

Factors Associated With Successful Publication for Systematic Review Protocol Registration: an Analysis of 397 Registered Protocols.

Le Huu Nhat Minh

Ho Chi Minh City University of Medicine and Pharmacy Faculty of Medicine

Huu-Hoai Le

Ho Chi Minh City University of Medicine and Pharmacy Faculty of Medicine

Gehad Mohamed Tawfik

Ain Shams University Faculty of Medicine

Omar Mohamed Makram

October 6 University

Thuan Tieu

McGill University

Luu Lam Thang Tai

City Children Hospital Ho Chi Minh City

Karim Mohamed Shahin

Alexandria University

Ali Ahmed-Fouad Abozaid

Ain Shams University

Jaffer Shah

Drexel University

Nguyen Hai Nam

Kyoto University: Kyoto Daigaku

Nguyen Tien Huy (✉ tienhuy@nagasaki-u.ac.jp)

Institute of Tropical Medicine (NEKKEN) <https://orcid.org/0000-0002-9543-9440>

Research

Keywords: Protocol, Systematic review, Meta-analysis, Protocol registration, Prospero, Cochrane, JBI

Posted Date: June 17th, 2021

DOI: <https://doi.org/10.21203/rs.3.rs-617877/v1>

License:  This work is licensed under a Creative Commons Attribution 4.0 International License.

[Read Full License](#)

Version of Record: A version of this preprint was published at Systematic Reviews on June 2nd, 2023.

See the published version at <https://doi.org/10.1186/s13643-023-02210-8>.

Abstract

Background: Meta-analyses are on top of the evidence-based medicine pyramid, yet many of them are not completed after they are begun. Many factors impacting the publication of meta-analysis works have been discussed, and their association with publication likelihood has been investigated. These factors include the type of systematic review, journal metrics, h-index of the corresponding author, country of the corresponding author, funding sources, and duration of publication. In our current review, we aim to investigate these various factors and their impact on the likelihood of publication. A comprehensive review of 397 registered protocols retrieved from five databases was performed to investigate the different factors that might affect the likelihood of publication. These factors include the type of systematic review, journal metrics, h-index of the corresponding author, country of the corresponding author, funding sources, and duration of publication.

Results: We found that corresponding authors in developed countries and English-speaking countries had higher likelihoods of publication: 206/320 ($p=0.018$) and 158/236 ($p=0.006$), respectively. Our models of multivariable logistic regression analysis revealed that two main factors impact the publication outcome: updated protocol status of the published review paper (OR: 1.7, 95% CI: 1.0-2.8, $p=0.037$) and external funding (OR: 2.1, 95% CI: 1.2-3.8, $p=0.01$). However, corresponding authors' location in developed countries (OR: 1.7, 95% CI: 0.8-3.4, $p=0.139$) and English-speaking countries (OR: 1.5, 95% CI: 0.9-2.6, $p=0.1$) were insignificant determinants.

Conclusion: Meta-analyses continue to be on top of the evidence hierarchy, rendering them the key to informed clinical decision-making. Therefore, more attention should be paid to the methodological quality of this type of publication.

Introduction

Systematic reviews (SR) and meta-analyses (MA) are considered to be the highest tier of the evidence-based medicine pyramid due to their ability to arrive at an empirical algorithm for diseases by combining results from different studies conducted over many years[1]. However, several studies point out that many registered SRs are not been published for a long time after registration. For example, about 20% of Cochrane protocols were not published as full reviews within eight years[2, 3], while another study reported that 26% of PROSPERO protocols (from 2/2011 to 2/2012) remained unpublished after at least 65 months[4].

Recently, some publications have shed light on the topic of registered protocols that were not published. One study related the lack of publication to financial factors, finding that funded reviews were more likely to be published[4, 5]. A survey in 2009, in which 625 authors participated, found that lack of time, funding, and organizational support were the main barriers to finishing the reviews[6]. Lack of time was also cited by a survey in 2018, which concluded that SRs with protocols took more than twice the time from search

to submission than SRs without protocols [2], while other studies need even more time – up to 2.4 years [3].

Nevertheless, these publications were limited in scope and the set of factors that may impact the likelihood of publication of registered protocols is not well understood. Studies may restrict their attention to reviews registered in one database, such as PROSPERO, and miss other databases such as Cochrane, Joanna Briggs Institute (JBI), Campbell, and others.

In our current review, we investigated various factors that might affect the likelihood of publication. These points included the different registration websites, types of SR, the h-index of the corresponding author, publication date, duration of study work time, funding sources, and journal metrics at the time of publication[7].

