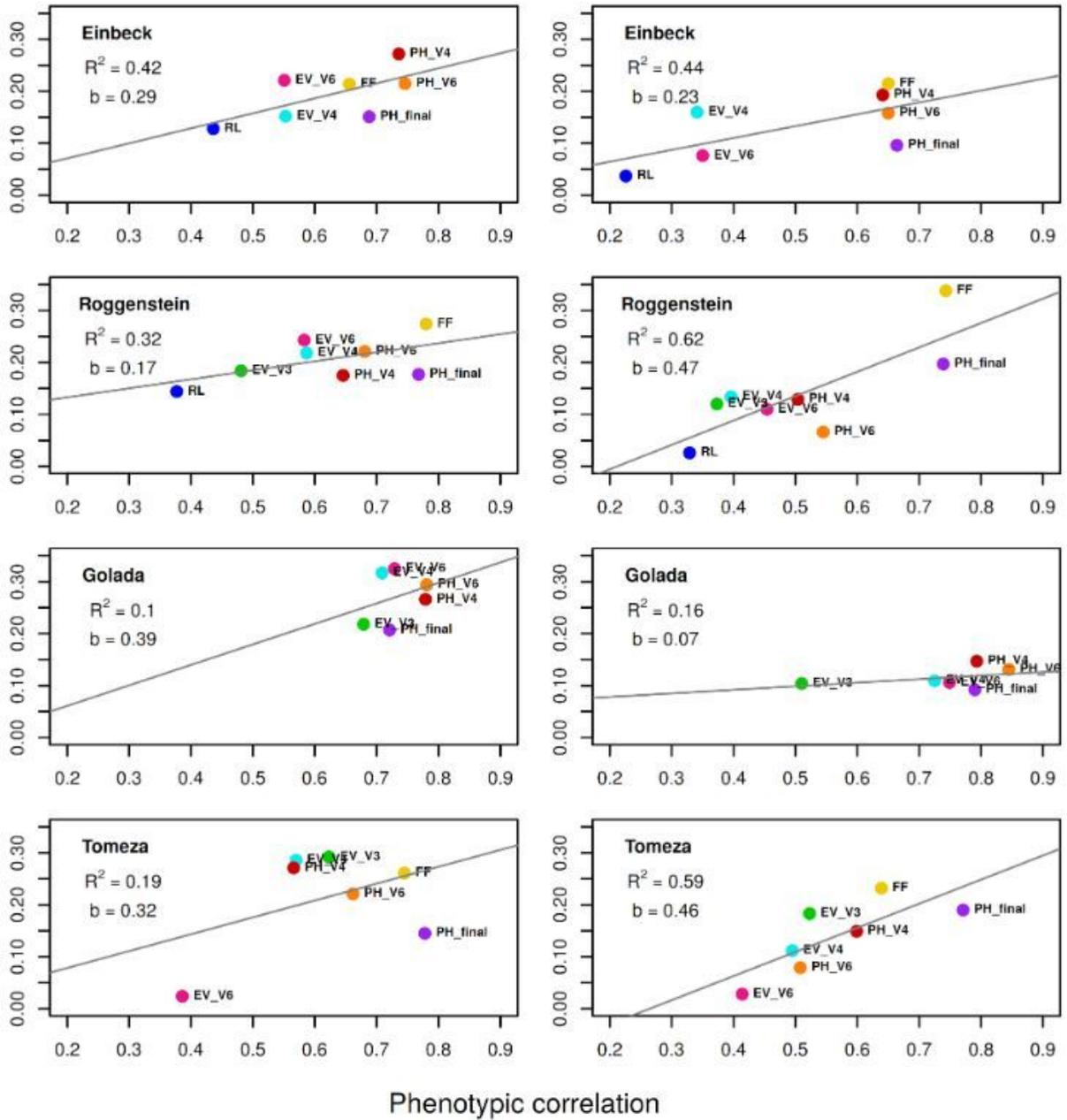


Figure 4

Regression of the absolute increase in predictive ability from univariate GBLUP to maximum bivariate sERRBLUP on the phenotypic correlation between 2017 and 2018 in KE (left) and in PE (right) for all studied traits. In each panel, the overall linear regression line (gray solid line) with the regression coefficient (b) and R-squared (R^2) are shown.