Methods

Identifying protocols: Search terms and inclusion/exclusion criteria

Five databases (EMBASE, PROSPERO, Cochrane, PubMed, and Scopus) were searched for SR/MA protocols published in 2013 with the following terms in their title: protocol, systematic, review, and meta-analysis. We selected this period because it represented the maximal time that a protocol would have the chance to be published (5 years) from the year in which we conducted this search method (26/07/2018), as estimated from the previous study[8]. After searching, two independent authors examined the results and excluded papers that were either abstract-only, not SR/MA protocols, or had no available full-text. Any discrepancies were resolved by discussion and consensus among the senior author (N.T.H). The protocols were later divided into five groups according to their registered websites: PROSPERO, Cochrane, JBI, SJR, and Others. The other databases include Annals of Cardiothoracic Surgery, BMC Medical Research Methodology, BMC Psychiatry, Clinical & Translational Allergy, BMJ Open, Environmental Evidence, Implementation Science, Injury Prevention, International Journal of Medical Informatics, International Journal of Stroke, Journal of Agricultural and Food Chemistry, Journal of Medical Internet Research, Primary Care, Respiratory Journal, Research Journal of Applied Sciences, Engineering and Technology, BMC Trials, and World Journal of Surgical Oncology. For groups with more than 100 protocols, 100 were randomly chosen for analysis. For groups with less than 100 protocols, all the protocols were examined. Figure 1 outlines the selection process.

Definition of variables

The unpublished protocols

We recorded various factors that might be associated with an increased likelihood of publication, as described below. Published protocols are protocols that were used in preparing at least one publication.

Unpublished protocols either a) produced no publications or b) had been registered for use in at least one publication, but the publication in question was withdrawn. Except where noted, this information came from the published protocol.

Journal metrics

Elsevier's Scopus has two older citation analysis metrics: Source Normalized Impact per Paper (SNIP) and SCImago Journal Rank (SJR)[9]. Citescore is a new metric of Scopus[10]. The factors recorded were the impact factor (IF), the number of citations, Citescore, SNIP, and SJR.

We used the journal metrics of the review's publication year. We performed our metrics search between July and December 2018. The Clarivate website was used to identify the IF, while the Scopus database was used for Citescore, SNIP, and SJR.

Countries

Based on the corresponding author of the protocol, we used the Institute for Scientific Information (ISI) list of developed countries to determine whether the author was based in a developing or a developed country[11], as it is known that developing countries usually have lower scientific production due to their financial burden. We also determined whether the corresponding author was based in a native English-speaking country, defined as Australia, Canada, New Zealand, the United Kingdom, or the United States. In our study, the countries of corresponding authors were working countries, not home countries.

Expertise of the corresponding authors

We collected information on the h-index of the corresponding authors of each protocol. The h-index was retrieved from Scopus (Elsevier, Amsterdam, Netherlands) between July and December 2018. We used Scopus because of its efficient author identification algorithm[12]. If there was more than one profile for the corresponding author, we selected the one with the higher h-index. We classified the authors into three groups based on the h-index: <6 , $6 < x < 12$, and > 12 [13].

Another variable – “Corresponding author is the first author” – was included, and it was considered positive if the first author and the corresponding author of the article were the same.

Publication date

The publication date of each protocol was retrieved from its respective registered website. If it was co-registered, we chose the publication year of 2013 so long as the actual publication occurred on or before that year.

In order to check whether the project was published after protocol registration, we employed the following steps. Firstly, the protocol was searched for the authors' indication that the project was published. If there was not a statement, the protocol's title, the first author name, and then the protocol's registered number were searched on Google. Finally, we looked for the first authors' ResearchGate account and checked if the protocol was published under another name. If a protocol had more than one publication, we selected

the earliest. The duration of publication was calculated from the publication date of the protocol to the publication date of the paper.

Type of systematic reviews

We adopted a recent classification of systematic reviews, published in 2018, in order to adequately categorize various reviews' protocols [7].

Funding support

Data regarding internal or external funds were collected according to the Cochrane handbook 5.1[1, 14].

Statistical analysis

Statistical analysis was performed in SPSS statistics version 25. Descriptive analysis was applied to summary protocol characteristics. We used the Mann-Whitney U test and the Kruskal Wallis H test to determine the difference of continuous factors between groups, while Fisher's exact test was performed to compare the dichotomous variables (yes or no variables). All factors associated with our outcome of interest, i.e. factors that directly related to the cause of un-publishing protocols in univariable analysis, would be included and re-analyzed in a multivariable logistic regression model. As multicollinearity was present, two multivariable logistic regression models were performed. A two-sided p-value of < 0.05 was considered statistically significant.

Results

A total of 3200 protocols were obtained after searching five databases – Embase, PROSPERO, Cochrane, PubMed, and Scopus – along with 196 protocols from Cochrane by manual searching. After applying inclusion/exclusion criteria, 2230 protocols entered the randomization selection stage, which resulted in a final dataset of 397 protocols. 121 protocols were registered in multiple webs (i.e. databases). The number of single-web-registered protocols (i.e. published in one single database) in PROSPERO, Cochrane, and JBI were respectively 100, 86, 51, and the remaining protocols were allocated in SJR databases ($n = 16$) or other websites ($n = 23$). For further analysis, all of the 397 protocols were later reclassified, as described in Fig. 1.

Regarding protocol characteristics, a majority of corresponding authors were working in developed countries (80.6%) and native-English-speaking countries (59.7%). The median value of the Scopus h-index was 8 with IQR of 14 (Range 0–94). There were 301 authors (75.8%) who held both the first author position and the corresponding author position. Apart from corresponding author information, the median number of authors per protocol was 5 with IQR of 3 (Range 3–6), and there were about 114 protocols (28.7%) where participation from external co-workers was recognized. Approximately 99 protocols (24.9%) declared no funding support, and 81 protocols (20.4%) did not report a study funding source. We

also found that 118 (29.7%) protocols were updated in the study process after registration. Other information was summarized in Table 1.

Table 1

Protocol characteristics comparison between published protocols and unpublished protocols (N = 397)

Protocol's characteristics	Total (N = 397)	Published protocols (N = 244)	Unpublished protocols (N = 153)	p- value
Information of corresponding author				
Scopus's h-index of corresponding author				0.118
Value	8 (14.5)	9.5 (15.5)	5 (3.0)	
Corresponding author based in developed countries				0.019
Yes	320 (80.6)	206 (84.4)	114 (74.5)	
No	77 (19.4)	38 (15.6)	39 (25.5)	
Corresponding author based in English-speaking countries				0.006
Yes	237 (59.7)	159 (65.2)	78 (51.0)	
No	160 (40.3)	85 (34.8)	75 (49.0)	
Corresponding author is the first author				0.278
Yes	301 (75.8)	180 (73.8)	121 (79.1)	
No	96 (24.2)	64 (26.2)	32 (20.9)	
Country of corresponding authors				0.033
United Kingdom	79 (19.9)	58 (23.8)	21 (13.7)	
Australia	62 (15.6)	44 (18.0)	18 (11.8)	
Canada	46 (11.6)	28 (11.5)	18 (11.8)	
United States	47 (11.8)	26 (10.7)	21 (13.7)	
China	24 (6.0)	14 (5.7)	10 (6.5)	
Others	139 (35.0)	74 (30.3)	65 (42.5)	
Other characteristics				

Protocol's characteristics	Total (N = 397)	Published protocols (N = 244)	Unpublished protocols (N = 153)	p-value
Number of authors registered in protocols				0.333
Value	5 (3)	5 (3)	4 (3)	
External co-worker				0.909
Yes	114 (28.7)	71 (29.1)	43 (28.1)	
No	283 (71.3)	173 (70.9)	110 (71.9)	
Registered database				0.186
Cochrane only	86 (21.7)	49 (20.1)	37 (24.2)	
PROSPERO only	100 (25.2)	72 (29.5)	28 (18.3)	
JBI only	51 (12.8)	28 (11.5)	23 (15.0)	
SJR	16 (4.0)	10 (4.1)	6 (3.9)	
Others	23 (5.8)	12 (4.9)	11 (7.2)	
Duo register (*)	121 (30.5)	73 (29.9)	48 (31.4)	
Study funding				0.016
No funding support	99 (24.9)	63 (25.8)	36 (23.5)	
Internal fund only	74 (18.6)	40 (16.4)	34 (22.2)	
Having external fund (with/without internal fund)	143 (36.0)	100 (41.0)	43 (28.1)	
No report	81 (20.4)	41 (16.8)	40 (26.1)	
Protocol was updated				0.033
Yes	118 (29.7)	82 (33.6)	36 (23.5)	
No	279 (70.3)	162 (66.4)	117 (76.5)	
Type of systematic review				0.667

Protocol's characteristics	Total (N = 397)	Published protocols (N = 244)	Unpublished protocols (N = 153)	p-value
Effectiveness	311 (78.3)	191 (78.3)	120 (78.4)	
Experiential or qualitative	22 (5.5)	12 (4.9)	10 (6.5)	
Etiology or risk	21 (5.3)	11 (4.5)	10 (6.5)	
Diagnostic test accuracy	16 (4.0)	11 (4.5)	5 (3.3)	
Others	27 (6.8)	19 (7.8)	8 (5.2)	
Statistical analysis tests: a: Fisher's exact test, b: Mann-Whitney U's test.				
Descriptive information was reported as: a: N(%), b: Median (IQR).				
(*): 121 duo-registered protocols were all registered in Prospero, along with another database: Cochrane (14), JBI (46), SJR (46), Others (15).				