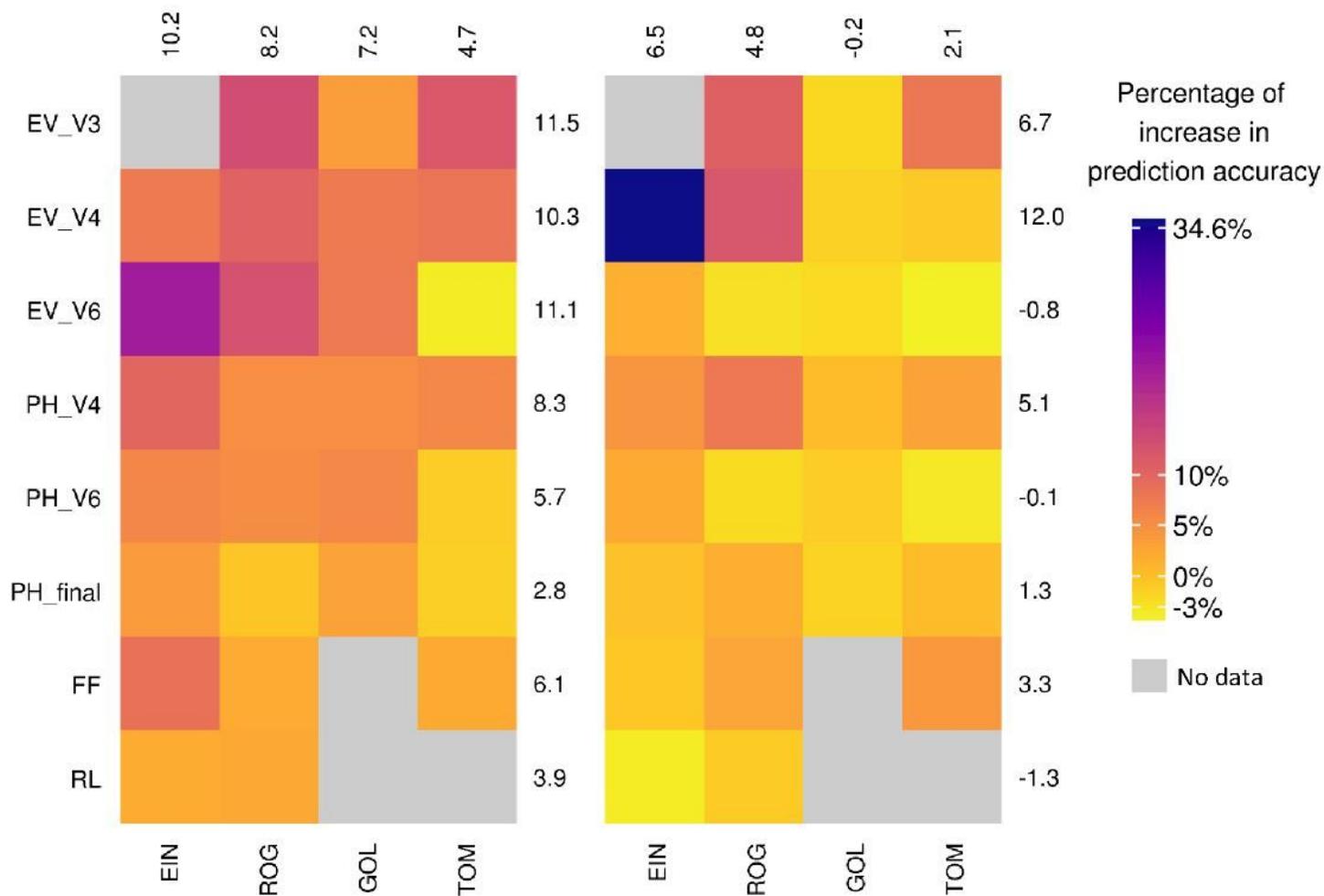


Figure 5

Percentage of change in prediction accuracy from bivariate GBLUP to the maximum prediction accuracy of bivariate sERRBLUP based on pruned set of SNPs in KE (left side plot) and in PE (right side plot). The average percentage of change in prediction accuracy for each trait and environment is displayed in all rows and columns, respectively.

Supplementary Files

This is a list of supplementary files associated with this preprint. Click to download.

- [ElectronicSupplementaryMaterial.pdf](#)