Our study results show that there were some differences in protocol characteristics among databases registered (**Supplementary Table 1**). We also found that protocols which were registered in the JBI database were more likely to be published in journals with a low Citescore, SNIP score, SJR score, and impact factor (**Supplementary Fig. 2–5**). However, there was no difference in publication chance among databases registered ($p = 0.186$). Approaching author information through the h-index of a corresponding author seems to be unrelated to the publication chance ($p = 0.118$). Protocols with corresponding authors working in developed countries ($p = 0.019$) and native-English-speaking countries ($p = 0.006$) were more likely to be published as papers. Besides that, study funding status ($p = 0.016$), along with an updated status of the protocol ($p = 0.033$), were also recorded as having a relationship with the outcome of interest. The overview information of other factors was presented in Table 1.

All potential protocol characteristics related to being published as papers, determined by univariable analysis, were summarized in Table 2. The multivariable logistic regression shows that the updated status of protocols [OR = 1.7, 95% CI = 1.0–2.8, $p = 0.037$], along with having external funding support [OR = 2.1, 95% CI = 1.2–3.8, $p = 0.010$], were associated with the publication of protocols. Notably, although being recorded as a potential factor, the country of corresponding authors showed no relationship at all with the outcome of interest.

Table 2
Factors associated with the paper publication of protocols

Protocol's characteristics	Univariable analysis		Multivariable analysis 1		Multivariable analysis 2	
	COR (95% CI)	p-value	AOR (95% CI)	p-value	AOR (95% CI)	p-value
Corresponding author based in developed countries						
Yes	1.9 (1.2–3.1)	0.016	1.4 (0.7–2.6)	0.347	1.7 (0.8–3.4)	0.139
No	reference	-	reference	-	reference	-
Corresponding author based in English-speaking countries						
Yes	1.8 (1.2–2.7)	0.005	1.5 (0.9–2.6)	0.105	-	-
No	reference	-	reference	-	-	-
Country of corresponding authors						
United Kingdom	2.4 (1.3–4.4)	0.004	-	-	1.8 (0.9–3.7)	0.080
Australia	2.1 (1.1–4.1)	0.020	-	-	2.0 (1–4.2.0)	0.054
Canada	1.4 (0.7–2.7)	0.368	-	-	1.0 (0.5–2.1)	0.938
United States	1.1 (0.6–2.1)	0.805	-	-	1.0 (0.5–2.0)	0.923
China	1.2 (0.5–3.0)	0.644	-	-	1.8 (0.6–4.8)	0.271
Others	reference	-	-	-	reference	-
Study funding						
No funding support	1.7 (0.9–3.1)	0.079	1.7 (0.9–3.1)	0.105	1.8 (0.9–3.3)	0.078
Internal fund only	1.1 (0.6–2.2)	0.669	1.1 (0.5–2.1)	0.866	1.0 (0.5–2.0)	0.992
Having external fund (with/without internal fund)	2.3 (1.3–4.0)	0.004	2.1 (1.2–3.8)	0.010	2.1 (1.2–3.9)	0.014
No report	reference	-	reference	-	reference	-
Protocol was updated						

Protocol's characteristics	Univariable analysis		Multivariable analysis 1		Multivariable analysis 2	
	COR (95% CI)	p-value	AOR (95% CI)	p-value	AOR (95% CI)	p-value
Yes	1.6 (1.0-2.6)	0.033	1.7 (1-2.8)	0.037	1.7 (1.0-2.8)	0.033
No	reference	-	reference	-	reference	-
Statistical analysis tests: Logistic regression. The multivariable analysis was divided into 2 models due to multicollinearity.						

Discussion

In this review, we investigated the main factors that made systemic review and meta-analysis protocols more likely to be published. Firstly, we found that corresponding authors in both developed countries (206/320, $p = 0.018$) and English-speaking countries (158/236, $p = 0.006$) had a higher chance of publishing the paper of their registered protocols. We found that the highest percentage of corresponding authors for both protocols and reviews were from the UK, representing 19.9% and 23.8% of the authors, respectively. Our models of multivariable logistic regression revealed that two main factors significantly impact the publication outcome, which included the updated protocol status of the published review paper (OR: 1.7, 95% CI: 1.0-2.8, $p = 0.037$) and external funding (OR: 2.1, 95% CI: 1.2–3.8, $p = 0.01$) (Table 1). This is most likely explained by more funding opportunities, better background research support, and the fact that English is the primary spoken language in these countries. However, our multivariable analysis did not show any significant difference among these countries compared to the other ones. It should also be noted that most reviews are published in English or included English studies. These results are supported by other studies that found that about 30% of corresponding authors of protocols registered in PROSPERO are from the UK[2, 5].

Contrastingly, another study published in 2016 found that UK corresponding authors were only 16%, while those of China were 21%. These results were based on searching the MEDLINE database[15]. Interestingly, China is demonstrating a rapid growth in the number of meta-analyses conducted, despite being neither a developed nor an English-speaking country[16].

We found that protocols with external funding had a higher chance of being published than those with internal funding. This result is consistent with that of Tsujimoto et al. (2017), who reported that PROSPERO protocols that received funding were associated with better publishing chances[4]. Lack of funding is also considered a barrier to publishing SRs[6]. However, in terms of quality, it was noticed that reviews funded by internal sources, such as academic institutions, were of a higher quality than those with external funding or those that failed to report funding status[17]. Our explanation for this issue could be summarized in one crucial finding, which is the correlation between funding and the number of team

members. In 2017, Borah et al. demonstrated that increasing the number of authors results in increased productivity[5].

Since it appears that a positive relationship between funding and quality has been established, protocols databases should exercise more requirements on declaring funding sources. Several reports have shown that 16.2% – 36% of SR/MA protocols fail to declare funding[2, 15, 18]. Our results revealed that about 20.4% of protocols have failed to report funding. More than 80% of JBI protocols failed to report funding. On the other hand, 99% of PROSPERO protocols reported funding.

It is believed that systematic reviews need to be kept up to date[19]. In 2016, a study revealed that about 10% of published reviews are updated, and remarkably, 81% of these updates were just Cochrane SRs[15]. Another result from Oral Health Cochrane Systematic Reviews showed that 14.5% of all reviews have been updated[20]. Univariate analysis reveals that there is a significant correlation between updated protocol and publication (OR = 1.6, 95% CI 1.0–2.60, $p = 0.033$), and that 33.6% of the published reviews had updated protocols.

Updated protocols were associated with a change in the author list, but not associated with the time from protocol registration to publication (**Supplementary Fig. 1**). Previously, it was found that a shorter time to publication might result in higher chances of the review being updated, and that a longer time to publication is often associated with the review having two published protocols, which hints at changes in the review plan[3].

In this study, we noticed that only 61.46% of protocols resulted in publications. In other words, more than one-third of the registered protocols either did not publish their results or are not yet finished. A recent study in 2018 found that about one-third of protocols did not have any publication within 3–5 years; however, only 80 papers were analyzed in the study, which is a relatively small sample size[2]. Two other studies found that this rate of nonpublication reached 12.4%, measured through a survey among investigators, and a second study found that 26% of PROSPERO protocols did not have any publication after 65 months of protocol registration[4, 6]. In 2008, a study found that 19.1% of Cochrane protocols were unpublished[3].

In our study, Cochrane protocols had one of the lowest rates of publication at 57%. This could be explained by the meticulous process that any Cochrane protocol or review must go through, and the editorial process strategies employed by Cochrane that promote good reporting, such as the MECIR standards[21]. JBI protocols had the lowest rate of publication, at only 55%. Many previous studies reported that Cochrane reviews had higher quality reporting methods than non-Cochrane reviews, which might take more time and reduce the overall chance of publication[15, 22]. A recent review suggests that the overall quality of registered reviews is higher than that of non-registered reviews [23].

In our analysis, we found that the median number of authors in a review was 5. A previous study estimated that at least 5 reviewers and an average of 67 weeks are required to complete a well-conducted systematic review[5]. Page et al. (2016) also found that the median number of authors is 5 (IQR 4–6)[15].

Two other studies have found that the number of authors is 7 (IQR 5–11) and 6. However, both of them had some limitations, such as a small sample size or including protocols from only one database[2, 18].

It has been reported that the best methodological quality SR/MAs were conducted by groups of authors with high levels of scientific experience, with a median h-index of 14[24]. In 2018, Schreiber et al. found that higher academic rankings among academic physicians in different specialties were associated with higher h-indices. On average, assistant professors have an h-index of 2–5, associate professors 6–10, and full professors 12–24[13]. In our study, we found that the median h-index of corresponding authors in protocols with published reviews was 9.5, while in unpublished protocols it was only 5. Since JBI is a nursing protocol database, it was associated with the lowest median h-index among corresponding authors.

We adopted a recent classification of systematic reviews, published in 2018, in order to adequately categorize various reviews' protocols[7]. Interestingly, we discovered that most of the published protocols belonged to the effectiveness group (78.3%). This result was supported by another study published in 2016, which found that about 55% of the SRs were classified as therapeutic, 25% as epidemiology, 11% as diagnosis/prognosis, and 10% as other[15]. Also, when reviews from the Cochrane databases were compared to ones from high-impact journals in cancer, it was found that Cochrane reviews were less likely to address questions concerning prognosis[25].

Remarkably, we found that 30.5% of the protocols in our random sample pool were registered in two protocols databases. Previous studies have found a higher proportion of co-registration of protocols. 45% of non-Cochrane protocols were published in both journals and PROSPERO[26]. 89.2% of protocols published in 'BMC Systematic Reviews' were also registered in PROSPERO[18].

In boxplot-based analyses of several different scoring systems, we observed considerable impact factor-like potential variability of publications between databases studied. The median scores of dual-published papers did not have substantial variation from the broader sample medians, indicating that dual-publication is not likely to influence the usage of a given article. However, individual databases such as Cochrane and JBI often had substantial and potentially significant deviations from the overall means. While a small overall sample size may contribute to some of Cochrane's deviation, caution should be taken in selecting the database one uses to generate and obtain meta-analyses due to potential bias in article quality and relevance to current topical discourse. Moreover, we found that 14.8% of the Cochrane protocols were withdrawn, compared to the previously reported rate of 12.7% [3]. In the past decade, new emerging journals have made it easier for authors to publish their SRs, even if there is already an SR on the same topic published at the same time. This has resulted in a large number of overlapping studies, and a previous study has estimated that about 67% of meta-analyses have at least one overlapping meta-analysis within three years[27].

There were some limitations to our study. We did not cover all the registered SR protocols in 2013. However, we performed a randomization of 100 Cochrane and PROSPERO protocols. Other studies have included more variables like page count, certificate of insurance (COI), funding by the pharmaceutical

industry, funding by academic institutions, randomized clinical studies (RCTs) as primary studies, meta-analysis included, journal bibliometrics, author bibliometrics, and others[17], but these were not available in PROSPERO records, so we could not adjust for them in our analysis. The main reasons reported for non-publication were lack of time, overly broad SR scope, and few studies eligible for SRs, as well as rejection[6].

Conclusion

Systematic review and meta-analysis studies require significant attention and a careful literature review before conducting them. Factors that increase chances of publication include a sufficient timeline, an adequate number of qualified authors, and a good source of external or internal funding. Choosing the most applicable database and periodically updating the review status is highly recommended to ensure a chance of publication among the highest Citescore journals, since this type of publication is placed on top of the hierarchy of evidence-based medicine and greatly influences clinical practice.

Declarations

Ethics approval and consent to participate: not applicable

Consent for publication: not applicable

Available of data and materials: all data generated or analysed during this study are included in this published article and its supplementary information files

Competing interest: none of the authors have any conflicts of interest to declare.

Funding: not applicable

Authors contributions: NTH was responsible for the idea and study design. All authors performed the screening and extraction under the supervision of NTH. Data analysis and its interpretation were made by LLTT and LHNM. Tables and figures were done by LHNM and LLTT. LHNM, GMT, and OMM revised the final revision versions. All authors contributed to the manuscript writing and approval of the final version.

Acknowledgments: We would like to offer our sincere condolence to Prof. Hosni Salem, who had recently passed away. Prof Hosni Salem has significantly contributed to the manuscript preparation and revision.

References

1. J. Higgins and S. G. (editors), "Cochrane handbook for systematic reviews of interventions version 5.1. 0 [updated March 2011]," The Cochrane Collaboration, 2011, ch. Chapter 4: Guide to the contents of a Cochrane protocol and review.
2. K. Allers, F. Hoffmann, T. Mathes, and D. J. J. o. c. e. Pieper, "Systematic reviews with published protocols compared to those without: more effort, older search," vol. 95, pp. 102-110, 2018.

3. A. C. Tricco, J. Brehaut, M. H. Chen, and D. J. P. O. Moher, "Following 411 Cochrane protocols to completion: a retrospective cohort study," vol. 3, no. 11, p. e3684, 2008.
4. Y. Tsujimoto *et al.*, "Majority of systematic reviews published in high-impact journals neglected to register the protocols: a meta-epidemiological study," vol. 84, pp. 54-60, 2017.
5. R. Borah, A. W. Brown, P. L. Capers, and K. A. J. B. o. Kaiser, "Analysis of the time and workers needed to conduct systematic reviews of medical interventions using data from the PROSPERO registry," vol. 7, no. 2, p. e012545, 2017.
6. A. C. Tricco *et al.*, "An international survey indicated that unpublished systematic reviews exist," vol. 62, no. 6, pp. 617-623. e5, 2009.
7. Z. Munn, C. Stern, E. Aromataris, C. Lockwood, and Z. J. B. m. r. m. Jordan, "What kind of systematic review should I conduct? A proposed typology and guidance for systematic reviewers in the medical and health sciences," vol. 18, no. 1, p. 5, 2018.
8. R. Borah, A. W. Brown, P. L. Capers, and K. A. Kaiser, "Analysis of the time and workers needed to conduct systematic reviews of medical interventions using data from the PROSPERO registry," (in eng), *BMJ Open*, vol. 7, no. 2, p. e012545, 02 2017, doi: 10.1136/bmjopen-2016-012545.
9. L. Colledge, F. de Moya-Anegón, V. P. Guerrero-Bote, C. López-Illescas, and H. F. Moed, "SJR and SNIP: two new journal metrics in Elsevier's Scopus," *Insights*, vol. 23, no. 3, p. 215, 2010.
10. J. A. T. Da Silva and A. R. Memon, "CiteScore: A cite for sore eyes, or a valuable, transparent metric?," *Scientometrics*, vol. 111, no. 1, pp. 553-556, 2017.
11. WorldBank, "World Bank Country Classifications," 2019.
12. B. Walker, S. Alavifard, S. Roberts, A. Lanes, T. Ramsay, and S. Boet, "Inter-rater reliability of h-index scores calculated by Web of Science and Scopus for clinical epidemiology scientists," vol. 33, no. 2, pp. 140-149, 2016, doi: 10.1111/hir.12140.
13. W. E. Schreiber and D. M. J. A. j. o. c. p. Giustini, "Measuring Scientific Impact With the h-Index: A Primer for Pathologists," vol. 151, no. 3, pp. 286-291, 2018.
14. J. Higgins, "Cochrane handbook for systematic reviews of interventions. Version 5.1. 0 [updated March 2011]. The Cochrane Collaboration," *www.cochrane-handbook.org*, 2011.
15. M. J. Page *et al.*, "Epidemiology and reporting characteristics of systematic reviews of biomedical research: a cross-sectional study," vol. 13, no. 5, p. e1002028, 2016.
16. J. P. A. Ioannidis, C. Q. Chang, T. K. Lam, S. D. Schully, and M. J. Khoury, "The Geometric Increase in Meta-Analyses from China in the Genomic Era," *PLOS ONE*, vol. 8, no. 6, p. e65602, 2013, doi: 10.1371/journal.pone.0065602.
17. F. Gómez-García *et al.*, "Systematic reviews and meta-analyses on psoriasis: role of funding sources, conflict of interest and bibliometric indices as predictors of methodological quality," vol. 176, no. 6, pp. 1633-1644, 2017.
18. T. Rombey, K. Allers, T. Mathes, F. Hoffmann, and D. J. B. m. r. m. Pieper, "A descriptive analysis of the characteristics and the peer review process of systematic review protocols published in an open peer

- review journal from 2012 to 2017," vol. 19, no. 1, p. 57, 2019.
19. D. Sutton, R. Qureshi, and J. J. J. o. c. e. Martin, "Evidence reversal—when new evidence contradicts current claims: a systematic overview review of definitions and terms," vol. 94, pp. 76-84, 2018.
 20. N. Pandis, P. S. Fleming, H. Worthington, K. Dwan, and G. J. P. o. Salanti, "Discrepancies in outcome reporting exist between protocols and published oral health Cochrane systematic reviews," vol. 10, no. 9, p. e0137667, 2015.
 21. J. Chandler, R. Churchill, J. Higgins, T. Lasserson, and D. J. M. e. o. C. i. r. Tovey, "Methodological standards for the conduct of new Cochrane Intervention Reviews Version 2.3 (2013)," 2016.
 22. M. Petticrew, P. Wilson, K. Wright, and F. J. B. B. M. J. Song, "Quality of Cochrane reviews: Quality of Cochrane reviews is better than that of non-Cochrane reviews," vol. 324, no. 7336, p. 545, 2002.
 23. S. Sideri, S. N. Papageorgiou, and T. Eliades, "Registration in the international prospective register of systematic reviews (PROSPERO) of systematic review protocols was associated with increased review quality," (in eng), *J Clin Epidemiol*, vol. 100, pp. 103-110, 08 2018, doi: 10.1016/j.jclinepi.2018.01.003.
 24. J. L. Sanz-Cabanillas *et al.*, "Author-paper affiliation network architecture influences the methodological quality of systematic reviews and meta-analyses of psoriasis," vol. 12, no. 4, p. e0175419, 2017.
 25. M. Goldkuhle, V. M. Narayan, A. Weigl, P. Dahm, and N. J. B. o. Skoetz, "A systematic assessment of Cochrane reviews and systematic reviews published in high-impact medical journals related to cancer," vol. 8, no. 3, p. e020869, 2018.
 26. I. Viguera-Guerra *et al.*, "Evolution of international collaborative research efforts to develop non-Cochrane systematic reviews," vol. 14, no. 2, p. e0211919, 2019.
 27. K. C. Siontis, T. Hernandez-Boussard, and J. P. J. B. Ioannidis, "Overlapping meta-analyses on the same topic: survey of published studies," vol. 347, p. f4501, 2013.

Supplementary Table

Supplementary Table 1 is not available with this version.

Figures

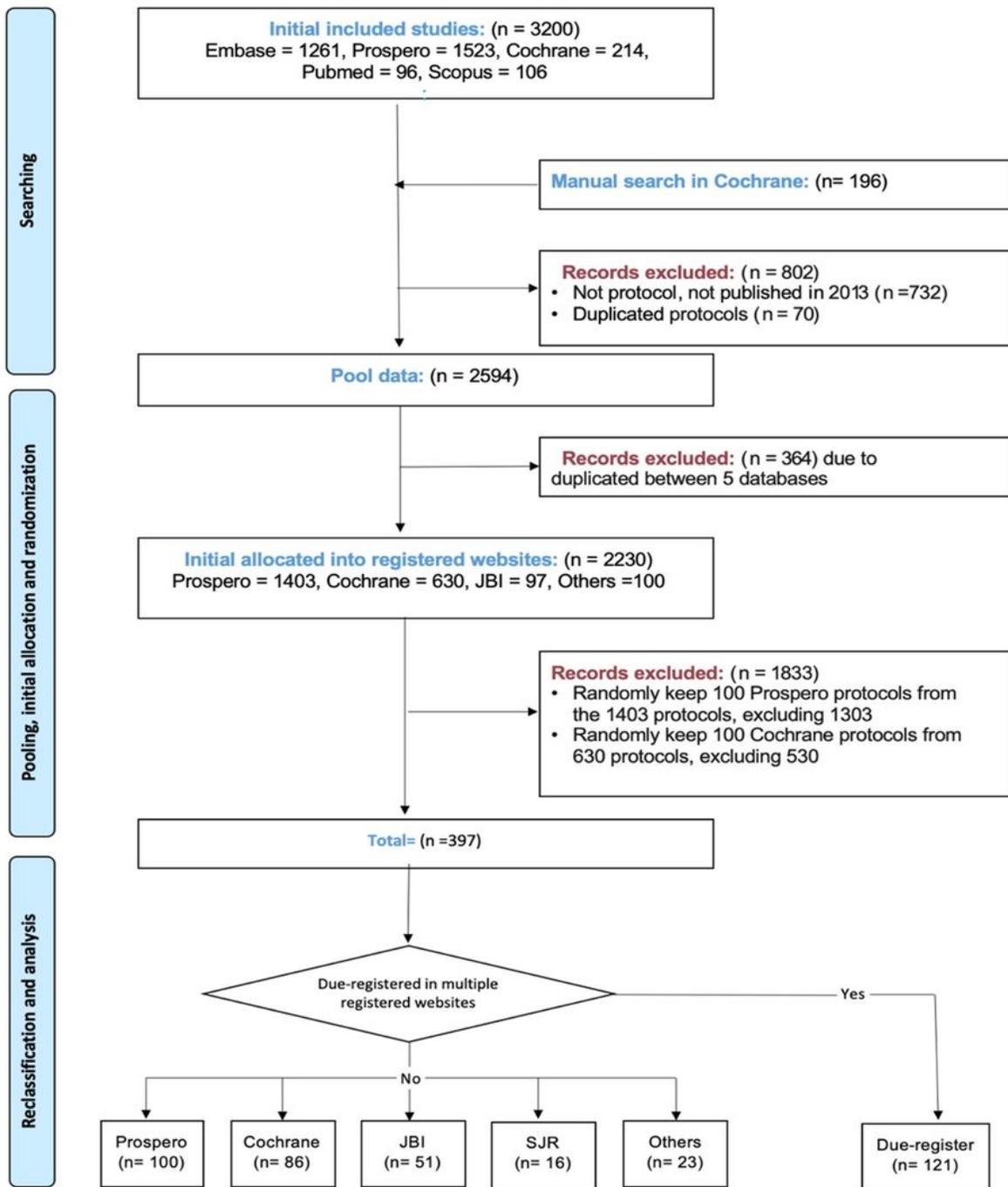


Figure 1

Flow diagram for study scheme steps

Supplementary Files

This is a list of supplementary files associated with this preprint. Click to download.

- [SupplementaryFigure1.tiff](#)
- [SupplementaryFigure2.tiff](#)
- [SupplementaryFigure3.tiff](#)
- [SupplementaryFigure4.tiff](#)
- [SupplementaryFigure5.tiff](#)